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THE OCCURRENCE OF TULARÆMIA IN BRITISH COLUMBIA

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The first diagnosed case of tularæmia in Canada was reported by McNabb, in February, 1930,¹ in a miner aged 34, living near Timmons, Ontario. Incidentally, it was also the first evidence that tularæmia was resident in the native fauna of Canada. In the present paper there is reported the recovery of *Bacterium tularense* McCoy and Chapin from a snowshoe rabbit (*Lepus americanus columbiensis* Rhoades), near Vavenby, British Columbia, in May, 1930. These two occurrences of tularæmia, the former in Ontario 400 miles north of the United States border, the latter in British Columbia over 200 miles north of the border, and the two localities over 1,500 miles apart from east to west, suggest the likelihood that tularæmia in the Canadian fauna is a widespread infection of many years' standing.

In the spring of 1930 one of us (Parker), upon request of the Dominion entomologist, Mr. Arthur Gibson, was detailed by the Surgeon General of the United States Public Health Service to visit British Columbia for the purpose of conferring with Hearle and Bruce concerning the occurrence of ticks and tick-borne infections in that Province. Incident to a trip into south central British Columbia during late April and early May, a snowshoe rabbit was autopsied which had an enlarged spleen and of which the liver showed lesions suggestive of tularæmia. This rabbit had been found near Vavenby, too weak to resist capture, by a local rancher and amateur naturalist, Mr. T. A. Moiliett, who, because of the animal's heavy infestation with ticks (*Haemaphysalis leporis-palustris* Packard), forwarded it to Hearle's laboratory at Kamloops.

Later, at the Public Health Service laboratory at Hamilton, Mont., some of the ticks and portions of the rabbit's spleen and liver (preserved in glycerin) were tested for the presence of *Bacterium tularense*.

¹ McNabb, A. L.: Tularæmia; The First Case Reported in Canada. Capadian Public Health Journal, vol. 21, February, 1930, pp. 91–92.

Three tick-injected guinea pigs and two of three injected with liver emulsion died in two to four days, and on necropsy all showed lesions typical of tularæmia. One liver- and three spleen-injected guinea pigs remained apparently well and when killed and autopsied were either negative or showed poorly defined lesions. Cultures recovered from one of the tick- and one of the liver-injected guinea pigs, when used as antigen, were agglutinated in high titer by known tularæmia immune sera, which similarly agglutinated a known *tularense* antigen.

In addition to these definite data, information suggestive of the occurrence of tularæmia in other parts of Canada was secured from Mr. Alex Dennis, of the Canadian Entomological Service, at Vernon, and from Maj. Allan Brooks, of Okanagan Landing. The former stated that in 1921, when living at Salmon Arm, British Columbia, he had killed snowshoe rabbits whose livers were "spotted"; and the latter said that when residing in Alberta "rabbit cycles" had been a familiar phenomenon, and that during epidemic years there was always an unusual amount of sickness among the resident settlers by whom jack rabbits were commonly used as food.

EFFECT ON LIFE INSURANCE MORTALITY RATES OF RE-JECTION OF APPLICANTS ON THE BASIS OF MEDICAL EXAMINATION

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The subject of the duration of medical selection due to rejections by life insurance companies on account of poor physical condition or disease has been discussed on many occasions in actuarial literature, but has not been taken up very often in publications relating to public health and vital statistics, although it is an important factor in the interpretation of life insurance mortality. Let that be the excuse for recurring at this time to a subject which was termed "threadbare" by George King in 1878.¹ It is always well, furthermore, to check up on previous conclusions whenever new data become available, especially since the relation between actuarial mortality and that of the general population appears to have changed greatly in the last half century.

The factor of selection which originally received particular attention was "the selection which the assured exercise against the companies by dropping policies on healthy lives and retaining those on lives which have become bad or doubtful."² However, by 1870, the

¹ In the discussion of a paper by T. B. Sprague read before the British Institute of Actuaries. Journal of the Institute and Assurance Magazine, January, 1879, Vol. XXI, p. 253.

