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THE INCIDENCE OF EPIDEMIC INFLUENZA. 1918-19*

A FURTHER ANALYSIS ACCORDING TO AGE, SEX, AND COLOR OF THE RECORDS OF MORBIDITY AND MORTALITY OBTAINED IN SURVEYS OF 12 LOCALITIES

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Introduction

This report, one of a series of papers from this office on the epidemiology of influenza, is devoted to a discussion of the incidence of the disease in the different sex, age, and color groups of the population during the 1918–19 epidemic, as indicated by surveys made at that time by the United States Public Health Service in certain localities. Summaries of the results obtained in these surveys were published shortly after the surveys were finished; but, as a further contribution to the epidemiological studies of the disease, it seems desirable to give a more detailed account of the results at this time.

The Public Health Service conducted special surveys in a number of widely scattered localities as soon as the 1918-19 epidemic in these

[•] From the Office of Statistical Investigations, in cooperation with the Office of Industrial Hygiene and Sanitation, United States Public Health Service.

¹ Influenza in Maryland: Preliminary Statistics of Certain Localities. By W. H. Frost and Edgar Sydenstricker. Pub. Health Rep., Mar. 14, 1919. Reprint No. 510.

Epidemiology of Influenza. By W. H. Frost and Edgar Sydenstricker. Pub. Health Rep., Aug. 15, 1919. Reprint No. 550. (Reprinted from J. A. M. A., vol. 73, No. 5, Aug. 2, 1919.)

Statistics of Influenza Morbidity, with Special Reference to Certain Factors in Case Incidence and Case Fatality. By W. H. Frost. Pub. Health Rep., Mar. 12, 1920. Reprint No. 586.

Variations in Case Fatality during the Influenza Epidemic of 1918. By Edgar Sydenstricker. Pub. Health Rep., Sept. 9, 1921. Reprint No. 692.

A list of epidemiological studies of influenza made by the Public Health Service will be found at the end of this article.

places appeared to have reached its close. The purpose was to determine for a population of known sex, age, and color composition the approximate incidence of the disease in sample areas of a number of widely scattered localities, and also to determine the relations between cases of influenza, cases of pneumonia, and deaths from these causes in so far as the number of observations would permit.

It was necessary to limit the surveys for the most part to localities in which the Public Health Service was at the time maintaining previously established organizations prepared to collect the requisite data reliably and efficiently; but in so far as practicable, the communities were chosen to represent the different geographical sections of the United States. Reference to Table 1 will show that, with the exception of the far West, this object was accomplished in a reasonably satisfactory manner. San Francisco was the only city west of San Antonio, Tex., and Des Moines, Iowa.

The survey included (a) 10 cities, varying in population from 22,500 to 680,000; (b) certain small towns of Maryland; and (c) one rural county of Maryland. The minor towns surveyed in Maryland are usually treated as a single statistical group in this report. In the case of Charles County, the entire population, rather than a sample of it, was made the basis of the survey. This particular survey was made by employees of the U. S. Bureau of the Census, funds having been transferred to that bureau by the Public Health Service for the purpose. The data were tabulated and analyzed by the Public Health Service. Although the canvass included the whole county, one of 12 enumeration districts was later dropped from the records, owing to the presence of a proving ground (Indianhead) which made that district unrepresentative of a general population.

In the case of Louisville, the canvass was made before the wave of the epidemic had run its full course; but in all the other localities the canvass is believed to have comprised practically the whole of the epidemic period. In Baltimore and San Francisco second surveys were made in January and February, respectively, to obtain a record of recrudescences which had taken place in the interval. The cases occurring during these recrudescences are included in the data here reported.

In the case of Spartanburg, S. C., some time after the completion of the canvass in the city itself, an additional survey was made of adjacent mill villages. These villages had a disproportionately large population of one selected class—mill workers—and for this reason the Spartanburg data are not altogether comparable with those collected in other localities.

The canvasses were made as soon as possible after the subsidence of the autumn (1918) wave of the epidemic in each locality. The following table will show the dates on which the surveys were begun and ended:

TABLE 1.—Localities in which 1918-19 surveys were made, with dates of surveys, estimated total populations, and number of persons canvassed

	Dates o	f canvass	Total	Population canvassed		
Locality	Begun	Completed	popula- tion (esti- mated) ¹	Number of persons	Per cent of total popula- tion	
New London, Conn. Baltimore, Md. Minor Maryland towns; 3	Dec. 2, 1918 Nov. 20, 1918	Dec. 18, 1918 Jan. 31, 1919 ²	25, 000 680, 000	7, 933 33, 361	31. 7 4. 9	
CumberlandFrederick	Dec. 1, 1918 Nov. 27, 1918	Dec. 6, 1918 Nov. 30, 1918	27, 300 11, 340	5, 194 2, 311	19. 0 20. 4	
Lonaconing	Dec. 4, 1918	Dec. 11, 1918	2,000	1,730	86. 5	
Salisbury	Dec. 10, 1918 Dec. 7, 1918	Dec. 12, 1918	9, 000 850	1, 727 718	19. 2 84. 4	
Linganore District (Frederick Co.) 4	Nov. 29, 1918	do	1, 000	688	68.8	
Quantico 4	Dec. 1, 1918	Dec. 10, 1918	2,000	114	5. 7	
Charles County, Md	(8)	(5)	6 18, 326	⁷ 18, 326	100. 0	
Spartanburg, S. C.	Dec. 5, 1918	Dec. 31, 1918	22, 500	5, 257	23. 4	
Augusta, GaMacon, Ga	Feb. 1, 1919 Dec. 4, 1918	Feb. 8, 1919 Dec. 14, 1918	55, 000 50, 000	4, 123 7, 905	7. 5 15. 8	
Des Moines, Iowa	Jan. 31, 1919	Feb. 8, 1919	115,000	5, 857	5. 1	
Louisville, Ky		Dec. 27, 1918	245, 000	12,002	4.9	
Little Rock, Ark	Dec. 2, 1918	Jan. 13, 1919	65, 000	9, 920	15. 3	
San Antonio, Tex.	Dec. 5, 1918	Dec. 22, 1918	150, 000	12, 534	8. 4	
San Francisco, Calif	do	Feb. 21, 19198	475, 000	18, 682	3.9	

Total number of persons canvassed in minor Maryland towns was 12,482.

Rural area. Census as of Mar. 12, 1919.

⁸ The population included in survey made in November and December was recanvassed in February in order to record cases occurring during a recrudescence of the epidemic,

The population estimates contained in the third column of Table 1 require some comment. Since the epidemic occurred while this country was at war, a number of factors (principally the withdrawal of males for military service) tend to make population estimates more than usually unreliable in the present instance.² By the time of the 1920 census the unusual distribution had given way to a more normal An estimate based on the 1910 and the 1920 censuses will thus not afford a reliable indication of the population of individual localities in the fall of 1918. Indeed, a satisfactory estimate is impossible, however it be derived. But since the data here presented deal almost entirely with actually enumerated populations in sample areas, estimetes of the total population are employed in only a few instances. The estimates adopted for use in the table are based on a number of factors, including an intercensal estimate of the population (calculated arithmetically), allowance having been made for the withdrawal of males for military service; population estimates based on the normal death rates from all causes, exclusive of respiratory infections; infor-

¹ Estimated as of July 1, 1918; revised on the basis of other data.

² The population included in survey made in November and December was recanvassed in January in order to record cases occurring during a recrudescence of the epidemic.

Actual count in February-March, 1919.
7 One enumeration district was later excluded from the study (see p. 304), leaving data for 16,147 canvassed persons

² This question has been given detailed consideration in the article, "Difficulties in Computing Civil Death Rates for 1918", by Edgar Sydenstricker and Mary L. King. Public Health Reports, Feb. 13, 1920. Reprint No. 583.

mation secured by Public Health Service officers located in the individual localities; and other available information.

Data were collected by intelligent inspectors working under specific instructions and careful supervision. In each locality these inspectors made a house-to-house canvass in 10 or more enumeration districts so situated geographically as to give, presumably, a fair sample of the general population of the city. Each district contained approximately the same number of families. Homes at which information was not available when the inspector called (owing to the absence of the adults, or for other reasons) were not counted. The effort was made to canvass in each city not less than 5,000 persons, in order to give a group sufficient for simple statistical analyses, and in cities of more than 100,000 population to increase this number so as to give not less than 5 per cent of the total population. These conditions were generally fulfilled.

Regarding each individual in the canvassed populations, the inspectors recorded the name, color, sex, and age at last birthday; whether or not sick since September 1, 1918, with "influenza," "pneumonia," or illness suspected to be influenza (classed as "doubtful"); date of onset, duration, and severity of such illnesses (whether "severe," "moderate," or "light"); and date of death, if death resulted. Regarding each household, the inspectors recorded the number of rooms occupied, and their impressions of the economic status of the family (whether "well-to-do," "moderate," "poor," or "very poor"). This point was recorded by the inspectors without instructions as to the possible definitions of each class.³

In making inquiry as to the type or nature of illness, the enumerators were instructed to ask which members of a family had "influenza," "flu," "grippe," "pneumonia," or "colds" since September 1, 1918. Persons who were said to have been only "feeling badly," or as having a "cold" were recorded as "doubtful" cases. If, however, the illness lasted not less than three days and was of such severity as to confine the patient to bed for the whole of one day, the case was classed as "influenza," unless otherwise diagnosed by the attending physician. Cases of illness, if definitely stated to be due to some cause other than "influenza," "pneumonia," or "colds," were not recorded. In view of the difficulties of diagnosis of influenza and the large number of mild cases indistinguishable from common colds, it was believed that the total morbidity from influenza during the epidemic period could be best represented by a figure which would include cases classified during the canvass as "influenza," "grippe," "pneumonia," and "doubtful." The widespread nature of the epidemic minimized the effect of minor

⁸ A special study of the data secured in relation to economic conditions has recently been issued: The Incidence of Influenza Among Persons of Different Economic Status during the Epidemic of 1918. By Edgar Sydenstricker. Pub. Health Rep., Jan. 23, 1931, vol. 46, No. 4, (Reprint No. 1444.)

respiratory illnesses unassociated with influenza. The inclusion of "pneumonia" in the figures was, of course, logical, since during the epidemic only a comparatively few pneumonia cases occurred which were not sequelae of influenza.⁴

The sources of error involved in the method of survey outlined are fully appreciated. Although the canvasses were made as soon as possible after current morbidity and mortality reports indicated that the wave of the epidemic had subsided, certain important points had been forgotten by the informants. Especially was this true in regard to the dates involved.

Another source of error arose from the fact that the families' statements were accepted as to diagnosis for a disease the diagnosis of which is especially difficult and uncertain. No other course was open; and it is confidently believed that, owing to the peculiar and widespread nature of the epidemic, the data obtained were sufficiently reliable when used in the mass.

A third source of error lay in the employment of enumerators not specially trained for this work. However, they were carefully selected and the inquiries were purposely made sufficiently simple to permit even untrained persons to obtain the data with such detailed written instructions as were furnished, if under careful supervision.

When due allowance is made for the inevitable errors incident to the method employed, it is still believed that the surveys gave data which represented with reasonable accuracy the influenza morbidity in the localities surveyed. This view is corroborated by a comparison of the chronological incidence of influenza cases in the surveyed populations and the chronological reported mortality for the population as a whole. In the following table this comparison is made for those surveyed localities for which death rates for the total populations were available by weeks.

^{&#}x27;That the inclusion of "doubtful" cases was justifiable for the epidemic period, for the purposes to which the data were to be put, is clearly indicated in the following table, from which it will be seen that, in Baltimore (the largest sample canvassed), cases classified as "influenza," "pneumonia," and "doubtful" show almost identical chronology. It is to be observed that the "doubtful" cases represent only 11 per cent of the total epidemic morbidity in Baltimore; for the surveys as a whole such cases were 7 per cent of the total—3.216 out of 42.920:

	Cases rep informa		Cases classified		Cases rep informa	Cases classified		
Week ended—	"Influ- enza," "grippe"	"Pneu- monia"	as "doubt- ful"	Week ended—	"Influ- enza," "grippe"	"Pneu- monia"	as "doubt- ful"	
Sept. 7. Sept. 14. Sept. 21. Sept. 28. Oct. 5. Oct. 12. Oct. 19. Oct. 26. Nov. 2.	28 52 126 271 1, 363 1, 605 1, 206 524 275	6 2 10 32 135 137 73 44 23	7 8 14 41 165 170 156 60 51	Nov. 9	86 47 29 24 5, 636 82. 1	12 7 6 3 490 7.1	18 15 16 15 736	

Table 2.—Weekly death rates per 100,000 from influenza-pneumonia in total population and weekly influenza case rates per 1,000 in canvassed populations of six localities, by weeks during epidemic of 1918-19 1

	Balti	more	Cumb	erland	Aug	rusta	Loui	sville	Little	Rock	San Fr	rancisco
Week ended-	Death rate in total popu- lation	Case rate in can- vassed popu- lation		Case rate in can- vassed popu- lation		vassed	Death rate in total popu- lation	Case rate in can- vassed popu- lation	total	Case rate in can- vassed popu- lation	Death rate in total popu- lation	Case rate in can- vassed popu- lation
1948 Sept. 7. St.pt. 14. Sept. 21. Sept. 22. Sept. 28. Oct. 12. Oct. 19. Oct. 25. Nov. 9. Nov. 16. Nov. 22. Nov. 30. Dec. 7. Dec. 14. Dec. 21. Dec. 22.	0 1. 0 2. 8 17. 2 82. 8 199. 6 157. 8 21. 6 7. 5 5. 9 8. 5 10. 9 8. 4	1. 2 1. 9 4. 5 10 3 49. 8 57. 3 43. 0 18. 8 10. 5 2. 1 1. 5 1. 3 1. 1 1. 4 2. 5	0 0 0 3.7 33.0 307.7 402.9 172.2 76.9 40.3 22.0 14.6 7.3 3.7 7.3 7.3	2. 1 3. 1 7. 7 33. 5 96. 6 123. 2 71. 4 25. 8 12. 7 6. 7 4. 8 4. 0 2. 1 1. 2	0 0 0 1.8 3.6 16.4 30.9 61.5 32.7 34.5 34.5 23.6 16.4 12.7 10.9	1. 5 .7 1. 7 2. 9 11. 2 14. 6 14. 1 7. 3 17. 0 11. 6 21. 6 15. 0 13. 1 11. 4 8. 2 16. 5 20. 1	0 1. 2 3. 7 1. 6 5. 7 37. 6 73. 5 28. 2 23. 7 15. 9 14. 9 25. 3 22. 4 37. 1 22. 4	1. 7 1. 1 1. 5 1. 4 22. 9 8. 6 13. 9 5. 5 5. 3 9. 6 7. 0 14. 3 18. 8 6. 6 1. 5	1. 5 0 0 18. 5 133. 8 146. 2 9. 2 7. 7 4. 6 13. 8 12. 3 12. 3 9. 2 10. 8	1. 0 1. 8 4. 9 8. 6 87. 3 95. 8 52. 1 22. 9 9. 8 9. 9 7. 5 7. 9 8. 1 3. 1 2. 4	1. 3 1. 3 2. 7 3. 2. 9 6. 3 27. 4 1165. 4 87. 2 41. 7 11. 8 10. 5 14. 8 28. 8 37. 5	1.5 2.5 3.5 3.7.2 13.2 27.5 28.8 16.2 9.5 4.3 9.0 8.2 6.8 12.4
1919 Jan. 4 Jan. 11 Jan. 18 Jan. 25 Feb. 1	7. 1 11. 0 12. 2 22. 1 20. 3	2. 5 4. 0 3. 3 1. 3	3. 7		29. 1 63. 6 70. 9 65. 5 25. 5	26. 7 44. 9 33. 2 16. 0 5. 1	9. 0 8. 2 8. 6 12. 2 8. 2		10. 8 13. 8 36. 9 21. 5 20. 0		40. 8 61. 1 65. 3 31. 4 12. 4	9.6 7.1 6.2 1.8

¹ Deaths classified according to date of death; cases classified according to date of onset.

The mortality rates are seen to follow the case incidence rates with considerable exactness, when one takes into account the necessary lag due to the difference between date of onset of the disease and death from it. So far as these few examples justify any conclusion it would appear that, for comparison between communities, with respect to chronology, mortality statistics give results quite similar to those derived from morbidity statistics. In the section on case fatality, however, it will be shown that entirely misleading results as to actual incidence of the disease would be obtained from judging by mortality alone.

Total Epidemic Morbidity (Influenza Incidence)

GENERAL ASPECTS

The observations made during the surveys relate to 146,203 persons, 42,920 cases, and 730 deaths. In view of the fact that the record of the morbidity from influenza practically disappears between epidemics and is extremely incomplete during epidemics, special significance must attach to the results of such a canvass. Although the data can not in themselves give an accurate picture of the incidence of the disease or of its case fatality in diverse parts of the country, they do indicate the incidence and fatality for the samples surveyed

and thus—in view of the correlation chronologically with the more general records noted in the introduction—for the particular cities in which the surveys were made. Accordingly, they serve as a check upon the precision of other morbidity data, and indicate in a general way certain highly important relations between morbidity and mortality.

The general incidence of influenza ("total epidemic morbidity") in the areas canvassed will be the first point to be taken up. In a later section of the report it will be shown that this incidence was not greatly different in the white and colored population. Because of this fact, and because of the small proportion of colored in most of the localities, no considerable error will be introduced into the following discussion by combining the white and colored rates.

Table 3.—Incidence of influenza in canvassed populations of each surveyed locality during the epidemic of 1918-19

Locality	Rate per 1,000	Num- ber of cases	Num- ber of persons	Locality	Rate per 1,000	Num- ber of cases	Num- ber of persons
All localities	294	42, 920	146, 203	Baltimore, Md Des Moines, Iowa	246 231	8, 199 1, 353	33, 361 5, 857
San Antonio, Tex. Minor Maryland towns Charles County, Md Little Rock, Ark Augusta, Ga	535 405 405 359 341	6, 701 5, 060 6, 546 3, 565 1, 405	12, 534 12, 482 16, 147 9, 920 4, 123	San Francisco, Calif Spartanburg, S. C. Macon, Ga. New London, Conn. Louisville, Ky. ¹	215 214 213 185 150	4, 021 1, 126 1, 681 1, 466 1, 797	18, 682 5, 257 7, 905 7, 933 12, 002

¹ Survey made before epidemic had ended.

The rate for all localities is 294 per thousand persons. words, one out of every three or four persons in the canvassed populations reported that they had influenza during the autumn wave of the epidemic and the recurrence. Other studies made by the same method in various parts of the country give substantially the same results, and a tabulation of these studies by Jordan is of interest at this point.

Table 4.—Incidence of influenza (autumn wave, 1918) in canvassed populations of various United States communities 1

Locality	Rate per 1,000	Num- ber of cases	Number of persons canvassed	Locality	Rate per 1,000	ber of	Number of persons canvassed
Oswego, N. Y. ²	470 406 289	6, 094 4, 749 3, 845	12, 952 11, 686 13, 319	Watertown, N. Y. ²	282 245 234	5, 765 2, 930 645	20, 473 11, 969 2, 757

It is of interest to contrast these results with those for the Army, remembering that in the latter case the population is concentrated at

From Epidemic Influenza, by E. O. Jordan, p. 190.
 Some Statistics of Influenza in Oswego and Watertown in 1918. Official Bull. N. Y. State Department

of Health, 4:53.

Report of Bureau of Local Health Administration. State Department of Health of New Jersey, 42:28.

Statistics of the 1918 epidemic of influenza in Connecticut. Winslow, C.-E. A., and Rogers, J. F. Journ.

those ages when the incidence was particularly high. The rates for four months of 1918 (September-December), corresponding approximately to the period covered by the Public Health Service surveys, are given in Table 5. The rates are for hospital admissions for influenza, bronchitis, broncho-pneumonia, and lobar pneumonia combined, and are exclusive of sickness occurring among the troops in Europe.