² On the Value of Selection as Exercised by the Policyholder Against the Company. John Adams Higham. Read before the British Institute of Actuaries, Mar. 31, 1851. The Assurance Magazine, vol. 1, No. III, April, 1851, p. 190.

importance of selection due to the medical examination was already recognized. For instance, at that time Sprague³ stated: "It is universally acknowledged that the rate of mortality among assured lives is very light during the first few years that follow the grant of the assurance; being extremely small in the first year and gradually increasing until, after the lapse of a greater or less number of years, the mortality becomes, according to some authorities, equal to that indicated by tables deduced from the population at large and according to others still heavier. This is, of course, satisfactorily explained by the medical examination of the lives proposed for assurance, which has the effect of eliminating those persons who are suffering from such acute or chronic diseases, dangerous to life, as can be detected by the medical officers of the assurance companies." The subject of withdrawals, however, continued to be the major topic of discussion with reference to selection until rather recently. At the present time discussions of selection relate rather to that due to the medical examination, including reports on the family history, personal history, and habits of persons applying for insurance (recently termed "temporary selection" by Elderton⁴), and a permanent force due to the class of lives involved ("permanent selection"). Without expressing an opinion as to whether the withdrawals ever did have an appreciable effect on the assured mortality rates, I believe we can follow Elderton's lead and disregard this phase of the subject. In passing, it may be said, however, that previously more reasonableness attached to the view that withdrawals did constitute a factor of selection against the company, because the mortality rate among insured lives was believed to be as high as or higher than that in the general population. As is well known, the contrary is true to-day for "ordinary" policyholders, a point which will be referred to later.

Henry Moir, in 1919,⁵ stated that: "More recently it has been urged that withdrawals do not have the effect of reducing the proportion of healthy lives; indeed, the direct contrary is sometimes accepted on the ground that withdrawal from a company in good standing is more frequently a result of financial embarrassment or irregular habits."

The duration and total effect of "temporary selection" on the mortality of assured persons have become important points in the

³ On the Rate of Mortality Prevailing Among Assured Lives, as Influenced by the Length of Time for Which They Have Been Assured. Thomas B. Sprague. Journal of the Institute of Actuaries and Assurance Magazine, Vol. XV, Part V, April, 1870, p. 328.

⁽a) Report on the Results of an Investigation of the Mortality Experience of Life Annuitants During the Period 1900-1920. W. Palin Elderton and H. J. P. Oakley. Journal of the Institute of Actuaries, Vol. LIV, Part I, p. 43, March, 1923.

⁽b) Notes on the Interpretation of "Select" Rates of Mortality. W. Palin Elderton and H. J. P. Oakley. Journal of the Institute of Actuaries, Vol. LV, Part I, p. 1, March, 1924.

⁴ Sources and Characteristics of the Principal Mortality Tables. Henry Moir. Actuarial Studies No. 1. ⁴ Published by the Actuarial Society of America, 1919, p. 44.

minds of actuaries in this country, following the adoption a few years ago by certain Canadian companies of the principle of insuring persons without medical examination.⁶ It is the general feeling that most of the selection wears off in the course of two years, but that a residue remains for some years. There is a great difference of opinion in regard to the matter. Henry Moir states that "it is the general opinion that the effects of the first selection never entirely disappear."⁷ On the other hand, Elderton believes that the period of "temporary selection" has frequently been overestimated because of the gradual decrease with time in the mortality rates analyzed.⁸ Thus there would seem to be a place for a further analysis of assured mortality data from this point of view. A measure of the degree of the selection in terms of mortality has not been completely worked out and is, no doubt, a changing element. For instance, Moir states that "the influence of medical selection is more persistent, and especially conspicuous amongst younger men." It might be expected that advance in medical science, with the development of urinalysis and other laboratory technique, would tend to make the selection more far-reaching in its effect.