Table 5.—Incidence of total respiratory diseases in Army in the United States, (admissions) September to December, inclusive, 1918 2

1

With this picture before us, we are able to establish in a broad way what the incidence of influenza was during the 1918 epidemic, and the results secured in the surveys by the Public Health Service seem to give a rather representative mean.

Detailed house-to-house surveys in England, comparable to the canvass by the Public Health Service, were made in a number of towns for the summer and autumn waves of 1918, giving considerably lower rates than those indicated for this country. Table 6 summarizes these results (also from Jordan).

Table 6.—Comparison of influenza incidence rates per 1,000 in English towns 1918 1

Locality	Summer	Autumn	Total	Persons
Manchester ² Leicester ³ Cambridge ⁴ Warrington ⁵ Newcastle-upon-Tyne ⁶	149 63 36 75 62	103 146 165 82 47	252 209 201 157 109	4, 666 4, 619 1, 626 4, 461

Returning again to the canvass made by the Public Health Service, it will be noted that the highest rate was in San Antonio, where one out of every two persons reported having the disease. The range of variation in the rates is considerable, the rate in San Antonio being nearly three times that in New London. The canvassed populations are so large that only a relatively small part of this fluctuation can be

Influenza, bronchitis, broncho-pneumonia, and lobar pneumonia.
 Compiled from data given in the Medical Department of the United States Army in the World War.
 Vol. IX. Communicable and Other Diseases. Prepared by Lieut. Col. Joseph F. Siler. Chapter 2: Inflammatory Diseases of the Respiratory Tract, by Maj. Milton W. Hall.

¹ From Epidemic Influenza, by E. O. Jordan, p. 194.

2 Analysis of the results of a block census undertaken in Manchester in December, 1918. 1920. Ministry of Health. Report on the Pandemic of Influenza, 1918-19. London, P. 456. By T. Carnwath.

3 Report on an inquiry into the recent epidemic of influenza in the county borough of Leicester. 1920. Ministry of Health. Report on the Pandemic of Influenza, 1918-19. London, P. 445. By M. B. Arnold.

4 Report on incidence of influenza in the University and Borough of Cambridge, and in the Friends School, Saffron Walden. 1920. Ministry of Health. Report on the Pandemic of Influenza, 1918-19. London, P. 388. By S. M. Copeman.

5 Report on an investigation of the incidence and effects of influenza among the population of Warrington (Lancs). 1920. Ministry of Health. Report on the Pandemic of Influenza, 1918-19. London, P. 530.

⁽Lanes.). 1920. Ministry of Health. Report on the Pandemic of Influenza, 1918-19. London. P. 539. By G. W. N. Joseph. Analysis of an influenza census at Newcastle-upon-Tyne. 1920. Ministry of Health. Report on the Pandemic of Influenza, 1918-19. London. P. 556. By S. J. Clegg.

explained as being due to chance. However, in several widely separated localities the incidence rate varied only within narrow limits.

A cursory examination of the rates in the different localities will show that no consistent relation is manifested between the rates and the geographic position of the localities. If the New England and Maryland localities are grouped together and contrasted with the central and southern localities, the rates in the two groups will be found to be practically identical, namely, 304 and 306, respectively.

AGE

A marked selective effect on the incidence of influenza was exerted by age during the epidemic of 1918-19. This observation, which is common to nearly all reports on the epidemic, is corroborated by the data secured in the surveys. What they show most clearly is a very heavy incidence in the younger ages and a definite contrast with the curve of mortality.

The influenza morbidity rates for each 5-year age group for all surveyed localities are given in Table 7.

Table 7.—Incidence of influenza among canvassed persons in each age group in all surveyed localities during the epidemic of 1918-19

Age group	Rate per	Number of	Number of
	1,000	cases	persons
All ages	294	1 42, 920	² 146, 20 3
Under 1	207	586	2, 838
	337	4, 016	11, 933
Under 5	312	4, 602	14, 771
	391	5, 755	14, 725
	381	5, 404	14, 182
	345	4, 448	12, 897
20-24	323	3, 967	12, 287
25-29	337	4, 127	12, 234
30-34	326	3, 805	11, 668
35-39	2 96	3, 276	11, 074
40-44	236	2, 219	9, 415
45-49	207	1, 688	8, 157
50-54	175	1, 162	6, 628
55-59	162	698	4, 323
60-64	143	537	3, 756
65-69	135	332	2, 456
70-74	111	189	1, 703
75 and over	88	145	1, 650

¹ Includes 566 of unknown age.

It will be noted that the incidence was highest in the age group 5 to 9, fell off progressively in the age groups from 10 to 24, rose to

$$0.6745\sqrt{\frac{pq}{n}}$$
 or $\sqrt{\frac{(\text{rate}) (1000-\text{rate})}{n}}$

where p is the chance that an individual will have a case, q the chance that he will not, and n the size of the canvassed population. The probable error is applicable because there were relatively few instances where one person reported having more than one case.

² Includes 4,277 of unknown age.

⁵ Even in the case of New London, which has one of the smallest surveyed populations, the probable error of the rate is less than 7 per 1,000 persons. This calculation is based on the formula

a minor second mode in the age group 25 to 29, and then declined progressively in successive age groups. Among old people the incidence appeared to be not more than one-third of that among the young.

Through the courtesy of the health officers of the States of Kansas and Maryland, reports of cases of influenza in these States were available for statistical analysis. Without going into the results of these studies in any detail, a comparison by age is of interest for corroborative purposes. There was, of course, no expectation that any great proportion of the cases occurring would be reported to the health departments of the States, but it was felt that the relative incidence by age might not be greatly affected by this limitation. In order to permit a comparison between the surveyed data and the data for the two States, the rates have been reduced to an index basis by dividing by the rate for all ages. Thus the three curves are put on a relative basis, and the actual height becomes of no significance.

Table 8.—Relative incidence of influenza by age in surveyed localities, in Kansas, and in Maryland during epidemic of 1918-19 (rate for each age group divided by rate for all ages)

Age group	Surveyed localities	Kan- sas	Mary- land	Age group	Surveyed localities	Kan- sas	Mary- land
Under 5	1. 04 1. 30 1. 27 1. 15 1. 08 1. 12 1. 09 . 99 . 79	0. 73 1. 28 1. 34 1. 34 1. 30 1. 36 1. 34 1. 09 . 88	0.73 1.29 1.36 1.47 1.40 1.29 1.22 .96	45-49. 50-54. 55-59. 60-64. 66-69. 70-74. 75 and over.	0. 69 . 58 . 54 . 48 . 45 . 37 . 29 1. 00	0.59 .34 } .19 1.00	0.58 { .40 .31 .24 } .20 1.00

The results are represented graphically in Figure 1. In general, the curves for Kansas and Maryland correspond to the curve for the survey, although the former show a tendency to fall off more rapidly with age. This may be due to a greater tendency not to report sickness among old people to the health authorities. At all events, it is the similarity of the three curves, rather than any differences, which is most striking.

The age curves in each of the surveyed localities may next be considered. These curves are given, in 5-year age groups, in Figure 2 and Table 9. For the graph, as in the preceding case, the ratios of the rate in each age group to that for all ages are used so that the age incidence in the different localities may be readily compared.

Table 9.—Incidence of influenza by age in each locality during epidemic of 1918-19
[Rate per 1.000]

Age group	New Lon- don	Balti- more	Minor Mary- land towns	Charles County, Md.	Sper- tan- burg	Au- gus- ta	Ma- con	Des Moines	Louis- ville	Little Rock	San An- to- nio	San Fran- cisco
Under 5	180	283	414	380	252	388	247	274	238	366	488	209
	230	366	493	448	253	480	318	350	268	463	609	281
	224	317	512	486	238	416	264	233	211	460	625	290
	177	289	493	508	232	325	219	220	142	384	598	235
	207	275	476	493	250	326	207	240	169	335	590	236
25-29	236	314	485	465	221	412	225	261	143	392	598	262
	210	295	488	441	217	388	202	249	188	378	590	258
	221	229	421	407	214	398	238	235	135	386	527	225
	173	185	321	349	168	242	196	219	113	262	464	185
45-49	169	158	300	277	158	298	142	162	93	278	410	157
	121	135	266	255	152	284	132	138	84	213	379	121
	72	131	211	229	124	298	160	161	64	199	330	97
	108	124	183	211	130	247	89	140	66	222	234	80
65-69	103	112	201	181	75	283	63	125	83	136	294	72
70-74	74	79	145	147	132	214	67	138	51	211	247	78
75 and over	20	56	109	119	150	275	45	36	49	236	230	33

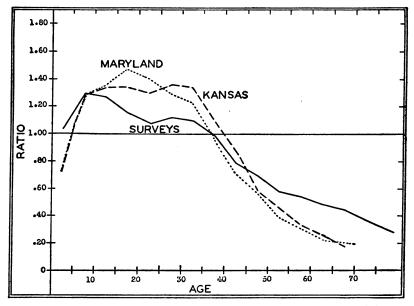


FIGURE 1.—Relative incidence of influenza by age in surveyed localities, in Kansas, and in Maryland, during 1918-19 epidemic (ratio of rate in each age group to that in all ages)

Although minor differences are noted in the incidence in various age groups, the essential similarity in the different localities—if we neglect the actual level of the rates already considered—is much more striking than these slight differences, indicating quite conclusively that the selective incidence in relation to age was a marked characteristic of this epidemic in each locality. The peak in the

younger ages, with a gradual decline in the rates after age 30 or 35, is found in every locality.

Perhaps of greatest interest is the suggestion that the double peak indicated in the data for all surveyed localities and in the reported morbidity for Kansas is really significant. The only curve which

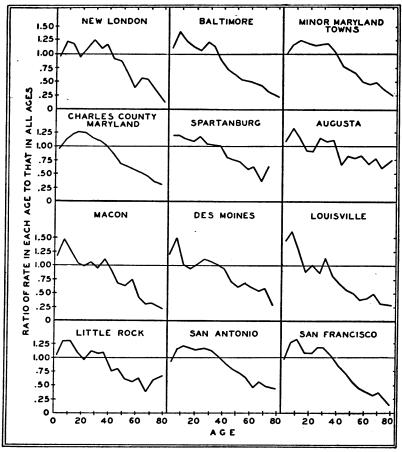


FIGURE 2.—Ratio of influenza case incidence in each age group to that in all ages in a canvassed population of each surveyed locality

does not give a suggestion of the two peaks is that for a rural area (Charles County). The first peak usually occurs in the age group 5 to 9 and the second peak in the age group 25 to 29. This bimodal tendency is analyzed in Table 10, giving the age group in which the two modes occur in each locality.

Table 10.—Age groups when first and second modes occur in each surveyed locality during epidemic of 1918-19

	Age group when—				
Locality	First mode occurs	Second mode occurs			
Spartanburg Baltimore. New London. Augusta. Des Moines. Little Rock Louisville. Macon. San Antonio. San Francisco. Minor Maryland towns.	1 0-9 5- 9 5- 9 5- 9 5- 9 5- 9 5- 9 10-14 10-14	20-24 25-29 25-29 25-29 25-29 25-29 30-34 35-39 25-29 25-29 30-34			

¹ Same rate for 0-4 and for 5-9.

In practically every case the second mode is quite definite, but it should be pointed out that in only one locality (New London) is the second mode higher than the first.

Reference may be made to the fact that W. T. Vaughan, in a house-to-house survey of 10,000 persons in Boston, also found two peaks of age incidence.

Question arose as to the advisability of adjustment of the rates for influenza in the various surveyed localities to a standard age or age and sex distribution of the population. Such adjustments were worked out, but found to be too slight in their effect to warrant their use in this paper, except for certain comparisons between the sexes.⁶

INCIDENCE OF INFLUENZA IN THE TWO SEXES

The morbidity rate of influenza as obtained in these canvasses was slightly higher for women than for men, the rate for all localities being 307 and 294, respectively, after adjustment to a standard age

⁶ To bring out the rather slight effect of adjustment for age and sex, the following table is reproduced. The rates for the different localities differ somewhat from those used previously, because in this case it was necessary to base the rate on persons of known ages.

	Actual	Rate per 1,000 adjusted to standard population (all surveyed localities)					
Locality	rate per 1,000 known	Ву	age	By age and sex			
	ages	Rate	Ratio to actual	Rate	Ratio to actual		
All localities	298	301	1.01	300	1. 01		
San Antonio, Tex		525	. 98	522	. 97		
Minor Maryland towns Charles County, Md.		418 405	1. 02 1. 00	417 405	1. 02 1. 00		
Little Rock, Ark	360	356	.99	354	. 98		
Augusta, Ga		362	1.01	359	1.00		
Baltimore, Md	253	260	1.03	258	1.02		
Des Moines, Iowa	232	235	1.01	233	1.00		
Spartanburg, S. C.	217	214	. 99	212	. 98		
San Francisco, Calif		219	1. 01 1. 01	218 212	1. 01 1. 00		
Macon, Ga		216 189	1.01	188	1.00		
New London, Conn Louisville, Ky	158	165	1.04	165	1.04		

distribution. The rates for women were higher in nearly every locality. The differences are brought out in Table 11. Adjustment seemed advisable, because of the possible effect of the withdrawal of males for military duty. As a matter of fact, this adjustment made little difference in the ratio between the two sexes, the unadjusted rates being 304 and 292 for women and for men for all known ages and 299 and 288 for all ages.

Table 11.—Incidence of influenza by sex in each surveyed locality (adjusted to standard age distribution) during epidemic of 1918-19

T 114	Locality Rate p		Ratio of female		Rate 1	per 1,000	Ratio of female
Locality	Male	Female	rate to male	Locality	Male	Female	rate to male
All localities	294	307	1. 04	San Francisco	213 514	222 530	1. 04 1. 03
Macon Minor Maryland towns	194 406	229 459	1. 18 1. 13	Augusta Charles County, Md	357 403	364 406	1. 02 1. 01
Spartanburg	200	220	1. 10	Des Moines	229	231	1.01
Baltimore New London	248 185	270 192	1. 09 1. 04	LouisvilleLittle Rock	166 352	164 345	. 99

When it is realized that in a large proportion of families the information was secured from the wife, it seems possible that this slight excess for women might be due to the fact that they were able to remember their own cases somewhat better than the cases of other members of the family. A tendency of this character has been noted in other studies where the information was secured in this manner.⁷ Thus the only conclusion which is really justified is that there was no marked difference in the rates of the two sexes.

In Table 12 and Figure 3 comparison is made by sex for the different ages.

Table 12.—Incidence of influenza among canvassed males and females in each age group, in all surveyed localities during epidemic of 1918-19

	Rate p	er 1,000	Number	of cases	Number	of persons
Age group	Male	Female	Male	Female	Male	Female
All ages	288	299	19, 742	23, 169	68, 684	77, 495
Under 1	214	199	301	284	1, 407	1, 427
	348	325	2, 081	1, 933	5, 984	5, 945
Under 5	322	301	2, 382	2, 217	7, 391	7, 372
	388	394	2, 845	2, 910	7, 342	7, 382
	379	383	2, 649	2, 755	6, 994	7, 187
	332	356	1, 985	2, 461	5, 986	6, 909
20-24	288	343	1, 267	2, 699	4, 405	7, 881
25-29	328	344	1, 624	2, 503	4, 953	7, 281
30-34	320	331	1, 723	2, 082	5, 385	6, 283
35-30	295	296	1, 638	1, 638	5, 546	5, 527
40-44	242	230	1, 112	1, 107	4, 592	4, 823
45-49.	200	215	850	838	4, 250	3, 907
50-54.	167	184	555	607	3, 319	3, 308
55-59.	157	166	334	363	2, 130	2, 192
60-64	128	157	237	300	1, 848	1, 908
	132	138	154	178	1, 170	1, 286
	114	108	85	104	744	959
	83	92	58	87	702	948

⁷ The Illness Rate Among Males and Females. By E. Sydenstricker. Pub. Health Rep., vol. 42, No. 30, July 29, 1927. (Reprint 1172.)

Except for the youngest ages, there is a tendency for the female rates to be higher, but, as just pointed out, the difference is slight. The ratios of female rates to male rates for broad age groups are as follows: Under 15 years, 0.99; 15 to 44, 1.07; 45 to 59, 1.07; and 60 and over, 1.09. The age curves are practically identical in the two sexes, the only difference between the two being the greater depression in the male curve between the two modes. In fact, the female curve shows only a bare suggestion of the second mode.

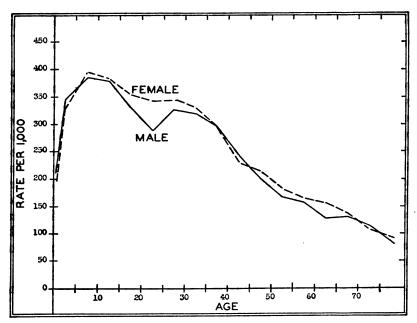


FIGURE 3.—Incidence of influenza among canvassed males and females in each age group (all surveyed localities)

INCIDENCE IN COLORED POPULATION

Since a number of the cities had a considerable colored population, it is of interest to determine whether a larger percentage of white or colored were attacked. We are faced immediately with the difficulty of getting as complete information from the colored as from the white in a canvass of this character; thus any results must be discounted. No rates have been used for all the surveyed localities, because of the varying proportion of colored persons in the different localities. The rates in the eight places where there was a sufficient number of colored to give somewhat reliable results are given in Table 13, adjustment having been made to a standard age and sex distribution.

Table 13.—Incidence of influenza in white and colored canvassed populations during the epidemic of 1918-19 (adjusted to a standard age and sex distribution)

* . 114	Rate p	er 1,000	Ratio of	Number of cases		Number of persons	
Locality	White	Colored	colored to white	White	Colored	White	Colored
Louisville	179	49	0, 27	1, 739	58	10, 534	1, 468
Baltimore	278	116	.42	7, 690	481	29, 085	4, 198
Augusta	456	212	. 47	1,044	361	2, 434	1, 689
Macon	220	137	. 62	1, 337	341	5, 971	1, 930
Spartanburg	224	173	. 77	1, 033	84	4,652	581
Minor Maryland towns	419	385	. 92	4, 794	249	11, 782	643
Little Rock	360	338	.94	2, 657	908	7, 262	2, 654
Charles County, Md	1 379	1 431	1. 14	3, 028	3, 518	7, 992	8, 155

¹ Rates for Charles County unadjusted; adjustment made only a slight difference in the ratios.

With the exception of Charles County, Md. (see p. 304 for information as to method of survey in this locality), the rates are consistently lower for the colored populations. In Louisville, Baltimore, and Augusta the rate is at least twice as great in the white as in the colored population. The fact that the colored population live generally under conditions presumably more favorable to the spread of contact infections would lead one to expect a higher rate of influenza among them. How much of the difference is to be ascribed to more complete reporting among the white populations is quite impossible to determine. Some confirmation of this difference between the incidence of influenza in white and colored is given by the rates for the Army while in the United States. The period covered in the table is September-December, 1918.

Table 14.—Incidence of total respiratory 1 disease by color in Army in the United States, September-December, inclusive, 1918 2

	White	Colored
Rate per 1,000	316	269
Number of cases	383, 498	40, 576
"Strength"	1, 215, 447	150, 569

¹ Influenza, bronchitis, broncho-pneumonia, lobar pneumonia.

² Compiled from data given in the Medical Department of the United States Army in the World War.

Vol. IX. Communicable and Other Diseases. Prepared by Lieut. Col. Joseph F. Siler. Chap. 2: Inflammatory Diseases of the Respiratory Tract, by Maj. Milton W. Hall.

One further table is presented giving the incidence of influenza by color in the two sexes. The tendency for higher rates in the white population is evidently present in both sexes.