One will realize that there is a corresponding selection in the case of annuitants, but due to a quite different cause. Persons who do not believe themselves to be in good health are not likely to take out annuity policies. Here again the maximum effect of this selection will be found in the early years of the policies. As the present study was not concerned with annuitants, no data on this phase of selection is included.⁹

The material for the present study is based on a joint investigation on occupational mortality by the Actuarial Society of America and the Association of Life Insurance Medical Directors.¹⁰ As a basis for the occupational comparisons, data were secured for ordinary business^{10a} during the years 1915–1926, involving \$546,357,000 in death claims. It should be noted that these data were based on the amounts insured, rather than on policies (the unit being taken as \$1,000, about the amount of the average policy), but it was not be-

⁶(a) Life Insurance Without Medical Examination. D. E. Kilgour. Transactions of the Actuarial Society of America, May 19 and 20, 1921. Vol. XXII, Part 1, p. 120.

⁽b) Actuarial Note. Insurance Without Medical Examination. Savings in Expense Compared With Expected Extra Mortality. Arthur Hunter. Transactions of the Actuarial Society of America, May 19 and 20, 1921. Vol. XXII, Part 1, p. 140.

⁷ Op. cit., p. 44.

^{*} Elderton and Oakley. Op. cit., 1924.

^{*} See Elderton and Oakley, 1923, Op. cit., for a recent study of selection in the case of annuitants.

¹⁰Joint Occupation Study: 1928. Compiled and published by the Actuarial Society of America and the Association of Life Insurance Medical Directors. New York: 1929. The chairman of the joint committee is Arthur Hunter, to whom grateful acknowledgment is made for review of the present paper.

¹⁴ Exclusive of industrial insurance where premiums are paid weekly or monthly.

lieved that this would result in any marked differences.¹¹ The data were graded to form tables of mortality rates from which to calculate the expected number of deaths in any occupation; but for the present purpose it seems preferable to employ the ungraded data (number exposed to risk and number of deaths by age of the policy and age at entry) to avoid possibility of errors entering into the calculations because of the method of grading. The number exposed to risk and the number of deaths are given in two appendix tables. The data were secured from 10 large insurance companies.

The basic data are given in the joint report in two ways: First, for the whole period 1915–1926 (being carried to the anniversaries of the policies in 1927), and second, for a part of this period for which some additional data were available, 1920–1926. It is a significant point that only policies *taken out* during the two periods are included, so that the maximum length of policy for the total period is 12 years and for the period 1920–1926 is seven years.

In determining the most logical way of handling the material, it was felt that the first consideration was the elimination of the effect of the influenza epidemic of 1918-19, since this not only greatly increased the rates, but exerted its influence mostly among young adults. The second period (1920-1926) was almost free from this effect, but had the unfortunate difficulty of being only seven years in duration. Furthermore, it was evident that in this case all of the deaths occurring during the seventh year of the policies would be in 1926, and all of the deaths during the sixth year of the policies would be in 1925 or 1926-i. e., at the very end of the period and based on relatively small numbers.¹² It seemed best, therefore, to use the 1920–1926 data for the first four years of the life of the policies, and the 1915-1926 data for the succeeding eight years. The effect of the epidemic was eliminated in this way; since even when the full period was used all deaths must have occurred after 1919, data for less than 4-year policies being used only for the period 1920-1926.

[&]quot;The point should be made, however, that the deaths are based upon death claims actually paid. Rejection of claims in the first year would therefore be one of the factors included in "temporary selection" as understood in this paper.

Quotation is made from the Joint Report in regard to the use of amounts, as follows:

[&]quot;Material for the mortality rates was furnished by each company in the same form as was used in the occupational classes. The latter were derived by policies, as the committee were satisfied that the morcality by policies would not differ essentially from that by amounts, especially in view of the small average policy in this investigation. In the case of the basic tables, however, the material had already been prepared in some of the companies by amounts for dividend purposes and in several of these institutions it would have been very laborious to obtain the exposures and deaths by policies. The data for amounts insured were therefore used for the basic tables." These are the data employed in the present analysis.

¹⁹ Similarly, of course, for the 1915-1926 data, the deaths during the twelfth year would also be in 1926, etc., but after the policies had been in force for so long a period as this, the lack of precise data did not seem of importance (even if they could have been secured).