Table 15.—Incidence of influenza by sex and color in certain canvassed localities during epidemic of 1918-19

	Rate per 1,000						
Locality	М	ale	Fei	nale			
	White	Colored	White	Colored			
Louisville Baltimore	169	43	162	37			
	255	98	272	129			
Augusta	427	197	430	225			
Spartanburg	207	135	235	152			
Maçon	198	171	246	180			
Little Rock Minor Maryland towns Charles County, Md	377	308	355	371			
	397	330	415	433			
	383	419	374	444			

The Frequency of Pneumonia as a Complication

GENERAL ASPECTS

The 1918-19 epidemic of influenza was notably different from the 1889-90 epidemic in a much higher frequency of pneumonia and consequently a much higher mortality, especially among young adults. The record of pneumonia cases in the areas canvassed by the Public Health Service is therefore of interest, particularly in view of the inadequacy of pneumonia morbidity reports during either epidemic or normal periods. As noted in the introduction, cases were classified in these surveys as "pneumonia" when so reported by the householder. No attempt could be made to diagnose the cases or to inquire of the physician in charge as to the diagnosis made by him. Deaths from influenza were classed as pneumonia cases even when not so specified on the census report.

The results obtained in Charles County are evidently not comparable to those obtained in the other localities, since in this county there were only 102 pneumonia cases recorded, whereas there were 147 deaths from influenza-pneumonia. The deaths in this instance were presumably complete, as the results of the survey were checked up with the death certificates in the State registrar's office; but since it may be assumed that epidemic deaths were due almost always to complicating pneumonia, and since by no means all of the pneumonia cases resulted in death, clearly the pneumonia cases were not complete. Because of these obvious inconsistencies, the records from Charles County have been omitted from all discussions of pneumonia morbidity.

The following table gives the pneumonia incidence for all localities (except Charles County) and for each locality.

⁸ If we were to assume completeness of recording nonfatal cases of pneumonia, we would have a fatality rate in Charles County of 82 per cent, whereas in the other localities the average is about 25 per cent.

Table 16.—Incidence of pneumonia in canvassed population of each surveyed locality during epidemic of 1918–19

Locality	Rate per 1,000	Num- ber of cases	Number of persons	Locality	Rate per 1,000	Num- ber of cases	Number of persons
All localities 1 Minor Maryland towns San Antonio	17. 6 25. 8 24. 2	2, 290 322 303	130, 056 12, 482 12, 534	New LondonLittle RockAugusta	17. 1 16. 0 15. 3	136 159 63	7, 933 9, 920 4, 123
Des Moines	23. 6 18. 0 17. 2	188 599 321	5, 857 33, 361 18, 682	Macon Louisville Spartanburg	13. 0 9. 2 6. 7	103 111 35	7, 905 12, 002 5, 257

¹ Exclusive of Charles County, Md.

The pneumonia case rate for all localities (except Charles County) was 17.6 per 1,000 persons, as compared with 280, the influenza rate, for the same localities. In other words, the percentage of influenza cases complicated by pneumonia, as determined in these surveys, was 6.3. A more detailed comparison with influenza morbidity will be taken up later. At this point it is desirable to summarize the pneumonia data themselves.

The most striking feature of the pneumonia rates is their wide range. The minor Maryland towns have a rate four times as great as that of Spartanburg (surveyed population, 5,257) and nearly three times as high as Louisville (surveyed population, 12,002).

Another point of interest is that the cities with the lowest rates are invariably in the south central part of the country, where, it is believed, the epidemic was somewhat less severe. The combined pneumonia rate for Augusta, Macon, Louisville, and Spartanburg was 10.7, whereas it was 19.6 in the other localities combined.

AGE

The toll of the epidemic in young adult life is depicted clearly by the rates for cases of pneumonia recorded in these surveys. The pneumonia incidence in each age group for all localities is presented in Table 17. The numbers are evidently sufficient for quite reliable results.

TABLE 17.—Incidence of pneumonia by age in all localities, exclusive of Charles County, Md., during epidemic of 1918-19

Rate per 1,000	Number of cases
17. 6	2, 290
24. 9 26. 0	60 264
25. 8 14. 8	324 186
15. 5 23. 1	137 173 256
25.7	352 279 213
13. 0 9. 8	112 73
9.3 6.5	82 51 19
	1,000 17. 6 24. 9 26. 0 25. 8 14. 8 11. 5 16. 5 23. 1 31. 1 25. 7 21. 0 9. 8 8. 3 9. 3

There are two marked peaks. The incidence is high in children under 5 years of age, although not any higher in the first year of life than in the years immediately following. The second mode occurs in young adult life, the highest point being found in the age group 25 to 29, where the rate is three times that in the age group 10 to 14. As age advances, the rate falls off rapidly. By 50 years of age it is already one-half of the rate for the age group 25 to 29. A direct comparison with the incidence of influenza as a whole is postponed until later, but it may be pointed out that the bimodal effect noted in the case of influenza is much more marked in the case of pneumonia alone. In both the incidence falls off steadily with age after the second peak.

So striking is this bimodal tendency for pneumonia curves according to age during the epidemic that it seems well to present the rates by age for the individual localities. The numbers are limited, and it has been necessary to combine certain age groups. The data are given in Table 18 and Figure 4.

Table 18.—Incidence of pneumonia in each canvassed locality, by age, during epidemic of 1918-19 1

		Rate per 1,000 persons canvassed									
Age group	New London	Balti- more	Minor Mary- land towns	Spar- tan- burg	Augus- ta	Macon	Des Moines	Louis- ville	Little Rock	San An- tonio	San Fran- cisco
Under 5. 5-9. 10-14. 15-19. 20-24. 25-29. 30-34. 35-39. 40-44. 45-59. 60 and over.	11. 1 9. 0 7. 8 12. 4 25. 4 44. 4 28. 7 19. 0 10. 4 6. 9 5. 4	27. 3 13. 4 11. 3 18. 5 21. 1 29. 4 21. 8 18. 4 10. 7 7. 2 9. 4	38. 2 21. 6 15. 2 19. 3 37. 0 39. 7 46. 2 38. 8 14. 9 9. 3 3. 9	10. 4 3. 2 3. 8 4. 3 6. 0 9. 1 8. 8 4. 9 6. 1	29. 6 23. 1 7. 4 13. 2 18. 0 26. 9 26. 7 7. 4 { 9. 2 4. 5	17. 5 11. 7 7. 3 10. 4 11. 2 14. 8 14. 6 { 15. 7 16. 5 8. 3 2. 1	37. 4 50. 0 14. 2 18. 0 25. 7 38. 7 24. 7 16. 3 10. 4 5. 8}	22. 2 8. 5 9. 2 6. 1 9. 4 11. 0 11. 7 10. 6 4. 7 3. 6	16. 3 7. 8 10. 9 11. 7 24. 1 22. 1 24. 7 18. 3 9. 8 9. 5 6. 9	20. 4 14. 7 11. 2 20. 1 39. 8 42. 2 34. 4 31. 9 13. 7 17. 4 8. 7	20. 2 8. 7 10. 6 15. 2 14. 7 30. 7 22. 8 17. 7 10. 6 9. 0

Inclusion of deaths from influenza as pneumonia cases was not possible in this table, except where the case was originally recorded as pneumonia. The rates, however, are not more than about 7 per cent too low.

The marked bimodal effect is noted in each locality without any exception. In all but one city the first peak comes in the under 5-year age group. Usually the second peak is in the age group 25 to 29, but in three instances it is in the age group 30 to 34, and in one in the age group 35 to 39. It is evident that the location of these modes is subject to a certain chance variation.

This strikingly high incidence of pneumonia in the young adult population, reaching a peak of nearly 5 per cent in some of the localities in the modal age group, is obviously at great variance with the normal age distribution of pneumonia. An idea of this difference may be obtained from a comparison of the age curve secured in this canvass with that for Hagerstown, Md., during a period (December 1, 1921, to April 1, 1924) without major epidemic waves, the data

having been secured in house-to-house canvasses during this period by the Public Health Service. No comparison of the actual level of the morbidity rates seems feasible or of consequence in this connection, in view of the varying periods for which the sickness data

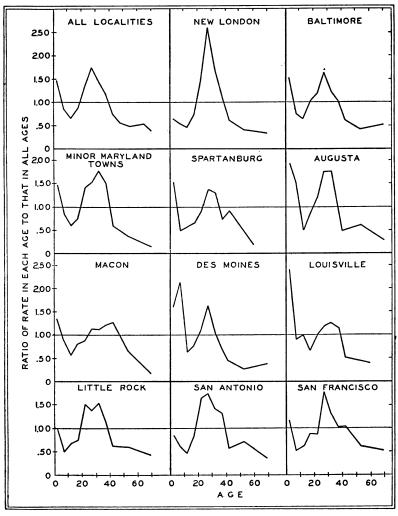


FIGURE 4.—Ratio of pneumonia case incidence in each age group to that in all ages in a canvassed population of each surveyed locality

in the various localities were secured. Comparison may be made most easily by reducing each series of rates to an index basis by dividing by the rate for all ages. These indices are given in Figure 5 and Table 19.

[•] The Incidence of Various Diseases according to Age. Hagerstown Morbidity Studies No. VIII. By Edgar Sydenstricker. Public Health Reports, May 11, 1923. (Per rint No. 1227.)

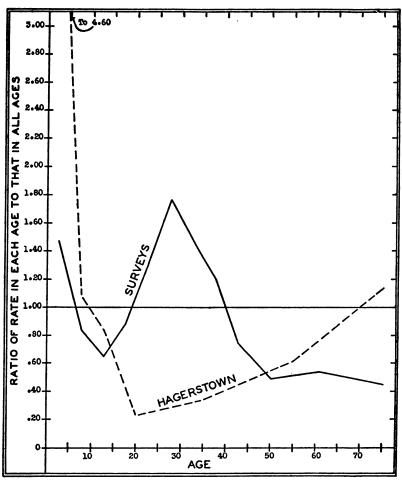


FIGURE 5.—Relative incidence of pneumonia by age in surveyed localities and in Hagerstown, Md. (data for Hagerstown from a previous sickness survey)

Table 19.—Relative incidence of pneumonia by age in surveyed localities during 1918-19 epidemic and in Hagerstown sickness study (rate for all ages=1.00)

	Rates p	er 1,000	Indices		
Age group	Surveyed localities	Hagers- town 1	Surveyed localities	Hagers- town	
Under 5	25. 8 14. 8 11. 5 15. 5 23. 1 31. 1 25. 7 21. 0 13. 0 8. 7 9. 5 7. 9	} 2.0	1. 47 .84 .65 { .88 1. 31 1. 77 1. 46 1. 19 .74 .49 .54 .45	4.60 1.09 .84 } .23 } .34 } .62 1.14 1.00	

¹ Annual rates.

In a nonepidemic period, pneumonia has its highest frequency at the beginning and end of life. In the pandemic of 1918 pneumonia showed its highest frequency in the age group 25 to 29, a subordinate peak in the age group under 5 years, and a relatively low incidence after 40 years of age. It should be observed that the contrast is really somewhat greater than that shown in the figure, since the curve for the epidemic contains a proportion of deaths from pneumonia not associated with the epidemic and therefore tending to follow the age curve as typified by the Hagerstown data.

SEX

In contradistinction to the material presented for the total morbidity during the epidemic, the pneumonia rates are slightly higher in the males, as shown in Table 20. Spartanburg is omitted, because only 35 cases were recorded in all, but is included in the total for all localities. The rates have been adjusted to a standard age distribution.

Table 20.—Incidence of pneumonia by sex in each surveyed locality during epidemic of 1918-19 \(^1\) (adjusted to standard age distribution)

f	Rate p	er 1,000	Ratio of female	Cases		
Locality	Male	Female	rate to male	Male	Female	
Augusta. Baltimore. New London Macon Louisville Minor Maryland towns Little Rock	16. 5 18. 7 16. 7 12. 9 11. 2 30. 6 17. 0 27. 0	17. 9 19. 1 16. 9 12. 5 9. 6 25. 1 13. 7	1. 08 1. 02 1. 01 . 97 . 86 . 82 . 81	26 267 64 44 52 159 81	35 327 72 56 58 163 77	
San Antonio	27. 0 20. 8 29. 4	21. 7 14. 8 19. 5	. 80 . 71 . 66	139 177 79	161 144 59	

¹ Spartanburg omitted because of small numbers.

In only one locality is the rate for females definitely higher. The fact that we do not find higher rates among females for these serious cases suggests that possibly the difference in the incidence of influenza as a whole was due to the tendency of the women to report a higher incidence for themselves than for other members of the family. That would hardly be expected in the case of illnesses severe enough to be classed as pneumonia, as they would probably be recalled whatever member of the family had the case.

A graph is added for pneumonia incidence by sex and age. (Fig. 6.) There is a suggestion that the excess among males occurs entirely during the ages where the epidemic exerted its greatest effect. The rates are presented in Table 21.

Table 21.—Incidence of pneumonia by sex and age in all surveyed localities during epidemic of 1918-19 (rate per 1,000)

Age group	Male	Female
All ages	18. 4	17. 0
Under 1	28. 8 27. 5 27. 7 14. 1 12. 1 17. 2 24. 1 29. 0 24. 1 11. 0 6. 2 6. 1 4. 2	21. 3 24. 5 23. 8 15. 5 11. 0 14. 2 22. 5 26. 9 22. 9 17. 9 12. 8 8. 6 10. 3 12. 2 8. 2

¹ Exclusive of Charles County.

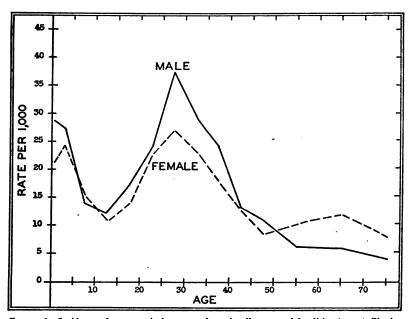


FIGURE 6.—Incidence of pneumonia by age and sex in all surveyed localities (except Charles County, Md.)

In view of these differences, it is of interest to compare the rates by sex and age in each locality. To do so, however, a broad grouping of ages is necessary to secure any degree of regularity. These broad groups have been chosen to bring out, as well as possible, the characteristics of the age curve (under 5 years, 5-19, 20-29, 30-39, 40 and over). The rates are given in Table 22. At the bottom of the table will be found ratios of the female rates to those of the males.

Table 22.—Incidence of pneumonia by sex and broad age groups in each surveyed locality during epidemic of 1918–19 ¹

	Under 5 years	5-19	20-29	30-39	40 and over
RATES PE	R 1,000				
All localities:					
Male	23. 3	12.2	26. 3	23.8	7.0
Female	20. 5	11.8	22. 3	18. 2	9.
New London:					
Male	11.3	12.6	34. 7	20.9	4.
Female	11.0	7. 2	34. 1	27. 3	9.
Báltimore:		10.0	07.0	۰.,	
Male	39. 5	13. 2	27. 9	19.4	6.
Female	24.0	15. 4	23. 3	20.6	10.
Minor Maryland towns:	00.0	20.7	48.4		
Male	36. 2 40. 1	16.9	31. 5	54. 5 32. 7	5. 12.
Female	40.1	10.9	31. 3	32.1	12.
Spartanburg:	3. 5	1.4	6.6	7.3	1
Male	17. 3	5.8	11.3	7.0	4.
Female	17.3	3.0	11. 3	1.0	7.
Augusta: Male	18.4	13.4	25, 4	15. 9	10.
Female	40. 2	15.8	20. 8	14.3	5.
Macon:	40. 2	10.0	20. 0	14. 0	1
Male	20.6	7.1	13.0	22.5	7.0
Female	14.8	11.9	13.0	8.2	10.
Des Moines:	14.0	11. 5	10.0	0. 2	10.
Male	49.1	37. 5	32.4	21.9	8.
Female	25. 4	17. 7	32.1	19. 2	6.
Louisville:	20. 2		U 1	-0.2	1
Male	19. 2	9.0	11. 5	15. 4	3.0
Female	25. 6	6.6	9.6	7.8	1 4.
Little Rock:		*. *			
Male	22.9	10.4	25. 0	26. 1	7.
Female	10.0	9.9	21.9	16. 5	9.
San Antonio:					
Male	26. 7	15. 3	44. 6	40.8	14.
Female	13.8	15. 5	40.6	26. 7	12.
San Francisco:					
Male	20.8	12.8	31. 1	24. 0	19.
Female	19.7	10.6	17. 6	16. 7	10.
RATIO OF FEMALE	RATE TO 3	AALE			1
All localities	88	97	85	76	12
New London	97	57	98	131	19
Baltimore	79	117	84	106	17
Minor Maryland towns	111	82	65	60	21
Spartanburg					l
Augusta	218	118	82	90	5
Macon	72	168	100	36	14
Des Moines	52	47	99 أ	88	8
Louisville.	133	73	84	51	11
Little Rock	44	95	88	63	12
San Antonio	52	101	91	65	8
San Francisco	95	83	57	70	5

 $^{^1}$ Inclusion of deaths from influenza as pneumonia cases was not possible in this table, except where the case was originally recorded as pneumonia.

The tendency is toward an excess in the male rate at the ages 20 to 39 and is evidently present in a great proportion of the localities.

COLOR

The recorded pneumonia incidence was generally greater among the white than among the colored population. The following table gives the cases and rates by color for each locality in which there was a considerable number of colored (except Charles County).

Locality		onia rate 1,900	Ratio of colored rate to	Number of cases		Number of persons	
	White	Colored	white	White	Colored	White	Colored
Louisville, Ky	10. 1 19. 7 19. 1 13. 7 26. 2 16. 9 6. 9	2. 7 8. 8 9. 3 10. 9 18. 7 13. 6 5. 2	0. 27 . 45 . 48 . 80 . 71 . 80 . 75	107 48 556 82 309 123 32	4 15 39 21 12 36 3	10, 534 2, 434 29, 085 5, 971 11, 782 7, 262 4, 652	1, 465 1, 689 4, 195 1, 930 643 2, 654 581

Table 23.—Incidence of pneumonia in canvassed white and colored populations of certain surveyed localities during epidemic of 1918-19

In some localities the colored population seemed almost to escape the disease, while the white population was severely affected. In Baltimore the white and colored rates were, respectively, 19.1 and 9.3, and in Louisville 10.1 and 2.7. This relation is consistent with the fact that, in the canvassed populations, the mortality was slightly higher in the white than in the colored.

Mortality and Case Fatality

Rates of mortality in the general population of this country during the pandemic of 1918 have been thoroughly analyzed. There is no occasion to refer to them in the present paper, or to utilize the record of deaths obtained in the canvass to corroborate such findings. The value of these records lies rather in the fact that by means of them we may have a fairly precise conception of the case fatality of the 1918 epidemic in the communities surveyed. The section will deal with the case fatality of the epidemic as a whole (the percentage which the influenza-pneumonia deaths are of the influenza cases) and the case fatality of pneumonia (the percentage which these deaths are of the pneumonia cases), together with some reference to the mortality rates themselves.

It has been previously pointed out that it is impossible to distinguish between deaths reported as due to influenza and those reported as due to pneumonia—in practically all cases both of these diseases contributed to the deaths. Therefore only a slight error will be introduced in taking the relation between the influenza-pneumonia deaths and the total epidemic or pneumonia cases. Obviously these deaths also include a small number of normal or nonepidemic deaths. In view of the small size of the samples and the lack of information as to the normal rate of pneumonia in these sample areas, it has been impossible to limit the study to epidemic deaths alone.

In the six communities in which a comparison was possible, it was found that the influenza-pneumonia death rate in the canvassed population was only about 70 per cent of that in the city as a whole during the same period. The discrepancy was found consistently in each

community, varying from 57 per cent in Louisville to 84 per cent in Baltimore. The data are recorded in Table 24, which gives also the mortality rates in the surveyed areas of the localities for which mortality rates for the whole city were not determined.