Recent investigations have brought out certain difficulties in determining the duration of selection due to the inclusion of data covering a long period of time during which the mortality rate and other relations may be changing. It is of interest to quote the following from Elderton:¹³

The simplest safeguard against misstating the period of selection probably lies in frequent investigations and the examination of the statistics obtained. We are inclined to take the view that by making past investigations over a long period of years in order to get a mass of data and thus reduce "errors of observation," we have introduced persistent errors which are of greater importance and have created for ourselves the inconvenience of showing temporary selection for a longer period than is justified by the statistics or necessary for calculations depending on an assumed future mortality. We feel that the true period of temporary selection can only be ascertained with certainty by the examination of homogeneous facts, and while we recognize that actuaries have always attached importance to homogeneity, we believe that there are more factors involved than has sometimes been assumed and that one of the most important of them is time.

It can be shown that the present material is reasonably free from such difficulties, especially since the total period (1915–1926) has only been used for policies of five years or more duration and therefore only deaths occurring during 1920–1926 have been used in the calculations. The mortality rate in this country has shown steady improvement in the past, but during the years 1920 to 1926 remained at a constant level. The crude mortality rates for these years in the total registration area were: 1920, 13.0 per 1,000; 1921, 11.6; 1922, 11.8; 1923, 12.3; 1924, 11.8; 1925, 11.9; 1926, 12.3.¹⁴

A fundamental difficulty in the present analysis will lie in the fact that the material is given according to the age at issuance, not the true age. For instance, Table 1 presents the annual death rates according to the age of the policy and age at issuance, and the reader will see at once that a directly vertical comparison is not possible, since persons classed in the age group 15 to 19, but with sixth year policies, were really in the age group 20 to 24. In other words, one year is added to the age for each step down the table. This fact is indicated by the figures in boldface, three age groups having been selected to emphasize the point. As the data were collected from the insurance companies in 5-year age groups, no direct correction is possible.

¹⁸ Elderton and Oakley, Op. cit., 1929.

¹⁴ In the spring of 1920 there was a recurrence of the influenza epidemic; but this could have affected only a relatively few assured persons—i. e., those taking out policies between Jan. 1, 1920, and the date of the wave of influenza and those taking out policies in the corresponding period of 1915.

	Age at issuance of policy											
Policy year	15 to 19	20 to 24	25 to 29	30 to 34	35 to 39	40 to 44	45 to 49	50 to 54	55 to 59	60 to 64	65 and over	
1920-1926												
First Second Third Fourth 1915–1926	1. 96 2. 21 2. 14 2. 31	1.76 2.06 2.28 2.26	1, 56 1, 89 2, 19 2, 35	1.76 2.26 2.56 2.88	2.34 3.67 3.81 3.71	2. 94 4. 07 5. 08 6. 43	4. 43 5. 65 7. 39 7. 79	6.06 9.17 11.47 10.55	10. 49 13. 88 17. 74 16. 93	14.66 18.84 17.52 21.65	19. 28 27. 09 46. 28 30. 96	
Fifth	2. 54 2. 68 2. 64 2. 54 2. 66 2. 52 2. 52 2. 70	2. 43 2. 37 2. 30 2. 22 2. 66 2. 62 2. 92 2. 33	2.79 2.75 2.61 2.85 3.10 3.42 3.10 3.57	3. 33 3. 42 3. 49 3. 82 4. 15 4. 01 4. 59 4. 69	4. 16 4. 72 4. 88 5. 87 6. 14 6. 14 7. 95 8. 01	5. 78 7. 38 7. 25 8. 35 10. 46 9. 51 9. 21 11. 19	9. 10 11. 15 10. 55 11. 11 13. 07 13. 32 18. 90 16. 47	13. 30 14. 89 17. 92 19. 01 21. 74 21. 02 21. 13 30. 62	16. 49 23. 31 26. 10 29. 46 25. 75 31. 19 41. 14 38. 74	27. 33 31. 40 35. 33 45. 51 56. 79 45. 36 36. 86 93. 25	35. 49 55. 85 65. 14 48. 71 52. 52 54. 56 103. 91 219. 21	

TABLE 1.—Annual mortality rates per 1,000 by policy year and age at issuance

If one follows down the rates given in boldface, or the intervening values, it will at once be evident that there is a factor of selection that is gradually dissipated—that as the policies become older, the mortality rates rise. Selective factors having to do with the type of person taking out insurance would exert a constant effect regardless of the number of years the policies have been in force. It appears a natural assumption that the selective factor which gradually fades out is that due to the medical examination (neglecting the small effect of withdrawals of healthy lives, already discussed). Furthermore, we can feel that this factor has ceased to be effective when the mortality rates for the same true age no longer rise as the policy years increase.

In the next table a comparison is made for the first, sixth, and eleventh year policies, because in that case it is possible to move the rates over one age group (in the case of the sixth year policies), and two age groups (in the case of the eleventh year policies). The table also gives the ratios to the first year policies. The indefinite age group, 65 years and over, is omitted.

	Attained age at specified policy year											
Policy year	15 to 19	20 to 24	25 to 29	30 to 34	35 to 39	40 to 44	45 to 49	50 to 54	55 to 59	60 to 64		
				RAT	E		<u> </u>	· · · · ·				
First Sixth Eleventh	1.96	1.76 2.68	1. 56 2. 37 2. 29	1. 76 2. 75 2. 92	2. 34 3. 42 3. 10	2.94 4.72 4.59	4. 43 7. 38 7. 96	6.06 11.15 9.21	10. 49 14. 89 18. 90	14. 66 23. 31 21. 13		
RA	TIO TO	O MOR	TALITY	T RATE	E IN FI	RST PO	DLICY	YEAR				
First Sixth Eleventh	100	100 152	100 152 147	100 156 166	100 146 132	100 161 156	100 167 180	100 184 152	100 142 180	100 159 144		

 TABLE 2.—Annual mortality rates per 1,000 for first, sixth, and eleventh policy years by true age, with ratios to first year

Even from this crude summary of the data two facts emerge: that there is an increase of more than 50 per cent in the mortality from the first to the sixth years, and that after this time there is no appreciable increase. This is shown very clearly from Figure 1 where these three curves have been plotted, allowing for the increase in age as the number of policy years increased.

It is suggested that somewhere between the first and the sixth year the effect of selection ceases to be operative so far as all causes of mortality are concerned. One may imagine that in the case of a specific disease, such as tuberculosis, this effect might persist for a much longer time, but information on that point is not at present

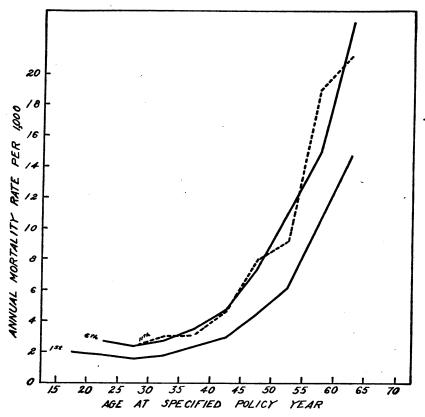


FIGURE 1.-Mortality rates for the first, sixth, and eleventh policy years, by true age

available. Furthermore, no question is raised now as to the factors of selection that differentiate the insured persons from the total population. We are concerned only with the effect of selection resulting from the medical examination itself. Later comparison will be made with the general population.¹⁵

¹⁵ For the reader who is doubtful about the results owing to the effect of time, it may be pointed out that the comparison between the first and sixth years is entirely free from such a factor. In both cases all deaths occurred during the period 1920-1926, since the first-year policies were based on that period and deaths for sixth-year policies would not have occurred before 1920. The same would be true in comparing the second and seventh year policies (based on 1921-1926), and the third and eighth year policies (based on 1922-1926), etc

In the next graph an attempt has been made to deal with the intermediate curves. It is evident that the age of persons in the second policy year will, on the average, be one year greater than of those in the first policy year. Therefore a lag of one year is allowed in plotting the second policy year curve, etc. In this case semilogarithmic paper has been employed to bring out the relative differences between the curves and especially to emphasize the differences in the earlier part of life, which would be more or less lost in a comparison on arithmetic paper.