Table 24.—Mortality from influenza-pneumonia during epidemic period in total populations of certain surveyed localities and in canvassed populations of same localities

Locality	Middle date of survey	Estimated popula- tion	Death rates per 1,000 based on reported deaths in total pop- ulation	Deaths reported from September 1 to middle date of survey	Death rates per 1,000 canvassed persons	Ratio of rate for canvassed population to that for total pop- ulation
Baltimore Cumberland Augusta Louisville Little Rock San Francisco New London Minor Maryland towns 2 Charles County, Md Spartanburg Macon Des Moines San Antonio	Jan. 151 Dec. 3 Feb. 4 Dec. 16 Jan. 3 Feb. 151 Dec. 10 Dec. 5 Mar. 12 Dec. 18 Dec. 9 Feb. 4 Dec. 14	680, 000 27, 300 55, 000 245, 000 65, 000 475, 000 25, 190 26, 190 18, 326 22, 500 115, 000 150, 000	6.2 10.8 6.3 3.7 5.1 7.8	4, 239 29.5 348 908 330 3, 700	5. 2 7. 1 4. 4 2. 1 3. 9 4. 8 5. 8 6. 4 9. 1 1. 9 3. 2 3. 8 4. 2	0. 84 . 66 . 70 . 57 . 77 . 62

¹ Middle date of recanvass.

There are a number of factors which may tend to explain the lower mortality rates in the canvassed populations: (a) Deaths of nonresidents in hospitals in the city have a tendency to raise the city mortality rates, but would not appear in the canvassed population; (b) there might be a tendency for persons visited to fail to mention deaths occurring in the family some time previously; (c) canvassed populations naturally do not include certain groups of the population in which mortality rates are likely to be excessive, such as boarding houses. Whatever the cause of this discrepancy, it is manifest that the case fatality rates to be discussed are affected by it in some degree.

The case fatality for all localities (percentage of total cases which were fatal) was 1.70. If we consider the pneumonia cases alone, it was 25.5 (omitting Charles County). The data by locality are given in Table 25.

² Exclusive of Cumberland (given above).

TABLE 25.—Influenza and pneumonia case fatality in canvassed populations of each surveyed locality during epidemic of 1918-19

Totalites		rate per cases	Per cent of influ- enza com-	Number	Number	
Locality	Influenza	Pneumo- nia	plicated by pneu- monia	Influenza	Pneumo- nia .	of deaths
All localities	1. 70	1 25. 5	1 6. 8	42, 920	1 2, 290	730
New London	3. 14 2. 25	33. 8	9. 3	1, 466 6, 546	136	46 147
San Francisco Baltimore	2. 24 2. 10	28. 0 28. 7	8. 0 7. 3	4, 021 8, 199	321 599	90 172
Minor Maryland towns Des Moines	1. 63	26. 1 15. 9	6. 4 10. 2	5, 060 1, 353	322 138	84 22
Macon Louisville Augusta	1. 39	24. 3 22. 5 28. 6	6. 1 6. 2 4. 5	1, 681 1, 797 1, 405	103 111 63	25 25 18
Little Rock Spartanburg	1.09	24. 5 28. 6	4. 5 3. 1	3, 565 1, 126	159 35	39 10
Sân Antonio	. 78	17. 2	4.5	6, 701	303	52

¹ Exclusive of Charles County, Md.

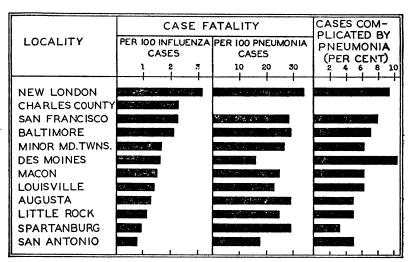


FIGURE 7.—Case fatality of influenza and of pneumonia, with percentage of cases complicated by pneumonia, in specified localities

A great variation in the fatality rates is observable, which is no doubt partly due to the small number of deaths. For total influenza, the fatality varies from 3.14 per cent in New London to 0.78 per cent in San Antonio. The coefficient of variability is 37.10 The pneumonia fatality showed much less variation, the coefficient being 23. The highest rate was in New London (33.8) and the lowest in Des Moines (15.9). These fatality rates are presented by graph in Figure 7, together with the percentage of cases complicated by pneumonia.

¹⁰ In making this calculation the minor Maryland towns were subdivided. See p. 305. Coefficient of variability is the standard deviation times 100 divided by the mean.

Examination of the graph shows that the influenza case fatality seemed somewhat lower in the south central part of the country. A map has been included (fig. 8) to bring this out more clearly. The fatality rate is indicated by symbols of varying degrees of density.

A comparison of the influenza case fatality with that obtained in certain other house-to-house canvasses is next given (Table 26). Since the available data are for the northeast section of the country, the only rates from the Public Health Service surveys which have been included in the table are for New London, Baltimore, and the minor Maryland towns.

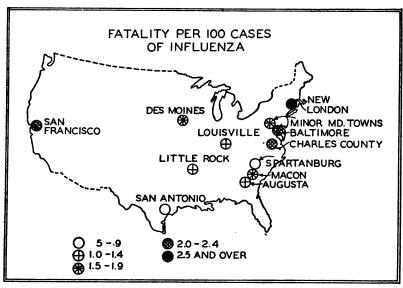


FIGURE 8.—Case fatality in different cities

Table 26.—Influenza case fatality rates during pandemic of 1918 in certain house-to-house canvasses

Locality	Case fatality (per- centage)	Number of persons surveyed	Locality	Case fatality (per- centage)	Number of persons surveyed
U. S. Public Health Surveys: New London Baltimore. Minor Maryland towns	3. 1 2. 1 1. 7	7, 933 33, 361 12, 482	New Britain, Conn. 1. Watertown, N. Y. 2. Boston 3. Oswego, N. Y. 3.	3.9 3.1 2.5 2.4	2, 757 29, 473 10, 050 12, 952

Statistics of the 1918 Epidemic of Influenza in Connecticut.
 Journ. Infec. Dis., 26:185.
 Winslow, C.-E. A., and Rogers, J. F.
 Some Statistics of Influenza in Oswego and Watertown in 1918-19.
 Off. Bull. N. Y. State Dept. of Health, 4:53.
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A question arises as to whether the incidence of influenza or the incidence of pneumonia determined the mortality rates in the sur-

veyed communities. This question can be considered from several angles. For instance, the fact that the case fatality of pneumonia was less variable than that of the epidemic as a whole (as previously noted) suggests that it was the presence of the secondary invaders which primarily determined the mortality. Another point of view is to consider the correlation of the rates of influenza, pneumonia, and deaths. The highest correlation is between the incidence of pneumonia and the mortality rates, but there is a definite correlation in the

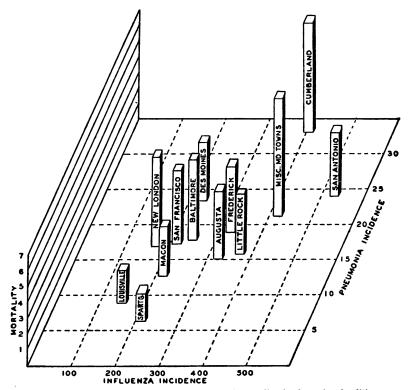


FIGURE 9.—Composite picture of incidence and mortality in the various localities

other two instances. The coefficients are as follows: Influenza incidence-pneumonia incidence, +.63; influenza incidence-mortality, +.66; pneumonia incidence-mortality, +.77. A composite picture of the interrelations by locality is shown in Figure 9, where the height of the vertical bars represents the mortality rates.

AGE

The age curve of mortality from the epidemic is given for all localities in Table 27, first for both sexes combined and then for males and females separately.

TABLE 27.—Mortality per 1,0	000 persons from	influenza-pneumonio	s by sex and age
TABLE 27.—Mortality per 1,0 in all surveyed	l localities during	gepidemic of 1918–19	1

Age group	Both sexes	Male	Female	Age group	Both sexes	Male	Female
All ages	5. 0	5. 3	4.7	20-24	6. 2 9. 9	6. 8 13. 3	5. 8 7. 6
Under 11-4	15. 2 6. 2	17. 1 5. 4	13. 3 7. 1	30-34 35-39 40-44	7. 9 6. 3 4. 0	9. 1 7. 9 4. 1	6.8 4.7 3.9
Under 5	7. 9 2. 2	7. 6 1. 9	8.3 2.4	45-49	2.9 2.6	3. 5 2. 8	2.3 2.4
10-14 15-19	2. 1 3. 4	1. 4 4. 0	2.8 2.9	70 and over	4. 3 5. 1	3. 3 4. 2	5. 6 5. 8

The most obvious point to be brought out is the extraordinary age curve of mortality during the epidemic. There is no necessity of emphasizing this fact here, since it has been thoroughly recognized in all accounts of the 1918–19 epidemic and the contrast with the usual experience has been apparent to everyone.

Discussion of the differences between the two sexes will be postponed until later. (See p. 334.)

The fatality of the epidemic according to age is of extraordinary interest, because it brings out so clearly the severe toll among young adults. The rates are presented in Table 28, for both the case fatality of influenza and that of pneumonia alone.

Table 28.—Fatality of influenza and of pneumonia by age, in all surveyed localities during epidemic of 1918–19 (percentage of cases which died)

Age group	Influenza	Pneu- monia ¹	Age group	Influenza	Pneu- monia ¹
All ages	1.7	25. 5	20-24	1.9	25. 0
Under 1	7. 4 1. 8	43. 3 18. 6	25-29 30-34 35-39	2.9 2.4 2.1	30. 1 28. 0 28. 6
Under 5	2.5	23. 1	40-44 45-49	1.7 1.4	28.6 27.4
5-9 10-14	0. 6 0. 6	11. 8 16. 1	50-59 60-69	1. 5 3. 1	28. 0 45. 1
15-19	1.0	19. 1	70 and over	. 5.1	57. 9

¹ Exclusive of Charles County, Md.

The very high incidence of pneumonia in young adult ages (previously discussed) is evidently the most important factor in the determination of the curves shown herewith. The fatality of influenza rises to nearly 3 per cent in the age group 25 to 29 and then falls to less than 1.5 per cent. In old age it rises again, reaching 5 per cent or more. Pneumonia cases themselves do not show this striking change in fatality in young adult life. As a matter of fact, the pneumonia fatality curve, except for an expected high value at the beginning of life, rises rather consistently from 12 per cent in the age group 5 to 9 to nearly 60 per cent in old age. It must again be stressed that the picture of pneumonia fatality includes the cases and deaths

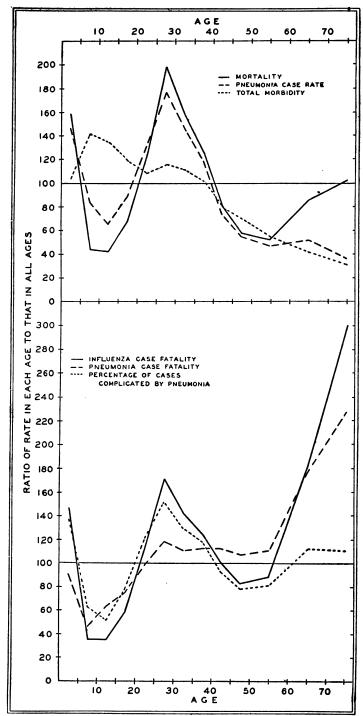


FIGURE 10.—Epidemic relations, by age, on relative basis (all ages=1.00). (Charles County omitted in rates involving pneumonia incidence)

which would have occurred at this time of year quite apart from the epidemic.

Perhaps it would be convenient to summarize in a single graph all the relations which have been brought out with respect to age, because the striking manner in which the epidemic affected young adults is so clearly depicted. Figure 10, accordingly, gives the age curves for influenza incidence, pneumonia incidence, mortality, percentage which the pneumonia cases were of the influenza cases, case fatality of the epidemic as a whole and case fatality of pneumonia. The indices (ratio of the rate in each age group to that for all ages) are given in Table 29.

Table 29.—Ratio of rates in each age group to those in all ages in all canvassed localities during epidemic of 1918–19

A ge group	Influenza incidence		Per cent compli- cated by pneumo- nia 1	Mortality	Influenza case fatal- ity	
All ages Under 5. 5-9. 10-14. 15-19. 20-24. 27-29. 30-34. 35-39. 40-44. 45-49. 50-59. 00-69. 70 and over	1. 00 1. 04 1. 42 1. 34 1. 18 1. 08 1. 16 1. 11 1. 02 . 80 . 70 . 55 . 42 . 31	1. 00 1. 47 .84 .65 .88 1. 31 1. 77 1. 46 1. 19 .74 .55 .47 .47 .47	1. 00 1. 37 . 62 . 51 . 78 1. 21 1. 51 1. 29 1. 17 . 92 . 78 . 81 1. 11 1. 10	1. 00 1. 59 .44 .42 .68 1. 24 1. 98 1. 58 1. 26 .58 .58 .58 .58 .58 .58	1. 00 1. 47 . 35 . 35 . 59 1. 12 1. 71 1. 41 1. 24 1. 00 . 82 . 88 1. 82 3. 00	1. 00 . 91 . 46 . 63 . 75 . 98 1. 18 1. 10 1. 12 1. 12 1. 07 1. 10 1. 77 2. 27

Exclusive of Charles County, Md.

As in the case of comparisons by locality, these relations indicate that the mortality is determined primarily by the incidence of pneumonia. The cause of the high mortality in young adult life evidently lies in the complicating pneumonia. All of the relations shown in this figure bear this out: The peak in the pneumonia case incidence in young adult life, coinciding almost completely with that of the mortality from the epidemic; the absence of a corresponding peak in the total epidemic morbidity (except a minor secondary mode) and (by corollary) a peak in young adult life for influenza case fatality and the percentage of cases complicated by pneumonia, but not for pneumonia case fatality itself.

SEX

Mortality and case fatility rates for influenza and for pneumonia were higher among men than among women, the differences being about 10 per cent on the average. In the case of influenza fatality, this may have been due to the fact that the reports were usually obtained from the female members of the household, giving

a relatively higher rate of influenza among them. But a similar explanation is hardly possible in the case of pneumonia fatality. Table 30 gives the relations between the two sexes for all the measures which have been employed in this report. All ratios based on pneumonia incidence are exclusive of Charles County, Md., as indicated. For the other cases, all 12 localities are used. Since it was found that adjustment for age made little difference in the ratio between the two sexes (see p. 324), these rates are given without adjustment.

Table 30.—Epidemic relations by sex in all surveyed localities during epidemic of 1918-19

	Male	Female	Ratio fe- male to male
Influenza incidence (per 1,000) Pneumonia incidence (per 1,000) Percentage of influenza cases which were complicated by pneumonia 1 Mortality (per 1,000) Case fatality—influenza (per cent) Case fatality—pneumonia alone 1 (per cent)	288	299	104
	18. 4	17. 0	92
	6. 8	5. 9	87
	5. 3	4. 7	89
	1. 8	1. 6	89
	26. 5	24. 5	92

¹ Exclusive of Charles County, Md.

The mortality rates by age and sex have already been given. (Table 27.) The excess among men would seem to occur at the ages when the epidemic took its severest toll (20 to 40). This is equally borne out in the fatality rates, which are given in Figure 11, especially in the case of influenza case fatality. The two sexes evidently present a quite different picture, which may be regarded as of importance in connection with the epidemiological problems raised by the disease. The data are given in Table 31. Table 32 gives corresponding figures for the percentage of cases complicated by pneumonia.

Table 31.—Fatality of influenza and of pneumonia by age and sex in all surveyed localities during epidemic of 1918–19

Age group	Fatality per 100 cases of cases of pneumonia 1		cases of		Age group	Cas	ty per 100 ses of uenza	cas	y per 100 ses of monia ¹
	Males	Females	Males	Females		Males	Females	Males	Females
All ages	1.8	1.6	26. 5	24. 5	20 to 24 25 to 29	2.4 4.1	1.7 2.2	27. 5 36. 3	23. 6 24. 5
Under 1 1 to 4	8. 0 1. 5	6. 7 2. 2	44. 1 12. 9	42. 3 25. 0	30 to 34	2.8 2.7 1.7	2.1 1.6 1.7	29. 9 31. 1 29. 1	25. 9 25. 3 28. 1
Under 5	2.4 .5 .4 1.2	2.8 .6 .7 .8	19. 0 11. 4 11. 4 22. 1	28. 0 12. 2 20. 9 16. 1	45 to 49	1.8 1.7 2.3 4.2	1. 1 1. 3 3. 8 5. 7	28.6 40.0 43.8 60.0	25. 8 21. 2 45. 7 57. 1

¹ Exclusive of Charles County.

Table 32.—Percentage of influenza cases which were complicated by pneumonia, by age and sex in all localities, during epidemic of 1918–19 1

Age group	Both sexes	Male	Female	Age group	Both sexes	Male	Female
All ages	6.8	6.8	5.9	20 to 24	7.6	9. 4 12. 1	6.8 7.9
Under 1	12.2 8.1	13. 6 8. 2	11. 0 7. 9	30 to 34	9. 5 8. 1 7. 4 5. 8	9.4 8.5 5.7	7. 0 6. 2 5. 8
Under 5 5 to 9 10 to 14	8. 6 3. 9 3. 2	8.9 3.7 3.4	8.3 4.1 3.0	45 to 49	4. 9 5. 1 7. 0	5.7 4.1 5.0	4.1 6.0 8.5
15 to 19	4.9	5.8	4.2	70 and over	6. 9	4.5	8.5

¹ Exclusive of Charles County.

COLOR

Outside of Charles County, Md., the fatality rate per 100 cases of influenza was about the same in the white and colored populations, 11.7 and 1.9, respectively. The pneumonia case fatality (excluding Charles County) in the white and colored was 28.8 and 39.8, respectively. Thus we are probably warranted in concluding that the case fatality was really higher in the colored populations of the surveyed communities.

Summary

The purpose of this report has been to make a permanent record, for future reference, of the statistics obtained by the surveys, not to offer any extended discussion of their meaning. Hence there is no necessity for any detailed summary of the findings. Certain major points, however, are of considerable interest.

Special surveys were undertaken at the close of the 1918-19 epidemic of influenza to determine for a population of known sex, age, and color composition the approximate incidence of the disease, and also to ascertain the relations between the epidemic morbidity, the incidence of pneumonia, and the mortality. Preliminary reports on the surveys were published at the completion of the work.

The incidence of influenza (including pneumonia and "doubtful" cases) was 294 per 1,000 for all localities, varying from 535 to 150. These rates correspond closely with what was found in other surveys of the same general character. There seemed to be no clear indication of a geographical difference in incidence.

The incidence was highest among very young persons (age group 5 to 9 years), with a secondary peak at about 30 years. The rate of attack fell off rapidly in older life. Among old people the incidence appeared to be not more than one-third of that among the young.

Slightly higher influenza rates were found among females (except in two localities), but it seemed possible that this was due to the fact that most of the reports as to illness came from the women, who

[&]quot; New London, San Antonio, Des Moines, and San Francisco excluded. In the case of these calculations by color, it was not possible to add to the pneumonia cases deaths reported as due to influenza.

might remember their own illnesses better than those of other members of the family. The colored had lower rates of influenza incidence, but it is possible that the reporting among them was less complete.

A special effort was made to determine the incidence of pneumonia as complicating the original case of influenza. For all localities the

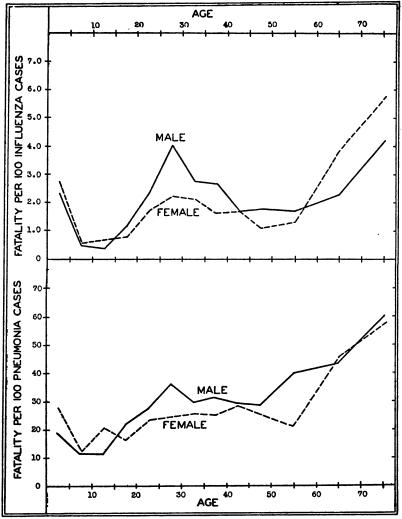


FIGURE 11.—Influenza and pneumonia fatality, by age and sex, in all surveyed localities during the 1918-19 epidemic. (Pneumonia fatality is exclusive of Charles County)

pneumonia rate was 17.6 per 1,000 persons, varying from 25.8 to 6.7. In other words, about 6 per cent of the influenza cases were complicated by pneumonia.

The peak in young adult life suggested in the epidemic morbidity as a whole comes out with remarkable clarity in the pneumonia

incidence. For all localities the rate is about 25 per 1,000 at the beginning of life, falls to about 11 in the age group 10 to 14, and then rises to a secondary mode of about 31 in the age group 25 to 29. After that the rate falls rather steadily to the end of life. This age distribution is, of course, fundamentally different from the normal course of pneumonia incidence, which is high among the very young and among the very old. The striking mode in young adult life is found in each locality without exception.

The pneumonia incidence rates were slightly higher among males, the difference being especially marked in young adult life. The recorded pneumonia incidence was higher among the white than among the colored.

The deaths from influenza-pneumonia during the epidemic period were obtained primarily to determine the relations as to case fatality. The fatality per 100 cases of influenza (total epidemic morbidity) was 1.70 for all localities, and that per 100 cases of pneumonia alone was 25.5. The fatality for the surveyed localities (total epidemic morbidity) seemed about the same as that recorded in other studies.

The fatality seemed lower in the southern and central localities, which is in line with other reports on this epidemic. It appeared that the incidence of pneumonia, rather than that of influenza as a whole, determined the mortality in the various localities.

The fatality of influenza (total epidemic morbidity) was very high among young adults, as would be expected in view of the high peak of pneumonia at these ages. The fatality of pneumonia did not show this peak, showing that the tendency to a severe toll at these ages was characteristic of the pneumonia itself, rather than of death from it.

The fatality rates, both for influenza and for pneumonia, were higher among men than among women. In the case of influenza, this may reflect the tendency of the women to report more adequately; but that would hardly explain the difference in the case of pneumonia fatality. The excess was most marked in young adult life.

The pneumonia case fatality was much higher among the colored than among the white.

Acknowledgments

Special acknowledgment is made to the Influenza Commission of the Metropolitan Life Insurance Co., which defrayed part of the expenses of the tabulation and analysis of the data collected in the surveys.

The surveys themselves were made under the direction of Senior Surg. W. H. Frost and Principal Statistician Edgar Sydenstricker, and the first papers reporting the results were prepared by them. Doctor Frost and Mr. Sydenstricker also supervised the more extended analysis on which this paper is based.

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The Incidence of Influenza among Persons of Different Economic Status During the Epidemic of 1918. By Edgar Sydenstricker. Pub. Health Rep., vol. 46, No. 4, January 23, 1931. (Reprint 1444.)

Mortality from Influenza and Pneumonia in 50 Large Cities of the United States, 1910–1929. By S. D. Collins, W. H. Frost, Mary Gover, and Edgar Sydenstricker. Pub. Health Rep., vol. 45, No. 39, September 26, 1930. (Reprint 1415.)

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DEATHS DURING WEEK ENDED JANUARY 16, 1932

Summary of information received by telegraph from industrial insurance companies, for the week ended January 16, 1932, and corresponding week of 1931. (From the Weekly Health Index, issued by the Bureau of the Census, Department of Commerce)

,	Week ended Jan. 16, 1932	Corresponding week, 1931
Policies in force	74, 179, 429	75, 092, 689
Number of death claims	15, 052	17, 116
Death claims per 1,000 policies in force, annual rate	10. 6	11. 9
Death claims per 1,000 policies, first 2 weeks of year,		
annual rate	9.9	11. 2

Deaths 1 from all causes in certain large cities of the United States during the week ended January 16, 1932, infant mortality, annual death rate, and comparison with corresponding week of 1931. (From the Weekly Health Index, issued by the Bureau of the Census, Department of Commerce)

[The rates furnished in this summary are based upon mid-year population estimates derived from the 1930 census]

	Wed	ek ended	l Jan. 16,	1932		ponding , 1931		ate ² for 2 weeks
City	Total deaths	Death rate ²	Deaths under 1 year	Infant mor- tality rate 3	Death rate?	Deaths under 1 year	1932	1931
Total (83 cities)	8, 402	12. 1	633	4 53	14.0	803	12. 5	14.0
Akron Albany * Atlanta * White Colored Baltimore * White Colored Birmingham * White Colored Birmingham * White Colored Boston Bridgeport Buffalo Cambello Camden Canton Chicago * Cincinnati Cleveland Colored Dallas * White Colored Dayton Denver Des Moines Detroit Duluth EI Paso Erie Fall River *7 Filint Fort Worth * White Colored Colored Grand Rapids Forted Colored	45 47 77 77 77 77 40 280 177 53 16 177 34 242 38 39 25 683 192 82 82 82 82 83 83 192 83 84 14 15 16 17 17 18 18 19 19 19 19 19 19 19 19 19 19	8. 9 14. 2 10. 3 21. 9 14. 7 18. 4 11. 5 8. 2 16. 0 16. 0 17. 1 14. 1 10. 1 14. 1 16. 0 17. 1 18. 0 19. 0 11. 8 8. 2 11. 5 11. 5	3 0 9 3 6 18 8 10 8 5 4 1 1 25 4 9 7 7 3 5 5 16 4 10 7 7 3 3 5 4 10 7 7 7 7 7 7 8 7 8 7 8 7 8 7 8 7 8 7 8	37 0 88 44 172 64 45 129 66 27 76 71 43 145 53 124 50 32 52 40 0 65 58	7. 9 14. 9 11. 9 12. 4 14. 2 20. 6 10. 3 18. 8 16. 6 12. 1 11. 4 11. 7 13. 2 10. 8 11. 7 13. 2 11. 7 13. 2 12. 8 17. 4 18. 8 17. 4 18. 8 17. 4 18. 8 19. 9 19. 9 1	4 2 9 5 4 9 111 8 10 2 2 8 8 20 4 13 5 5 2 2 2 2 2 4 1 1 3 5 5 5 2 2 2 2 2 4 1 1 3 5 5 5 2 2 2 2 2 4 1 1 3 5 5 5 5 2 2 2 2 2 2 4 1 1 3 5 5 5 5 2 2 2 2 2 4 1 1 3 5 5 5 5 2 2 2 2 2 2 4 1 1 3 5 5 5 5 2 2 2 2 2 2 4 1 1 3 5 5 5 5 2 2 2 2 2 2 4 1 1 3 5 5 5 5 2 2 2 2 2 2 4 1 1 3 5 5 5 5 2 2 2 2 2 2 4 1 1 3 5 5 5 5 2 2 2 2 2 2 4 1 1 3 5 5 5 5 2 2 2 2 2 2 4 1 1 3 5 5 5 5 5 2 2 2 2 2 2 4 1 1 3 5 5 5 5 5 5 5 5 5 5 5 5 5 5 5 5 5	9.6 16.6 17.2 12.5 26.5 14.0 18.4 10.7 15.9 15.9 11.3 16.4 11.5 11.5 11.6 9.8 11.1 11.3 11.6 9.8 11.4 11.6 11.6 11.6 11.6 11.6 11.6 11.6	8. 7 16. 00 12. 6 16. 00 12. 6 16. 00 12. 6 12. 6 16. 00 11. 2 12. 6 16. 00 11. 2 12. 6 16. 00 11. 2 12. 6 16. 00 12. 9 11. 3 13. 8 13. 7 13. 6
White Colored Indianapolis 6 White Colored	46 30 92 79 13	10. 1 18. 3 12. 8 12. 6 14. 7	2 2 7 7 0	57 64 0	13. 1 14. 2 10. 1 15. 4 14. 6 20. 8	4 0 9 0	11. 8 10. 0 16. 8 13. 6 12. 8 19. 3	12.9 13.2 11.9 14.9 14.4 19.0

See footnotes at end of table.

Deaths 1 from all causes in certain large cities of the United States during the week ended January 16, 1932, infant mortality, annual death rate, and comparison with corresponding week of 1931—Continued.

	Wee	ek ended	Jan. 16,	1932		ponding , 1931		ate ' for 2 weeks
City	Total deaths	Death rate 2	Deaths under 1 year	Infant mor- tality rate 3	Death rate?	Deaths under 1 year	1932	1931
Jersey City Kansas City, Kans. White Colored Kansas City, Mo	66 27 16	10. 8 11. 4 8. 4 24. 3	5 0 0	41 0 0	11. 8 15. 7 14. 2 22. 2	9 5 3 2	11. 7 14. 1 13. 3 17. 6	12. 3 15. 9 14. 4 22. 2
White	11 91 31 24 7	11. 4 14. 5 13. 4 20. 0	0 7 6 6	79 152 167	15. 4 15. 3 14. 3 20. 5	8 6 5	10. 2 11. 7 10. 6 17. 1	15. 0 15. 5 13. 7 24. 9
Colored Long Beach Los Angeles Louisville 6	30 326 103 80	9. 7 12. 3 17. 4 16. 0	0 16 2 2	0 47 18 21	11. 3 13. 5 16. 4 14. 4	1 22 7 6	11. 2 12. 8 15. 5 14. 2	10. 6 14. 1 20. 9 19. 2
White Colored Lowell 7 Lynn	23 29 22 101	25. 2 15. 1 11. 2 20. 0	0 2 0 13	0 52 0 142	27. 3 12. 0 15. 2 15. 9	1 4 1 2	22. 4 14. 1 12. 4 17. 8	30. 1 13. 3 14. 7 16. 8
Memphis 6	40 61 31 24	12. 8 31. 7 14. 2 14. 2	3 10 1	51 301 28 39	12. 1 22. 1 14. 4 16. 1	0 2 3 2	12. 5 26. 2 14. 5 13. 9	10. 5 14. 7 20. 3 12. 3 13. 5
Willie Colored Milwaukee Minneapolis Nashville 6	7 101 83 34	14. 5 8. 8 9. 0 11. 3	0 6 6 3	0 29 39 45	8. 2 10. 0 11. 6 18. 1	1 9 14 2	16. 5 9. 8 9. 6 13. 7	8. 2 10. 3 12. 8 17. 3
WhiteColored	26 8 23 47	11. 9 9. 8 10. 7 15. 1	3 0 1 1	59 0 29 20	16. 2 23. 1 12. 5 10. 3	1 1 2 1	13. 3 14. 6 12. 3 14. 1	14. 8 23. 8 13. 9 12. 2
New Haven New Orleans 6 White Colored New York	137 85 52 1,499	15. 1 13. 2 19. 8 10. 9	9 4 5 126	51 35 82 56	21. 0 18. 5 27. 1 15. 7	12 4 8 165	16. 2 13. 8 22. 1 11. 5	21. 4 18. 6 28. 3 14. 9
Bronx Borough Brooklyn Borough Manhattan Borough	220 525 558 152	8. 3 10. 2 16. 4 6. 6	10 49 53 10	29 54 76 42	11. 4 14. 8 23. 9 10. 0	19 60 64 18	8. 9 10. 3 17. 5 7. 6	10. 3 14. 1 22. 6 9. 8
Queens Borough Richmond Borough Newark, N. J Oakland	44 96 70 40	13. 7 11. 2 12. 2 10. 2	4 9 4	79 49 50 55	13. 1 13. 5 13. 0 11. 9	4 8 2 6	15. 6 11. 2 12. 3 11. 0	14. 4 13. 1 14. 4 12. 2
Oklahoma City Omaha Paterson Peoria Philadelphia	58 40 21 451	13. 9 15. 0 9. 9 11. 9	1 2 1 25	11 36 28 39	15. 6 17. 3 18. 8 16. 5	6 4 5 43	14. 1 15. 4 11. 5 13. 1	14. 7 15. 2 16. 8 16. 1
Philadelphia. Pittsburgh. Portland, Oreg. Providence Richmond 6.	157 85 89 57	12. 1 14. 3 18. 1 16. 1	15 2 8 8	69 26 77 121	16. 7 13. 8 13. 9 16. 1	24 4 11 7	14. 1 14. 3 18. 3 17. 2	16. 6 14. 7 15. 3 16. 4
Colored	31 26 85 288	12. 2 25. 7 13. 3 18. 1	4 4 9 23	90 183 86 82	11. 5 27. 6 11. 5 16. 3	2 5 9 23	14. 4 24. 3 12. 4 15. 8	13. 9 22. 7 13. 1 16. 5
St. Louis St. Paul Salt Lake City San Antonio San Diego San Francisco Schenectady Seattle	52 29 83 52	9. 7 10. 4 17. 6 16. 6	5 0 12 3	53 0 65	11. 0 14. 2 16. 9 14. 7	4 4 16 5	9. 5 11. 9 15. 0 15. 7	11. 4 14. 0 16. 1 17. 0
Somerville	163 15 86 22	12. 9 8. 1 11. 9 10. 8	1 4 1	28 29 40 40	16. 6 8. 1 16. 4 8. 9	5 1 5 0	15. 0 9. 8 12. 1 12. 1	14. 6 8. 4 14. 5 11. 6
South Bend	19 29 38	8. 9 13. 0 12. 9 13. 1	2 1 5 4	58 27 84 52	5. 8 17. 5 9. 6 12. 2	1 5 0	8. 5 14. 5 13. 4 12. 0	5. 8 14. 1 10. 6 13. 0
Tacoma Tampa White Colored	54 22 26 18 8	10. 6 12. 6 11. 0 18. 3	2 3 2 1	55 86 70 158	14. 5 14. 4 12. 0 23. 5	5 2 4 2 2 3	10. 1 11. 4 10. 7 13. 8	13. 8 17. 1 15. 1 24. 7
Toledo	68	11.8	41	43	11.2	3	11.6	11.8

See footnotes at end of table.

Deaths 1 from all causes in certain large cities of the United States during the week ended January 16, 1932, infant mortality, annual death rate. and comparison with corresponding week of 1931—Continued

	Wee	ek ended	Jan. 16,	1932		ponding , 1931	Death rate 2 for the first 2 weeks	
City	Total deaths	Death rate ¹	Deaths under 1 year	Infant mor- tality rate 3	Death rate ³	Deaths under 1 year	1932	1931
Trenton Utlea Washington, D. C.4 White Colored Waterbury Wilmington, Del.7 Worcester Yonkers Youngstown	33 33 167 111 56 15 32 49 23 30	13. 9 16. 8 17. 7 16. 2 21. 4 7. 7 15. 7 12. 9 8. 5 8. 9	0 1 10 4 6 0 2 3 2 5	0 28 56 33 107 0 45 42 52 81	12.6 20.4 17.7 15.2 24.3 6.7 10.3 16.9 11.6 9.3	1 3 8 5 3 1 6 1 4	16. 4 14. 2 15. 9 14. 2 20. 3 8. 2 15. 9 13. 7 8. 6 9. 2	19. 6 17. 6 18. 6 16. 3 24. 7 7. 8 14. 7 14. 9 11. 1

¹ Deaths of nonresidents are included. Stillbirths are excluded.

Deaths under 1 year of age per 1,000 live births. Cities left blank are not in the registration area for births.

Data for 78 cities.
Deaths for week ended Friday.

² These rates represent annual rates per 1,000 population, as estimated for 1932 and 1931 by the arithmetical method.

Deaths for week ended Friday.
For the cities for which deaths are shown by color, the percentages of colored population in 1930 were as follows: Atlanta, 33; Baltimore, 18; Birmingham, 38; Dallas, 17; Fort Worth, 16; Houston, 27; Indianapolis, 12; Kansas City, Kans., 19; Tampa, 21; Knoxville, 16; Louisville, 15; Memphis, 38; Miami, 23; Nashville, 28; New Orleans, 29; Richmond, 29; and Washington, D. C., 27.
Population Apr. 1, 1930; decreased 1920 to 1930, no estimate made.

PREVALENCE OF DISEASE

No health department, State or local, can effectively prevent or control disease without knowledge of when, where, and under what conditions cases are occurring

UNITED STATES

CURRENT WEEKLY STATE REPORTS

These reports are preliminary, and the figures are subject to change when later returns are received by the State health officers

Reports for Weeks Ended January 23, 1932, and January 24, 1931

Cases of certain communicable diseases reported by telegraph by State health officers for weeks ended January 23, 1932, and January 24, 1931

	Diph	theria	Influ	jenza	Me	esles		ococcus ngitis
Division and State	Week ended Jan. 23, 1932	Week ended Jan. 24, 1931	Week ended Jan. 23, 1932	Week- ended Jan. 24, 1931	Week ended Jan. 23, 1932	Week ended Jan. 24, 1931	Week enged Jan. 23, 1932	Week ended Jan. 24, 1931
New England States: Maine. New Hampshire Vermont Massachusetts. Rhode Island. Connecticut.	2 2 55 4	5 3 2 61 8	181	17 114 1 140	633 44 334 349 1, 056	20 25 8 643	0 0 0 1 0	0 0 0 4 1
Middle Atlantic States: New York New Jersey Pennsylvania East North Central States:	168 30 112	126 58 127	1 29 11	1 1, 140 744	884 104 1, 030	329 388 1, 022	6 5 9	20 4 8
Ohio. Indiana. Illinois. Michigan. Wisconsin.	86 69 170 46 19	39 64 162 48 24	15 29 33 1 28	7 33 263 2 82	141 213 68 217 89	140 251 905 143 172	1 6 8 6 3	5 12 6 6 1
West North Central States: Minnesota	17 26 57 6 11 44	8 39 5 26 8 28	7 4	71 	68 3 28 86 56 14 79	28 3 1, 109 3 12 30 53	1 0 0 0 0 0	1 2 5 0 0 1 2
South Atlantic States: Delaware	3 35 19	4 25 11	1 41 1	1 1, 226 28	2 11 3	3 229 25	0 3 0 3	0 0 1
Virginia West Virginia North Carolina South Carolina Georgia Florida	42 38 12 24 10	13 33 16 33 11	64 23 389 126 7	150 177 1, 968 267 42	336 137 20 9	30 163 27 108 63	0 3 0 1	3 0 3 0 2 1

New York City only.
 Week ended Friday.
 Typhus fever, week ended Jan. 23, 1932, 6 cases: 1 case in Maryland, 1 case in Georgia, and 4 cases in Alabama.