One finds a definite excess in mortality for persons in the second year of their policies compared with those in the first year. The third

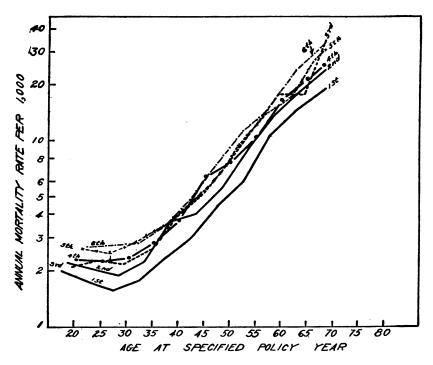


FIGURE 2.-Mortality rates for the first policy years, by true age

year also shows an additional increase in mortality, but of no great amount. The curve for the fourth year is, if anything, slightly under that for the third, but the fifth again shows a small increase. The mortality rate seems quite stabilized by the sixth year.

It was realized that a more intelligible expression of this relation would be secured if the number of years the policies had been in force could be used as the base line rather than the age. Although it would not be possible to disregard age, it did seem practicable to deal with three or four broad age groups showing curves for each. The difficulty, of course, lies in allowing for the fact that the data were according to age at issuance of the policies. An approximate but simple method was devised to allow for this factor. Suppose we take the age group 25 to 34. For the first policy year, the rate could be easily calculated from the number of persons exposed to risk and the number of deaths in the two 5-year age groups. For the second policy year, it seemed accurate enough to add to these figures onefifth of the persons and deaths in the age group 20 to 24, and to deduct from them one-fifth of the persons and deaths in the age group 30 to 34. For the third policy year, in the same way, two-fifths would be added from the 20 to 24 age group and two-fifths deducted from the 30 to 34 age group; and so on. This method was applied consistently for the succeeding broad age groups. Table 3 gives the results.

 TABLE 3.—Annual mortality rates by year of policy for three age groups, with ratio to first year 1

	At	tained age gr	oup
Policy year	25 to 34	35 to 44	45 to 54
RATIO TO FIRST YEAR			
First	100	100	10
Second	124	140	12
Third to fourth	143	148	14
Fifth to sixth Seventh to ninth	160	153	16
Tenth to twelfth	153 162	157 151	16
Tenta to twenta	102	101	16
DEATH RATE PER 1,000		•	
First	1.67	2.61	5.02
Second	2.07	3.65	6.3
Third to fourth	2.39	3.87	7.4
Fifth to sixth	2.68	3.99	8.14
Seventh to ninth Tenth to tweifth	2. 55 2. 71	4. 10 3. 93	8. 25 8. 36
EXPOSED TO RISK		<u>/</u>	
First	7, 836, 580	7, 480, 665	3, 501, 176
Second	5, 275, 376	5, 540, 249	2, 890, 597
Third to fourth	6, 258, 637	7, 136, 339	4, 082, 861
Fifth to sixth	5, 412, 464	6, 829, 047	4, 149, 447
Seventh to ninth	3, 384, 605	5, 056, 611	3, 565, 928
Fenth to twelfth	690, 287	1, 294, 748	1, 037, 609
NUMBER OF DEATHS?			
First	12 101	10 100	
econd	13, 101 10, 932	19,496	17, 579
Phird to fourth	14, 937	20, 247 27, 630	18, 236
Fifth to sixth	14, 512	27, 247	30, 496 33, 792
eventh to ninth	8,614	20,752	33, 792 29, 543
Fenth to twelfth	1,874	5,090	29, 543 8, 671
	_,	-,	0,0/1

First 4 insurance years, based on 1920-1926 data, the remaining years based on 1915-1926 data.
 Based on amounts insured and death claims paid, converted into persons on a unit of \$1,000.

In addition to the number exposed to risk and the number of deaths, the table gives the corresponding annual death rates, and also the ratio to the rate for the first year. A rather broad grouping of policy years has been followed in order to give regularity to the results.