Cases of certain communicable diseases reported by telegraph by State health officers for weeks ended January 23, 1932, and January 24, 1931—Continued

	Diph	theria	Infl	uenza	Me	asl es		gococcus ingitis
Division and State	Week ended Jan. 23, 1932	Week ended Jan. 24, 1931	Week ended Jan. 23, 1932	Week ended Jan. 24, 1931	Week ended Jan. 23, 1932	Week ended Jan. 24, 1931	Week ended Jan. 23, 1932	Week ended Jan. 24, 1931
East South Central States: Kentucky Tennessee Alabama ³ Mississippi West South Central States:	1 21	16 15 60 14	43 95	14 187 87	94 16 17	76 110 458	0 4 4 1	7 2 5
West South Central States: Arkansas Louislana Oklahoma 4 Texas Mountain States:	35 50	14 21 26 32	18 4 73 63	209 91 155 102	2 4 84 10	9 2 74 141	0 1 0 1	3 1 0 2
Montone	6	1	11		103 2 1	2	0 1 0	0 1 3
Montana. Idaho. Wyoming. Colorado. New Mexico. Arizona Utah 2 Pacific States:	9 17 2 1	9 4 14 3	250 42	1 22 1	8	29 21 125 2	0 0 0	0 1 3 0 0 9 1
Washington Oregon California	9 5 79	25 5 62	70 235	56 93	443 40 252	62 115 546	0 0 5	2 0 6
	Polion	yelitis	Scarle	t fever	Sma	llpox	Typhoi	id fever
Division and State	Wcek ended Jan. 23, 1932	Week ended Jan. 24, 1931	Week ended Jan. 23, 1932	Week ended Jan. 24, 1931	Week ended Jan. 23, 1932	Week ended Jan. 24, 1931	Week ended Jan. 23, 1932	Week ended Jan. 24, 1931
New England States: Maine. New Hampshire Vermont. Massachusetts. Rhode Island. Connecticut. Middle Atlantic States:	0 1 0 2 0 1	4 0 0 3 1 0	25 14 7 549 36 87	36 5 2 325 65 74	0 1 26 20 0 4	0 0 0 6 0	1 0 1 3 0	2 0 1 0 1
New York New Jersey Pennsylvania East North Central States:	1 1 3	0 0 2	909 209 589	739 252 580	• 0 0	1 0 2	15 3 26	6 1 1
Ohio	0 0 5 1 3	1 1 4 1 0	323 100 398 319 111	363 391 521 381 145	34 31 29 16 5	73 108 51 88 4	10 0 14 2 2	8 0 7 4 0
Minnesota Iowa Missouri North Dakota South Dakota Nebraska Kansas	1 1 0 3 1 1	2 1 2 2 1 3 0	87 64 89 15 7 15 74	4 89 178 27 6 51 68	0 67 23 1 17 5	12 46 24 10 38 28 87	0 0 1 2 4 0 4	3 2 5 0 2 5
South Atlantic States: Delaware Maryland 23 District of Columbia Virginia West Virginia North Carolina South Carolina	0 1 0	0	8 92 21	33 82 32	0	0 0 0	0 12 3	0 3 1
Virginia West Virginia North Carolina South Carolina Georgia Florida	1 0 2 0 1 1	2 0 1 1 0 0	46 57 11 32	57 58 17 68 7	4 1 2 0 0	19 0 0 0 0	7 5 11 5	12 1 4 7 2

Week ended Friday.
 Typhus fever, week ended Jan. 23, 1932, 6 cases: 1 case in Maryland, 1 case in Georgia, and 4 cases in Alabama.
 Figures for 1932 are exclusive of Oklahoma City and Tulsa, and for 1931 are exclusive of Tulsa only.

Cases of certain communicable diseases reported by telegraph by State health officers for weeks ended January 23, 1932, and January 24, 1931—Continued

	Poliomyelitis		Scarle	t fever	Sma	llpox	Typhoid fever	
Division and State	Week ended Jan. 23, 1932	Week ended Jan. 24, 1931						
East South Central States:								
	1	0	124	114	8	16	19	9
KentuckyTennessee	Ō	Ŏ	62	42	16	5	21	9 3 14 2
Alabama 3		3	30	62	16	6	24	14
Mississippi	l ī	Ŏ	23	25	58	12	6	2
West South Central States:	_							1
Arkansas	. 0	1 1	14	35	20	42	6	8 2
Louisiana		ī	15	28	4	10	9	1 2
Oklahoma 4	l õ	Õ	26	34	36	97	2	10
Texas	i	Ŏ	98	65	72	31	8	7
Mountain States:	-					-	1	1
Montana	0	0	45	59	2	2	3	2
Idaho.	Ŏ	i	12	20	2	2	0	l o
Wyoming	Ŏ	ī	18	50	Ō	2	Ō	l ó
Colorado	Ŏ	Ō	46	45	4	19	2	l i
New Mexico		Ŏ	9	7	1	2	ī	1 0
Arizona		ň	5	4	ē	14	Ō	l i
Utah 3	Ŏ	ŏ	18	6	Ō	ī	Ō	li
Pacific States:		ľ			1	_		l "
Washington	0	1	39	50	32	36	3	1 2
Oregon.	ĭ	î	27	14	20	19	3	Ō
California	2	7	149	142	21	82	5	l š

SUMMARY OF MONTHLY REPORTS FROM STATES

The following summary of cases reported monthly by States is published weekly and covers only those States from which reports are received during the current week:

State	Menin- gococ- cus menin- gitis	Diph- theria	Influ- enza	Ma- laria	Mea- sles	Pel- lagra	Polio- myelitis	Scarlet fever	Small- pox	Ty- phoid fever
November, 1931 Colorado	4 1	22 311	1	· · · · · · · · · · · · · · · · · · ·	102		1	129 282	31	26 12
Alabama. Arkansas Idaho Illinois. Indiana Maryland Minnesota Missouri New Jersey North Carolina Pennsylvania Pento Rico Rhode Island West Virginia	6 2 1 25 39 5 7 16 7 9 24	263 132 7 603 325 288 121 411 153 360 544 68 27 188	83 49 14 104 71 94 6 25 51 88 80 18	87 39 23 4 10 19,859	73 49 5 168 121 38 71 37 126 187 2, 791 126 2, 249 1, 085	54 13 2 1 114 1 3	11 0 0 42 3 26 3 9 5 5 22 2	207 103 63 1,378 406 420 260 381 591 1,914	2 40 23 88 36 0 35 38 2 1 0	72 41 4 82 29 44 15 24 13 25 92 15 0

Week ended Friday.
 Typhus fever, week ended Jan. 23, 1932, 6 cases: 1 case in Maryland, 1 case in Georgia, and 4 cases in Alabama.

⁴ Figures for 1932 are exclusive of Oklahoma City and Tulsa, and for 1931 are exclusive of Tulsa only.

November, 18	91	Cases	Mumps:	Cases
Colorado:			Alabama	
Paratyphoid fever		. 1	Arkansas	
Kansas:		979	Idaho	
Chicken pox			Illinois Indiana	
Impetigo contagiosa			Maryland	
Mumps			Missouri	
Paratyphoid fever			New Jersey	
Scables			Pennsylvania	
Septic sore throat			Porto Rico	10
Tetanus		2	Rhode Island	128
Trench mouth			Ophthalmia neonatorum:	
Tularaemia		7	Alabama	1
Undulant fever		7	Illinois	7
Vincent's angina			Maryland	4
Whooping cough		113	Minnesota	1
			Missouri Pennsylvania	2
December, 193	11		Porto Rico.	18 20
Chicken pox:			Rhode Island	1
Alabama		133	Paratyphoid fever:	•
Arkansas		52	Idaho	1
Idaho		130	Illinois.	2
IllinoisIndiana		598	New Jersey	1
Maryland		274	North Carolina	1
Minnesota		470	Porto Rico.	5
Missouri		382	Rhode Island	1
New Jersey		759	Puerperal septicemia:	
North Carolina		508	Illinois	20
Pensylvania			Pennsylvania	19
Porto Rico		8	Porto Rico	4
Rhode Island		78	Rabies in animals:	
West Virginia		2 55	Illinois Maryland Maryland	3 3
Diarrhea:			Missouri	2
Maryland		10	Rhode Island	1
Dysentery: Illinois		5	Rabies in man:	-
Illinois (amoebic)		4	Illinois	2
Maryland		8	Scabies:	
Minnesota		3	Maryland	10
Minnesota (amoebic)		2	Septic sore throat:	
Missouri		1	Illinois	32
New Jeisey		1	Maryland	4
Pennsylvania		1	Missouri	22
Porto Rico		67	North Carolina Rhode Island	12 3
Filariasis:			Tetanus:	0
Porto Rico		23	Illinois.	14
German measles: Illinois		21	Maryland	1
Maryland		1	New Jersey	1
New Jersey		31	Pennsylvania	1
North Carolina		13	Porto Rico	2
Pennsylvania		159	Tetanus, infantile:	2.1
Rhode Island		11	Porto Rico	4
Impetigo contagiosa:			Trachoma:	:
Maryland		30	Arkansas	1
Lead poisoning:			Illinois	10
Illinois		5	Indiana	1
New Jersey		2	Missouri	29
Lethargic encephalitis:		3	New Jersey Pennsylvania	4
Alabama		7	Porto Rico	11
Maryland		il	Trichinosis:	
Pennsylvania		4	New Jersey	1

Tularaemia:	Cases	Vincent's angina:	Cases
Alabama	. 1	Illinois	27
Arkansas	. 1	Maryland	. 12
Illinois	. 54	Whooping cough:	
Indiana	. 11	Alabama	. 19
Maryland	. 9	Arkansas	. 22
Minnesota	. 3	Illinois	1, 250
Missouri	. 13	Indiana	208
New Jersey	. 1	Maryland	. 637
Pennsylvania	. 3	Minnesota	. 96
Typhus fever:		Missouri	446
Alabama	. 11	New Jersey	. 738
North Carolina	. 1	North Carolina	. 55 2
Undulant fever:		Pennsylvania	. 1,984
Alabama		Porto Rico	. 206
Illinois	. 3	Rhode Island	. 28
Indiana	. 1	West Virginia	. 111
Maryland	. 3	Yaws:	
Minnesota	2	Porto Rico	. 9
Missouri	6		
New Jersey	2		
Pennsylvania	6		

Cases of certain communicable diseases reported for the month of November, 1931 by State health officers

	,								
State	Chick- en pox	Diph- theria	Mea- sles	Mumps	Scarlet fever	Smallpox	Tuber- culosis	Ty- phoid and para- ty- phoid fever	Whoop- ing cough
Maine	193	17	782	10	139	0	40	16	80
New Hampshire		21			23	ŏ		1	00
Vermont	253	30	141	53	58	75	15	0	277
Massachusetts	488	243	390	627	906	0	437	15	474
Rhode Island	70	34	571	38	71	0	53	0	19
Connecticut	206	17	99	118	167	0	89	18	148
New York	1, 439	419	904	398	1, 787	70	1, 471	100	1, 109
New Jersey	524	134	122	87	499	ĭ	384	21	641
Pennsylvania	2, 504	508	1, 352	1, 108	1, 603	Ō	607	242	1,743
Ohio	1 000	-00	004	- 45	0.00=		004	100	1 001
OhioIndiana	1, 836 408	568 361	234 138	547	2, 005	55	334	136	1, 321
Illinois	1, 077	586	210	63 133	415 1, 176	31 71	222 620	27 87	13 7 1, 15 8
Michigan	763	210	326	310	787	64	257	48	694
Wisconsin	1, 172	95	101	518	294	29	111	16	667
	•								
Minnesota	353	114	72		197	10	180	12	38
Iowa.	363	83	13	14	201	258	33	16	111
Missouri	246 126	412	80	19	465	9	217	66	440
North Dakota	120	16 39	7 216	33 33	79	73	. 8	20	22 33
Nebraska	165	93	42	43	108	44 29	11 19	11 5	52
Kansas	352	311	102	98	282	31	68	14	113
		011	102		202	0.1	00	11	110
Delaware	13	144	2	5	36	0	23	3	28
Maryland	250	289	21	1 2 3	432	0	147	94	599
District of Columbia	22	60	9		92	0	84	14	67
Virginia	419	1, 335	211		729	6	126	130	738
West Virginia	283	228	730		249	2	65	152	213
North Carolina	389	691	184		714	4		64	5 36
South CarolinaGeorgia	84 61	340 179	38 25	58 20	62 149	1	113 105	43 80	72 35
Florida	12	89	25 35	13	24	2	37	80 12	35 5
						-	ا "		•
Kentucky 1									
Tennessee	73	5 24	26	29	345	20	156	117	277
Alabama	67	397	26	22	247	2	295	89	61
Mississippi	238	391	27	45	154	44	75	40	302

¹ Reports received weekly.

Cases of certain communicable diseases reported for the month of November, 1931, by State health officers—Continued

State	Chick- en pox	Diph- theria	Mea- sles	Mumps	Scarlet fever	Small- pox	Tuber- culosis	Ty- phoid and para- ty- phoid fever	Whoop- ing cough
Arkansas Louisiana Oklahoma ³ Texas	34 12 46	234 243 431 364	35 28 8	7 3 30	137 148 192 183	11 14 23	² 10 ² 177 38	50 97 107 48	55 19 25
Montana Idaho Wyoming Colorado New Mexico Arizona	174 87 31 118 99	18 20 22 78 73	571 6 9 5	4 72 20 17 10	127 46 31 129 51 26	. 6 3 2	53 18 38 88	11 3 1 27 37 9	18 2 14
Utah ¹ Nevada Washington Oregon California	2 442 1 253 1, 031	50 8 456	135 26 574	97 56 409	5 235 71 579	58 36 28	179 43 707	0 22 13 58	11 62 27 351

¹Reports received weekly.

Case rates per 100,000 population (annual basis) for the month of November, 1931

									
State	Chick- en pox	Diph- theria	Measles	Mumps	Scarlet fever	Small- pox	Tuber- culosis	Ty- phoid and para- typhoid fever	Whoop- ing cough
Maine	293	26	1, 188	15	211	0	61	24	122
New Hampshire Vermont		55			60	0		3	
Massachusetts	854 138	101 69	476 110	179	196	253	51		935
Rhode Island	122	59	996	177	256 124	0	124 92	4	134
Connecticut	153	13	74	66 88	124	ŏ			33
Connecticut	100	13	1 14	80	124	U	66	13	110
New York	136	40	86	38	169	7	139	9	105
New Jersey	154	39	36	26	146	ó	113	6	188
Pennsylvania	313	63	169	138	200	Ŏ	76	3ŏ	218
01.7									
Ohio	331	102	42	99	361	10	60	24	238
Indiana	152	134	51	23	154	12	82	10	51
Illinois	169	92	33	21	184	11	97	14	181
Michigan	186	51	80	76	192	16	63	12	169
Wisconsin	479	39	41	212	120	12	45	7	273
Minnesota	166	54	34		93	5	85	- 6	18
Iowa	178	41	6	7	99	127	16	8	54
Missouri	82	137	27	6	155	3	72	22	146
North Dakota	224	28	12	59	140	130	14	36	39
South Dakota	209	68	376	57	106	77	19	19	57
Nebraska	145	82	37	38	95	25	17	4	46
Kansas	226	200	66	63	181	20	44	g	73
								Ĭ,	
Delaware	66	729	10	25	182	0	116	15	142
Maryland District of Columbia	184	213	15	90	318	0	108	69	441
District of Columbia	54	148	22		227	0	207	35	165
Virginia	209	667	105		364	3	63	65	369
West Virginia	195	157	504		172	1	45	105	147
North Carolina	146	259	69		268	1		24	201
South Carolina	59	237	26	40	43	1	79	30	50
Georgia Florida	26	75	10	.8	62	0	44	33	15
r IVI IUA	10	71	28	10	19	2	29	10	4
Kentucky 1						l		1	
Tennessee	34	241	12	13	158	9	72	54	127
Alabama	30	180	12	10	112	i	134	40	28
Mississippi	142	234	16	27	92	26	45	24	180
1 Paparta received weekly		•							

¹ Reports received weekly.

² Pulmonary.

¹ Exclusive of Oklahoma City and Tulsa.

Case rates per 100,000 population (annual basis) for the month of November, 1931—Continued

State	Chick- en pox	Diph- theria	Measles	Mumps	Scarlet fever	Small- pox	Tuber- culosis	Ty- phoid and para- typhoid fever	Whoop- ing cough
ArkansasLouisiana	22 7	152 138	23 16	5 2	89 84	7 8	2 7 2 101	33 55	36 11
Oklahoma ¹	27	250 74	5	17	112 37	13	22	62 10	15
Montana Idaho	394 237	41 54	1, 292	9 196	287 125	14 8	120 2 22	25 8	136
Wyoming	164	26	32	106	164 150	11		5 31	95
New Mexico	333 269	220 198	25 14	48 27	144 71	3 5	107 239	104 24	6 38
Nevada	26				66	0	, 13		144
Washington Oregon California	339 316 211	38 10 93	103 32 117	74 70 84	180 89 118	44 45 6	137 54 145	17 16 12	47 34 72

¹ Reports received weekly.

PATIENTS IN INSTITUTIONS FOR THE CARE OF EPILEPTICS, JANUARY TO MARCH. 1930

Reports for the first quarter of the year 1930 were received by the Public Health Service from 13 institutions for the care and treatment of epileptics, located in 13 States. The total number of patients, including those on parole or otherwise absent, but still on the books, on March 31, 1930, was 8,677.

The first admissions were as follows:

Month	Male	Female	Total
January, 1930 February, 1930 March, 1930	62 76 65	35 41 44	97 117 109
Total	203	120	323

Of the new admissions during the three months, 62.8 per cent were males and 37.2 per cent were females, giving a ratio of 169 males per 100 females.

During the quarter 120 patients were discharged—71 males and 49 females. Seventy-four male patients and 76 female patients died. The annual death rates, based on the total number of patients of the institutions on March 31, 1930, were: Males, 65.1 per 1,000; females, 75.8 per 1,000; persons, 70.1 per 1,000.

At the end of March there were 4,613 males and 4,064 females on the rolls of the institutions, giving a ratio of 114 males per 100 females

² Pulmonary.

³ Exclusive of Oklahoma City and Tulsa.

The following table shows for the 13 institutions the numbers of patients in the hospitals and on parole on January 1, 1930, and at the end of each month of the first quarter of the year:

	Jan. 1,	Jan. 31,	Feb. 28,	Mar. 81,
	1930	1930	1930	1930
Patients in hospitals: Male Female	4, 196	4, 321	4, 861	4, 375
	8, 827	3, 867	8, 897	8, 909
Total	8, 023	8, 188	8, 258	8, 284
Patients on parole: Male Female	325	227	220	238
	215	169	151	155
Total	540	396	371	393
Total patients on books: Male Female Total	4, 521	4, 548	4, 581	4, 613
	4, 042	4, 036 '	4, 048	4, 064
	8, 563	8, 584	8, 629	8, 677
Per cent of total patients on parole: MaleFemale	7. 2	5. 0	4. 8	5. 2
	5. 3	4. 2	3. 7	3. 8
Total	6. 3	4.6	4. 3	4. 5

GENERAL CURRENT SUMMARY AND WEEKLY REPORTS FROM CITIES

The 92 cities reporting cases used in the following table are situated in all parts of the country and have an estimated aggregate population of more than 33,460,000. The estimated population of the 85 cities reporting deaths is more than 31,903,000. The estimated expectancy is based on the experience of the last nine years, excluding epidemics.

Weeks ended January 16, 1932, and January 17, 1931

	1932	1931	Estimated expectancy
CASES REPORTED Diphtheria:			
46 States	1,740	1.331	I
92 cities	561	465	889
Measles:	301	300	009
45 States	5, 739	5, 959	İ
92 cities	1, 786	2, 050	
Meningococcus meningitis:	2, 100	2,000	
46 States	70	144	İ
92 cities	31 1	68	
Poliomyelitis:		•••	
46 States	38	69	l
Scarlet fever:		•••	
46 States	5, 243	5, 265	
92 cities	2,031	1, 968	1,411
Smallpox:		-,	
46 States	550	1, 375	
92 cities	22	100	43
Typhoid fever:	- 1		
46 States	237	150	
92 cities	29	21	28
DEATES REPORTED			
Industrial and an array and a	ì		
Influenza and pneumonia:			
	851	1, 530	
Smallpox:	_	_	
85 cities	0	1	
Omaha, Nebr	0	1	

City reports for week ended January 16, 1932

The "estimated expectancy" given for diphtheria, poliomyelitis, scarlet fever, smallpox, and tyhpoid fever is the result of an attempt to ascertain from previous occurrence the number of cases of the disease under consideration that may be expected to occur during a certain week in the absence of epidemics. It is based on reports to the Public Health Service during the past nine years. It is in most instances the median number of cases reported in the corresponding weeks of the preceding years. When the reports include several epidemics, or when for other reasons the median is unsatisfactory, the epidemic periods are excluded, and the estimated expectancy is the mean number of cases reported for the week during nonepidemic years.

If the reports have not been received for the full nine years, data are used for as many years as possible, but no year earlier than 1923 is included. In obtaining the estimated expectancy, the figures are smoothed when necessary to avoid abrupt deviation from the usual trend. For some of the diseases given in the table the available data were not sufficient to make it practicable to compute the estimated expectancy.

		Diph	theria	Influ	lenza			
Division, State, and city	re- esti- Cases		Cases reported	Deaths reported	Measles, cases reported	Mumps, cases re- ported	Pneu- monia, deaths, re- ported	
NEW ENGLAND								
Maine: Portland	11		2	1	0	134	0	7
New Hampshire:		,		•	ő	0	0	1
Concord Manchester	l ö	Ö	6		l ö	ŏ	ŏ	3
Nashua	Ŏ	Ó	Ó		0	0	0	0
Vermont: Barre		0						
Burlington	0	Ō	0		0	44	0	0
Massachusetts: Boston	84	35	17	8	5	8	19	12
Fall River	7	4	2	2	1	1	0	.2
Springfield Worcester	12 9	5 5	0		0	5 3	22 81	-2 3 6
Rhode Island:	-	-	_		0	0	0	0
Pawtucket Providence	0 10	2 7	0		Ö	642	ŏ	3
Connecticut:		-						
Bridgeport Hartford	1 4	6	0 3	1 1	1	1	0 42	3 3 3
New Haven	24	ĭ	ŏ		Ŏ	Ō	27	3
MIDDLE ATLANTIC								
New York:								
Buffalo New York	38 210	12 199	1 157	1 28	2 14	13 33	0 72	28 176
Rochester	8	6	1		0	57	24	3
Syracuse	25	2	0		0	9	11	7
New Jersey: Camden	8	8	7	1	1	1	1	3
Newark	59	17	4	5	0	1 1	37	13 5
Trenton Pennsylvania:	3	2	2		٠,	- 1	' '	9
Philadelphia	128	64	5	9	5	7	20 43	46
Pittsburgh Reading	53 23	19	9	1	4	139 2	15	12 8
Scranton	5		ŏ		Ŏ	1	i	
EAST NORTH CENTRAL				İ				
Ohio:			_			ا		
Cincinnati	10 154	9 30	5 13	1 26	2 0	0 176	0 110	9 20
Columbus	22 53	4	6		0 1	2	0	20 3
ToledoIndiana:	53	7	0	2	2	2	1	4
Fort Wayne		4						
Indianapolis South Bend	36	7	1 0		0	0	63	11 0
Terre Haute	5	i	ŏ		ŏ	ŏ	ŏ	3
Illinois:	112	104	48	62	2	46	3	45
Chicago Peoria	7		1 1	02	0 [0	0	5
Springfield	41	0	2 !		0 1	0 1	4 1	1

		Diph	theria	Infi	nenza .	١.,		_
Division, State, and city	Chick- en pox, cases re- ported	Cases, esti- mated expect- ancy	Cases reported	Cases reported	Deaths reported	Mea- sles, cases re- ported	Mumps, cases re- ported	Pneu- monia, deaths re- ported
EAST NORTH CEN- TRAL—continued	,							
Michigan: Detroit		_	33	Ι.			l	
Flint	68 19	53 3	1	1	3 1	9	17 74	29 8 0
Grand Rapids Wisconsin:	10	1	0		1	42	4	Ŏ
Kenosha	9	.0	0		0	.0	1	0
Milwaukee Racine	101 32	16 3	0		0	17 1	62 54	0 12 0 0
Superior	8	Ò	Ó		Ō	Ō	13	ŏ
WEST NORTH CENTRAL						•		
Minnesota: Duluth								
Minneapolis	19 41	0 16	0 5		0	0 8	0 47	1 7
St. PaulIowa:	18	5	Ŏ		ō	ĭ	2	
Davennort	2	1	2			0	1	:
Des Moines Sioux City	0	2	1 4			0	0	
Waterloo	4	ō	õ			ĭ	ő	
Missouri: Kansas City	35	6	14		o	2	1	•
St. Joseph	6	1	1		0	ō	0	1 4
St. Louis North Dakota:	15	41	12		0	1	2	11
Fargo. Grand Forks	4	0	Õ		0	26	0	0
South Dakota:	0	0	0			0	0	
Aberdeen Nebraska:	7	0	0			16	0	
Omaha	11	4	4		o	o	2	7
Kansas: Topeka	6	2	6		0	1	2	
Wichita	39	2	10		ŏ	6	ő	0
SOUTH ATLANTIC	1	1		1	1	1		
Delaware:	_	_	_			i		
Wilmington Maryland:	2	2	0		0	0	1	3
Baltimore	65	22	18	6	2	4	56	33
Cumberland Frederick	0	8	0		8	0	0	1
District of Columbia:	12	17	13		- 1	į.	1	_
Washington Virginia:	_	- 1			0	1 /	0	24
Lynchburg Norfolk	1 8	1 2	0		8	2 0 0 1	8	3
	0	6 (6].		1	ŏ	0	7
Roanoke. West Virginia: Charleston.	8	2	4 -		. 1	1	0	1
Charleston	7	1	1 .		0	6	0	1
Huntington Wheeling	Ž -	i	2			11	8 -	
North Carolina: Raleigh	i	1	l			- 1	- 1	•
ALTHUMEN COLL	1	i	·····i				·····	······i
Winston-Salem South Carolina:	4	1	2		8	8	0	Ī
Charleston	1	31	11	16	8	o l	9	
Columbia	Ì	81	8		0	8	8	0
Georgia: Atlanta	- 1					- 1	- 1	
Brunswick	8	18	81.	28	8	8	8	4
Savannah Florida:		. i						
Miami	2	31	3	1	• 1	9	8	1
Tampa	• j	1	0 _		81	1	ð	3

		Diph	theria	Infl	uenza			
Division, State, and city	Chick- en pox, cases re- ported	Cases, esti- mated expect- ancy	Cases reported	Cases reported	Deaths reported	Measles, cases reported	Mumps, cases re- ported	Pneu- monia, deaths re- ported
Bast South Central								
Kentucky: Covington Lexington Tennesses:	2	1	<u>i</u>	i	0	i	12	i
Memphis Nashville Alabama:	2	, 1	3		i	0	0	3
Birmingham Mobile Montgowery	4 0 1	5 1 1	5 1 0	3 1 2	1 1	0 0 1	5 0 14	5
WEST SOUTH CENTRAL								
Arkansas: Fort Smith Little Rock Louisiana:	0	0 1	1 1		1	0 2	0	6
New Orleans	0	13 1	17 3	6	4 0	0 18	0 1	9
Oklahoma: Muskogee Tulsa	4 3	2	3 4			0	1	
Texas: Dallas	5	9	12	-,	0	1	0	7
Fort Worth Galveston	5	7	9 6		0 1	2 0	0	7 4 1
Houston San Antonio	0 1	8 3	16 3		1 2	0	1 0	3 14
MOUNTAIN	.							
Montana: Billings Great Falls Helena Missoula	0 5 0	0	0		0 0 0	3 1 52 0	0 0 0	0 2 0 1
Idaho: Boise	0	o	0	-	0	0	1	0
Colorado: Denver Pueblo	14 15	8	5 0		9	4 0	10 0	15 0
New Mexico: Albuquerque	3	0	2		o	0	o	1
Arizona: Phoenix Utah:	0 -		1		0	0	o	3
Salt Lake City Nevada:	37	3	0		2	0	1	1
Reno	0	0	0		0	0	0	2
PACIFIC	I		ļ			ļ	1	
Washington: SeattleSpokaneTacoma	63 17 5	4 1 3	0		0	161 2 1	17 0 3	5
Oregon: Portland	26 5	8	1 0	1 3	0	1 0	6	6 2
California: Los Angeles	86	36	45	108	5	o	6	43
Sacramento San Francisco	13 49	3 14	1 4	1 14	5	89 33	ŏ	13 7

	Scarle	t fever		Smallpo	ox .	Tuber-	T;	rphoid f	ever	Whoop-	
Division, Stata, and city	Cases, esti- mated expect- ancy	Cases re- ported	Cases, esti- mated expect- ancy	Cases re- ported	Deaths re- ported	culo- sis, deaths	Cases, esti- mated expect- ancy	Cases re- ported	Deaths re- ported	cough, cases re- ported	Deaths, all causes
NEW ENGLAND											
Maine: Portland	8	1	0	0	0	0	1	0	0	6	32
New Hampshire: Concord	0	4	0	0	0	0	0	0	0	0	5
Manchester Nashua	2	2	Ŏ	ŏ	ě	ŏ	ŏ	ě	Ŏ	Ŏ 1	16
Vermont:	-	1		١	U	· ·		ľ	U	1	
Barre Burlington	0	2	0	2	0	1	0	0	0	·····o	ii
Massachusetts: Boston	92	152	o	1	0	22	0	0	0	32	
Fell River	4 9	2 5	Ö	0	Ŏ	0	1 0	Ŏ'	Ŏ	0	242 82 80 49
Springfield Worcester	14	36	ŏ	ŏ	ŏ	2	ŏ	ŏ	ŏ	15	49
Rhode Island: Pawtucket	2	0	0	o l	0	0	0	o	0	0	17
Providence Connecticut:	16	23	٥	٥١	0	3	1	٥١	0	21	89
Bridgeport Hartford	10	8	0	8	0	8	0	0	0	5 31	38 42
New Haven	ő	8	ŏ	ŏ	ŏ	ž	ŏ	ŏ	ĭ	3	42 47
MIDDLE ATLANTIC		ĺ	1	İ				l			•
New York: Buffalo	27	76	0	0	0	8	0	0	9	35	142
New York	227	404	0	0	0	101	7	6	0	151	1, 499
Rochester Syracuse	10 13	56 26	8	8	0	2	8	2 0	0	102	84 54
New Jersey: Camden	6	50	o	0	0	1	1	0		6	39
Newark Trenton	25	19	0	0	Ŏ	15	8	1 0	0	58 1	101 83
Pennsylvania:	97		0	- 1	1	_	2	1	1	- 1	
Philadelphia Pittsburgh	36	150 65	Ŏ	0	0	24 7	Ō	1 0	0	276 43	451 157
Reading Scranton	4	15	0	. 0	0	0	0	8	0	14	87
EAST NORTH CENTRAL											
Ohio: Cincinnati	22	50	0	0	اء	11	اه	اه	0	8	101
Cleveland	44	61	0	0	0	20	1	2	0	172	181 192
Columbus Toledo	12 14	13	1 2	8	8	8	8	0	0	23 62	82 68
Indiana: Fort Wayne	5 .		0 -				0				
Indianapolis South Bend	12 3	4	5	1	0	2 8	0	0	0	13 5	
Terre Haute	2	ő	ő	ŏ	ŏ	ő	ŏ	ŏ	ŏ	î	19 25
Illinois: Chicago	187	191	1	1	0	37	1	0	o	156	683
Peoria Springfield	8	10 -		0	8	8 -		8	0	10	683 21 24
Michigan: Detroit	108	151	- 1	1	1	22		- 1	0	113	
Flint	18	15	2	8	8	0	8	8	0	13	270 18 80
Grand Rapids_ Wisconsin:	13	10	0	- 1	- 1	1	٥	1	0	2	3 0
Kenosha Milwaukee	35	35	0	8	8	8	0	8	0	133	101
Racine	5	1	Š	8	8	8	8	8	ě	2	

	Scarle	t fever		Smallp	ox .	Tuber-	T	phoid i	lever	Whoop-	1
Division, State, and city	Cases, esti- mated expect- ancy	Cases re- ported	Cases, esti- mated expect- ancy	Cases re- ported	Deaths re- ported	culo- sis, deaths re-	Cases, esti- mated expect- ancy	Cases re- ported	Deaths re- ported	ing cough, cases re- ported	Deaths, all causes
WEST NORTH CENTRAL											
Minnesota: Duluth Minneapolis St. Paul Iowa:	11 47 29	4 34 14	0 1 1	0 0 0	0	1 1 3	0 1 0	0 0 0	0 0 0	3 7 7	15 83 54
Davenport Des Moines Sioux City Waterloo	3 8 3 2	11 6 0 1	1 2 0 1	0 1 6 0			0 0 0	0 0 0		0 0 3 8	33
Missouri: Kansas City St. Joseph St. Louis	18 3 44	25 0 25	0 1 1	0 0 1	0 0 0	4 0 13	0 0 1	0	0 0 0	44 0 69	91 27 288
North Dakota: Fargo Grand Forks South Dakota:	3 0	5 0	0	0	0	1	0	1	0	0	4
Aberdeen Nebraska: Omaha Kansas:	7	1 7	0 2	0 2	0	1	0	0	0	3 6	58
Topeka Wichita SOUTH ATLANTIC	3 5	0 1	0 1	0	0	0	0	0	0	12 5	15 35
Delaware: Wilmington Maryland:	6	5	0	0	0	1	0	0	0	2	32
Baltimore Cumberland Frederick District of Col.:	34 1 0	45 4 2	0	0	0	14 0 0	2 0 0	1 0 0	0	144 3 3	230 8 4
Washington Virginia: Lynchburg Norfolk Richmond	26 1 3 7	23 2 6 23	0	0	0 0 0	9 0 1 3	0 0 0	2 0 0	0	15 3 0 4	167 12 51
Roanoke West Virginia: Charleston Huntington	4 1	1 1 0	0	0	0	0	0	0	0	0 4 0	17 19 0
Wheeling North Carolina: Raleigh Wilmington	3 1 1	ĭ 0	0 1 0	ŏ	0	0	1 0 0	ŏ	0	6	26 9
Winston-Sal- em South Carolina:	2	7	1	0	0	1	0	0	0	13	16 14
Charleston Columbia Greenville Georgia:	1 0	0 1 0	0 0 1	0 0 2	0	0	0	0 0	0	1 0 1	-
Atlanta Brunswick Savannah Florida:	6 0 1	5 0	1 0 0	0	0	0	0	0	0	2 0	77 2
Miami Tampa EAST SOUTH CEN-	3	0	0	8	0	5 2	0	2	0	0	31 27
TRAL Kentucky: Covington	2		0				0				
Lexington Tennessee: Memphis	8 .	0	1	0	0	1	1	0	0	6	11
NashvilleAlabama: Birmingham Mobile	3 5 1	0 8 1	1 0	0	0	0 2 0	0	0 4 0	0 0	10 2 0	34 61 21
Montgomery	il	3 1	0	ŏ l.			Ŏ I	٥١.		4 '-	

	Scarle	t fever		Smallp	ox	Tuber-	T:	phoid i	iever	Whoop	
Division, State, and city	Cases, esti- mated expect- ancy	Cases re- ported	Cases, esti- mated expect- ancy	Cases re- ported	Deaths re- ported	culo- sis, deaths re-	Cases, esti- mated expect- ancy	Cases re- ported	Deaths re- ported	ing cough, cases re- ported	Deaths, all causes
WEST SOUTH CENTEAL											
Arkansas: Fort Smith Little Rock Louisiana:	1	0 2	0	0	0	i	0	0	<u>o</u>	8 1	
New Orleans Shreveport Oklahoma:	7	9	0 1	3 0	0	13 2	3 0	0	1 1	2 4	187 28
Muskogee Tulsa Texas:	2	0 5		0 1			0	0		4 2	
Dallas Fort Worth Galveston Houston San Antonio	7 8 1 2 2	11 11 0 7 0	1 2 0 3 0	0 2 0 2 0	0 0 0	0 1 4 7 4	1 0 0 0	1 1 0 0	1 0 0 1 1	1 0 0 0	59 39 16 76 83
MOUNTAIN											
Montana: Billings Great Falls Helena Missoula	2 5 1	0 0 0 4	0	0	0 0 0	0	0	0000	0 0 0	0	5 10 5 9
Idaho: Boise Colorado:	1	1	1	0	0	0	0	0	0	0	9
Denver Pueblo New Mexico:	18 1	28 0	1 0	0 1	0	7	0	0	8	2 1	80 12
Albuquerque Arizona: Phoenix	1	1	0	0	0	1	0	2	0	0 1	15
Utah: Salt Lake City_	5	2	0	0	0	0	0	0	0	1	41
Nevada: Reno	1	0	0	0	0	0	0	0	0	0	8
PACIFIC	İ										
Washington: Seattle Spokane Tacoma Oregon:	10 8 8	5 1 6	2 2 2	0	0	-	0	0	0	7 0 0	22
Portland Salem California:	8	4	7	12 0	8	1	0	0	0	3	85 12
Los Angeles Sacramento San Francisco	40 2 19	48 0 8	812	8 0 1	0	16 2 12	1 0 1	0	0	10 8 0	826 163

			1						
	CO	ningo- ccus ingitis	Leths cep	argic en- halitis	Pe	llagra	Polion tile	oyelitis paraly	(in fan- /sis)
Division, State, and city	Cases	Deaths	Cases	Deaths	Cases	Deaths	Cases, esti- mated, expect- ancy	Cases	Deaths
NEW ENGLAND									
Maine: Portland	0	0	0	0	0	0		1	
Massachusetts:						ľ		1	U
Boston	0	0	0	0	0	0	1 0	1	0
Connecticut:	-			-	-	1	1	i -	v
Hartford	1	0	0	0	0	0	0	0	0
MIDDLE ATLANTIC									
New York:									
Buffalo New York ¹	2 5	0 2	0 2	0 2	0	0	0	0	0
Pennsylvania:								1	•
Philadelphia Pittsburgh	1 2	0	1 0	1 0	0	0	0	0	0
EAST NORTH CENTRAL			-		-			٠	·
Indiana:									
IndianapolisSouth Bend	7 0	2	0	0	0	0	0	0	0
Illinois:			1	1	_	-	-	-	
Chicago	2	1	0	0	0	0	1	1	0
Detroit	4	1	0	1	0	0	0	0	0
Flint	2	0	0	0	0	0	0	0	0
SOUTH ATLANTIC			į						
Maryland:							_		_
Baltimore	0	0	1	0	0	0	0	0	0
Charleston 3	0	0	0	0	2	0	0	0	0
EAST SOUTH CENTRAL		ł		1	1				
Tennessee:		اء					_		_
NashvilleAlabama:	2	0	0	. 0	0	0	0	0	0
Birmingham	0	0	0	0	1	1	0	0	0
WEST SOUTH CENTRAL						- 1	j		
Louisiana:	ı	- 1			l	I	1	- 1	
New Orleans Texas: 2	2	2	0	0	0	0	0	0	0
Dallas	0	1	0	0	0	0	0	1	0
Houston	0	1	0	0	0	1	. 0	0	0
MOUNTAIN Arizona:		l		1			l		
Phoenix	0	1	0	o	0	0		0	0
PACIFIC		i	.	1	- 1	- 1			
California:		_		_ [_ [ا _ ا		_
Los Angeles San Francisco	0	1	0	0	0	8	0	0	0
		-1				٦ ا			

The following table gives the rates per 100,000 population for 98 cities for the 5-week period ended January 16, 1932, compared with those for a like period ended January 17, 1931. The population figures used in computing the rates are estimated mid-year populations for 1931 and 1932, respectively, derived from the 1930 census. The 98 cities reporting cases have an estimated aggregate population of more than 34,000,000. The 91 cities reporting deaths have more than 32,400,000 estimated population.

¹ Typhus fever: 1 death at New York City, N. Y.
² Dengue: 3 cases at Charleston, S. C., and 2 deaths at San Antonio, Tex.

Summary of weekly reports from cities, December 13, 1931, to January 16, 1932-Annual rates per 100,000 population, compared with rates for the corresponding period of 1930-31 1

DIPHTHERIA CASE RATES

	D	IPHTE	IERIA	CASE	RATES	3				
					Week	ended-				
	Dec. 19, 1931	Dec. 20, 1930	Dec. 26, 1931	Dec. 27, 1930	Jan. 2, 1932	Jan. 3, 1931	Jan. 9, 1932	Jan. 10, 1931	Jan. 16, 1932	Jan. 17, 1931
98 cities	103	2 94	72	71	3 72	80	483	81	4 88	74
New England Middle Atlantic East North Central West North Central South Atlantic East South Central West South Central Mountain Pacific	84 71 104 187 118 157 189 96 82	143 62 116 89 108 84 202 18	65 57 69 134 99 111 115 26 41	75 47 102 54 86 84 143 62 40	84 56 64 130 71 107 129 44 11 64	116 68 91 83 62 72 136 62 55	79 50 76 131 114 162 204 4 136 65	79 63 96 98 85 117 142 35 61	* 87 82 7 68 106 * 95 * 98 195 43 97	91 50 91 82 60 70 106 52 47
		MEA	SLES	CASE	RATES					
98 cities	128	2 194	126	181	192	281	4 301	351	• 279	324
New England Middle Atlantic East North Central West North Central. South Atlantic East South Central West South Central West South Central Nountain Pacific	637 79 60 25 26 52 44 740 294	271 87 28 1,416 138 275 18 167 6	945 66 32 50 14 17 41 339 259	305 70 27 1, 277 124 323 24 229 16	1, 207 93 93 38 79 • 31 64 513 11 445	268 101 55 1,894 322 921 24 317 24	1, 706 146 142 157 53 17 43 41, 530 784	490 178 62 2, 156 435 869 20 226 33	1, 916 116 7 182 78 8 35 10 9 73 517 544	310 158 87 1, 829 500 1, 004 7 374 55
98 cities	214	2 234	187	222	226	231	4 274	277	* 317	316
New England Middle Atlantic East North Central West North Central South Atlantic East South Central Most South Central Mountain Pacific	438 202 264 138 201 157 101 261 94	351 208 306 279 208 197 173 300 83	389 205 227 126 107 157 41 113 61	353 190 285 246 178 341 59 379 85	539 240 233 115 221 119 108 209 11 109	327 229 261 238 262 299 108 220 73	549 286 298 229 227 225 69 4 351 141	433 242 363 297 277 399 68 322 73	* 596 380 7 335 220 * 247 * 109 99 259 129	539 282 398 321 305 470 129 331 73
	1	SMALI	POX	CASE	RATE	3				
98 cities	5	19	4	7	*3	7	•6	13	43	16
New England Middle Atlantic East North Central West North Central South Atlantic East South Central West South Central West South Central West South Central Pacific	55 0 4 4 0 0 3 0 2	0 0 6 48 0 0 2 15 115 10	14 0 4 10 0 0 7 0 8	0 0 2 43 0 0 17 35 20	12 0 7 4 0 0 0 0 9	0 0 5 46 0 17 9	26 0 1 6 0 23 26 411 19	0 0 15 63 2 6 37 9	6 2 0 7 1 17 6 0 10 0 16 9 8	0 0 10 98 0 18 27 78 29

¹ The figures given in this table are rates per 100,000 population, annual basis, and not the number of cases reported. Populations used are estimated as of July 1, 1932, and 1931, respectively.

2 Shreveport, La., not included.

2 Covingtom, Ky., and Spokane, Wash., not included.

4 Salt Lake City, Utah, not included.

5 Barre, Vt., Fort Wayne, Ind.; Raleigh, N. C.; Savannah, Ga.; Covington, Ky.; and Memphis, Tenn., not included.

6 Barre, Vt., not included.

7 Fort Wayne, Ind., not included.

8 Raleigh, N. C., and Savannah, Ga., not included.

9 Covington, Ky., not included.

9 Covington, Ky., and Memphis, Tenn., not included.

10 Covington, Ky., and Memphis, Tenn., not included.

11 Spokane, Wash., not included.

Summary of weekly reports from cities, December 13, 1931, to January 16, 1932— Annual rates per 100,000 population, compared with rates for the corresponding period of 1930-31—Continued.

TYPHOID FEVER CASE RATES

					Week	ended-				
	Dec. 19, 1931	Dec. 20, 1930	Dec. 26, 1931	Dec. 27, 1930	Jan. 2, 1932	Jan. 3, 1931	Jan. 9, 1932	Jan. 10, 1931	Jan. 16, 1932	Jan. 17, 1931
98 cities	5	18	6	7	* 5	5	14	4	5	Ž
New EnglandMiddle Atlantic	7 5	10	2	2 3	12	2	2 5	5 2	• 0	
Rast North Central West North Central	1 0	9 8	2	12 6	4 2	4 2	2 2	2	72	
South Atlantic	10	12	14	16 18	6	4	8	10	8 14	10 53
East South Central	10 23 34	36 126	12 44	18	138 3	48 3	0 13	12 20	10 36 10	53 14
Mountain	0	20	70	9	Ō	18	411	17	10	
Pacific	2	6	4	6	11 8	6	4	2	Ô	

INFLUENZA DEATH RATES

91 cities	8	2 10	9	11	• 13	16	4 18	24	• 13	36
New England Middle Atlantic	5	2 5	7	2 10	2 5	7	10 12	5 29	• 17 12	10 59
East North Central West North Central	6	10 15	5 3	7 9	10 9	7 3	14	12 21	7 5	9 18
South Atlantic East South Central	12 6	20 32	12 32	24 19	18 27	20 26	35 31	28 45	8 8 10 31	42 64
West South Central Mountain	17 17	³ 23 18	24 70	32 0	45 131	93 18	30 4 125	76 44	30 103	79 35
Pacific	14	10	7	17	14	10	23	22	26	10

PNEUMONIA DEATH BATES

91 cities	106	1111	101	126	• 121	164	1144	187	▶ 126	219
New England	111	116	94	119	91	160	165	113	6 104	159
Middle Atlantic	116	127	101	126	126	184	148	233	133	311
East North Central	63	69	77	94	84	103	104	110	7 82	124
West North Central	103	96	118	117	103	180	131	200	119	212
South Atlantic	142	138	132	174	174	230	196	267	206	237
East South Central	120	110	113	149	151	207	169	267	10 156	229
West South Central	142	2 135	131	189	152	199	128	238	148	228
Mountain.	200	220	226	194	165	264	4 329	244	181	270
Pacific	122	127	89	135	175	135	167	134	158	118

² Shreveport, La., not included.
³ Covington, Ky., and Spokane, Wash., not included.
⁴ Salt Lake City, Utah, not included.
⁵ Barre, Vt.; Fort Wayne, Ind.; Raleigh, N. C.; Savannah, Ga.; Covington, Ky.; and Memphis, Tenn., not included.
⁶ Barre, Vt., not included.
⁷ Fort Wayne, Ind., not included.
⁸ Raleigh, N. C., and Savannah, Ga., not included.
⁹ Covington, Ky., not included.
¹⁰ Covington, Ky., and Memphis, Tenn., not included.
¹¹ Spokane, Wash., not included.

FOREIGN AND INSULAR

CANADA

Quebec Province—Communicable diseases—Week ended January 16, 1932.—The Bureau of Health of the Province of Quebec, Canada, reports cases of certain communicable diseases for the week ended January 16, 1932, as follows:

Disease	Cases	Disease	Cases
Cerebrospinal meningitis	1 180 55 7 2 322	Mumps Poliomyelitis Scarlet fever Tuberculosis Typhoid fever Whooping cough	88 6 110 44 21 48

LATVIA

Communicable diseases—October, November, 1931.—Cases of certain communicable diseases were reported in Latvia during the months of October and November, 1931, as follows:

	C	ases		Ct	3.Ses
Disease	Octo- ber	Novem- ber	Disease	Octo- ber	Novem-
Anthrax Botulism Cerebrospinal meningitis Diphtheria Erysipelas Influenza Leprosy Measles	1 2 67 28 62 3 11	7 79 16 101	Mumps Poliomyelitis Puerperal septicemia Scarlet fever Tetanus Trachoma Typhoid fever Whooping cough	58 9 35 4 83 88 57	116 3 14 62 2 101 66 71

PHILIPPINE ISLANDS

Manila—Rat bite fever.—According to information dated January 15, 1932, there was a mild outbreak of rat bite fever in Manila, P. I. Eight cases were identified bacteriologically, and it was thought that there were probably many more unrecognized cases. The distribution of the disease was said to be the same as the former distribution of plague.

CHOLERA, PLAGUE, SMALLPOX, TYPHUS PEVER, AND YELLOW PEVER

CHOLERA

From medical officers of the Public Health Service, American consuls, International Office of Public Hygiene, Pan American Sanitary Bureau, health section of the League of Nations, and contracts. The reports contributed in the following tables must not be considered as complete or final as regards either the list of countries the many of the figures for the page.

ior the particular countries for winds reports are given.		[C in	dicates c	[C indicates cases; D, deaths; P, present]	deaths;]	, prese	rf]										
										Week ended	pepa						
. Place	June 28- July 25, / 1981	July 26- Aug. 22, 1981	Aug. 23- 36pt. 19,	Sept. 20- Oct. 17, 1931	October, 1931		Š	November, 1931	1931		Å	December, 1931	1931		January, 1932	13, 18	g
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India (Portuguese)	201-			258	<u>∞</u> ∞	80	8-		000		H	H	\mathbb{H}	$\frac{\cdots}{1}$	H	Π	

CHOLERA, PLAGUE, SMALLPOX, TYPHUS FEVER, AND YELLOW FEVER-Continued

CHOLERA—Continued

[C indicates cases; D, deaths; P, present]

																	1
									P	Week ended	pepu						
Place	June July 25, 1931	July 26- Aug. 22, 1931	Aug. 23- 19, 19,	Sept. 20- 20- Oct. 17, 1931	October, 1931	- er,	ž	November, 1931	1881		Ď	December, 1931	1831		January, 1932	73, 198	23
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Indo-China (see also table below): Cochin-China—Rachgia Prompean	9	ተሪ	2	1		2	8										
Saigon and Cholon	647	1	7							$^{++}$	+	<u>A</u>	-	+			
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Indo-China (French) (see also table above): Cambodis 1. Cochin-China 1.	DCDC 11088	8 241 0 143 6 423	2282	4. 4. 8. 8. 8. 8. 8.	1112	16 2 2 2	8111	w 104			∞ ×0	0-00	

1 On Oct. 23, 1931, cholers was reported at Mohammerah, Abadan, and Ahwas, Persia. During the period from Oct. 22 to Nov. 7, 1931, 141 cases and 97 deaths were reported.

* Figures for cholers in the Philippine Islands are subject to correction.

* Reports incomplete.

CHOLERA, PLAGUE, SMALLPOX, TYPHUS FEVER, AND YELLOW FEVER—Continued

PLAGUE

[O indicates cases; D, deaths; P, present]

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									W 86	Week ended-	f						
Place	July 26- Aug. 22, 1931	Aug. 23- Sept. 19, 1931	Sept. 20- Oct. 17, 1931	October, 1931	¥	ž	November, 1931	, 1931		D D	December, 1931	1931		Jen	January, 1932	8	
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D Belgian Congo				İ	Ħ	Ħ	$\dagger \dagger$	<u> </u>	-	67	╫	$^{+}$	H		-	$\dagger \dagger$	
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Canary Islands: Palma Island—Los Llanos	88	202	270	8	T	25	28			<u>; ;</u> g ;	+	╁	$^{++}$	$\frac{11}{11}$	$\frac{1}{11}$	00	
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1 On July 27, 1931, 1, 220 cases of plague were reported in Chiobe and Changchow, China, since April. On Sept. 19, 1931, 18 deaths were reported in Changchuanpu and new cases in Kaitung and Fengine and Fengine State and Fengine Province, China, with 2,000 deaths at Hsinghsien.

* On Oct. 17, 1831, plague epidemic was reported in western Shansi Province, China, with 2,000 deaths at Hsinghsien.

CHOLERA, PLAGUE, SMALLPOX, TYPHUS FEVER, AND YELLOW FEVER—Continued

PLAGUE-Continued

[C indicates cases; D, deaths; P, present]

									W	Week ended-	- Pa				:		1
Place	July 26- Aug. 22, 1931	Aug. 23- Sept. 19, 1931	Aug. Sept. 22- 20- Sept. Sept. October, 19,1931 1931	Octo 199	ber,	ž	vemb	November, 1931		Ă.	сешре	December, 1931		.gr	January, 1932	1962	
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Peru (see table below). Senegal (see table below).)	•	'	<u>' </u>											
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Spain: Hospitalet—Barcelona Province	200	-10	- 12	10-		-	-	İ	+		İ	+		İ			
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Orange Free State	7	P	Ъ				ÌŢ	ď									
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Place	June, 1931	July, 1931	Au- gust- 1931	Sep- tem- ber, 1931	Octo- ber, 1931	No- vem- ber, 1931	Place	June, 1931	July, 1931	Au- gust, 1931	Sep- tem- ber, 1931	Octo- ber, 1931	No vein-
British East Africa (see also table above): Kenya. Kenya. Carlamor Parish—Los Hoyos. Carlamora. C	2	26 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1	8	4 8 11 12 13 14 14 17 18 18 18 18 18 18 18	2 82 1 1 2 7-1 80 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1	4	Peru—Continued Huancabamba—Ayacaba C Huancabamba—Ayacaba C Huanra—Chancay C Lia Samana—Hualgayoc C Lima—Lima (haclendas) D Paijan—Trujillo D Patian—Trujillo D Patian—Trujillo D Patian—Trujillo D Patian—Trujillo D Patian—Trujillo D Patian—Trujillo D Patian—Trujillo D Patian—Huancabamba D San Pedro—Pacasmayo C C C D D D D D D D	488 488 <u>1</u> 28888	22322 8-1% ST-88	102 82 4 4 5 5 5 5 5 5 5 5 5 5 5 5 5 5 5 5 5	11 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1	PO H H 0044 H00H0	4444446
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¹ Reports incomplete.

CHOLERA, PLAGUE, SMALLPOX, TYPHUS FEVER, AND YELLOW FEVER-Continued

SMALLPOX

[C indicates cases; D, deaths; P, present]

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										₩ 86k 6	Week ended-	·					.
Place	June 28- July 26, 1931	July 26- Aug. 22, 1981	Aug. 28- Sept. 19, 1931	80pt. 20-Oct. 17, 1931	October, 1981	ber,	ž	November, 1931	1831		Dece	December, 1931	1931	·	annar	January, 1982	
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CHOLERA, PLAGUE, SMALLPOX, TYPHUS FEVER, AND YELLOW FEVER-Continued

SMALLPOX—Continued

[O indicates cases; D, deaths; P, present]

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										Week ended-	-pep					
Place	June 28- July 26, 1931	July 26 Aug. 22, 1931	Aug. 28- Bept. 19, 1931	Sept. 20-Oct. 17, 1931	October, 1931	¥	ž	November, 1981	, 1981		Decet	December, 1981	- E	Jan	January, 1932	28
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Jamaica Japan: Yokohama. Mexico (see also table below):	<u> </u>	Ш						$^{+}$	<u> </u>	<u> </u>	<u> </u>	<u> </u>	<u> </u>	<u> </u>	-	
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	181					-		-	! 	<u> </u>	•			Ц		
Portugal: Lisbon	*	33	8	3	2	91	17	8	8	8-	- 	- 8	*	8	8 -	
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	Δ.	November, 1931	11-20	87	June, 1931	89 -
	<u>a</u>	No	1-10	84	May, 1931	49 1,345
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	<u>a</u>	1931	21-30	.2 8		Morocco Rumania Turkeys Union of S
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e ::::::::::::::::::::::::::::::::::::	0,0,	September, 1931	1-10	1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1	Novem- ber, 1931	8 5
		st,		55-4	Octo- ber, 1931	
	64 Pr Pr	August,	193		Sep- tem- ber, 1931	4
	D-D-	July,	1831	∞-#	Au- gust, 1931	
AOOOA 	00000 00 0	June,	1931	47	July, 1931	81-8
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(see table below).	rleans fr from Jabana,				May, 1931	≅- 2 ∞-
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Rumania (see table below). Slam. Spain. Spain. Straits Settlements Sudan (Anglo-Egyptian). Turkay (see table below). Turkay (see table below). Turkay (see table below). Turkay (see table below). Turkay (see table below). Turkay (see table below). Turkay (see table below). Turkay (see table below).	Cape Province. Cape Province. Natal. Orange Free State Transval On vessel: Brazilian ship Jaboatao at New Orleans from Brazil. S. Talf (pigrin ship) at Stakin from Jedah. S. S. Hellasco at Mobile, from Habana, Cuba, and	Place		Indo-China (see also table above) Ivory Coast Syria: Beirut		China: Harbin Chosen France Greece Greece Maxico (see also table above)
Rumania (see table below). Spain. Spain. Straits Settlements. Straits Settlements. Straits (see table below). Turkey (see table below). Turkey (see table below). Turkoy (see table below). Turkoy (see table below). Turkoy (see table below). Turkoy (see table below).	Cape Province Cape Province Oratel Crange Free State Orange Free State Orange Free State Free State Orange Free State Fre			ee also t	Place	China: Harbin
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Stamania Siam	Oration Oratio			Indo-China (see also Ivory Coast		China: Chosen France. Greece. Mexico
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¹ Imported case.

CHOLERA, PLAGUE, SMALLPOX, TYPHUS FEVER, AND YELLOW FEVER-Continued

TYPHUS FEVER

[C indicates cases: D. deaths: P. present]

	_		C indicates cases; D, deaths; P, present	D, des	. F.	presen	,										
	į									Week	Week ended-						ı
Place	147 25. 1831	July 26- Aug. 22 1931	July 26- Aug. 23- Aug. 22, Sept. 19, 2 1931 1931	80t. 20-0st. 17, 1931.	October, 1981	l ter,	ž	November, 1931	, 1831		Dece	December, 1831	183		January, 1933	1983	
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San Luis Potosi Torreon Morocoo				- 	548	<u>2</u> ∞	→	00 (NO)44	0 mm-			-	- =	- 6	64			
Palestine Paraguay: Asuncion Peru ¹ Poland					20 m	 = &°	7- =	6 4.		H 600		38	32	n g	-n &	30 1	- g-	
Portugal: Oporto				20000	0 1100	n 80	E-16	38 13 33 33	-626	N 600	3 -8	-	50	ନ୍ଧ୍ୟ		1111		
Turkey (see table below). Union of Sooislies Soviet Republics (see table below) Union of South Africa: Cape Province Municipality of East London Natal.	see table	e below	. !!!	0000	р, р.	<u> </u>	рын			P1 P1	ρι	р. р.	<u>α</u> , α,					
Orange Free State Transvaal Yugoslavia (see table below).					P P P	ы	Ы	4 T	0,0,	<u>:</u>	P. P	P	d					
Place	May, 1931	June, 1931	July, 1931	Au- gust, 1931	Sep- tem- ber, 1931	Octo- ber, 1931	No- Vem- 1931		ä	Place		May, 1931	June, 1931	July, 1931	Au- gust, 1931	Sep- tem- ber, 1931	October, ber, 1931	No vem- ber, 1931
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Typhus fever has been reported in Peru from May to November, 1831, 183 new cases being reported during the months of October and November. The disease has not spread to the coastal regions.

CHOLERA, PLAGUE, SMALLPOX, TYPHUS FEVER, AND YELLOW FEVER-Continued

YELLOW FEVER

[C indicates cases; D, deaths; P, present]

	May	June	July	Α116.				l		Week ended-	ded-					
Place	31- June 27,	क्ष्में इंद्रेड़	Aug. 22,	Sept.	Sept.		Octo	October, 1931	ដ		ž	November, 1931	er, 193		December, 1931	ber,
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Senegal: Podor (Hinterland)										- :				:
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