The death rates from this table are reproduced in Figure 3. The higher age groups naturally have the higher mortality rates, but the point of particular interest is the rise in the curves during the earlier years of the policies. The significance of these curves is much better brought out by the ratios given in the preceding table, since these

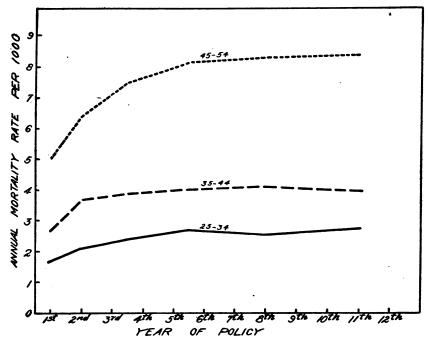


FIGURE 3.-Mortality by years of policy for three age groups

ratios bring the three age groups together at the beginning of policy life. The ratios are plotted in Figure 4. The results are quite consistent for the three age groups, and point unmistakably to the fact that the duration of selection due to the medical life insurance examination, taking all causes together, is hardly more than three or four years.

An estimate of the lessened mortality rates in insurance data as a result of the medical examination can be obtained from this graph. The ratio reaches a level of about 158. In other words, in the first year the mortality is 37 per cent less than it would be if we could imagine the insuring of people without examination, other factors remaining the same.¹⁶ In the second year the percentage is about 18; in the third, about 10; in the fourth, about 5. After that the difference is nominal.

In this comparison no reference has been made to the highest ages. It was felt that persons applying for life insurance much above the age of 50 formed a special group and were given a more thorough examination. The numbers were also small for these advanced ages and inconsistent results were noted on analysis. For these reasons, no consideration has been given to the higher ages other than that already given in the curves according to age. (Figs. 1 and 2.)

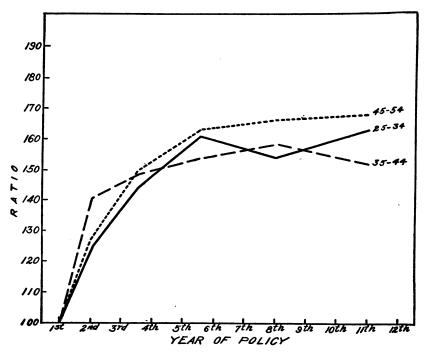


FIGURE 4.—Ratios of mortality in succeeding years of policy life to that in first year, for three age groups

No doubt the reader will wish to know how these mortality rates compare with those in the general population. In the next table a brief comparison of this sort is made. The rates for the sixth policy year ¹⁷ are used, since it has been shown that by this time the effect of selection due to the medical examination has been eliminated and because these rates could be used in a comparison according to age by adding five years to the age group as originally classified. This has been done. The general mortality is based on white males and

^{16 (158-100)} divided by 158.

¹⁷ The fifth, sixth, and seventh year policies were taken together to form these actuarial mortality rates in order to give smoothness to the ratios.

females in the registration States for the years 1920–1926. The rates for sixth year policies for the period 1915–1926 are based on deaths occurring since 1920, and therefore comparable chronologically. The table also gives the ratio of the rates in the Registration States to those for the actuarial data.

 TABLE 4.—Annual mortality rate by true age, for actuarial data (average for fifth, sixth, and seventh policy years) and registration States, 1920-1926

	20 to 24	25 to 29	30 to 34	35 to 39	40 to 44	45 to 49	50 to 54	55 to 59	60 to 64
			· J	RATE					
Actuarial Registration States	2.61 3.79	2. 38 4. 18	2. 73 4. 99	3. 40 5. 99	4. 53 7. 64	6. 67 9. 66	10. 12 13. 72	14.95 20.12	21. 04 29. 36
RATIO OF RATI	FOR F	EGISTI	RATION	STATES	з то те	IAT FO	R ACTU	ARIAL	DATA
	145	176	183	176	169	145	136	135	140

It appears that, even after the effect of selection due to the medical examination has been dissipated, the mortality rates for people insured at ordinary rates are much lower than those for the general population, the excess for the latter being in the vicinity of 50 per cent. This is to be accounted for by a difference in social or economic level, and in general indicates the effect of "permanent selection" due to the class of lives involved, referred to by Elderton. (See p. 47.) It is notable that the mortality for insured wage earners (industrial policies) is higher than that in the general population,¹⁸ the most marked difference occurring as here, in the younger adult ages, which are also the ages showing the greatest difference in mortality rates by economic status (England and Wales).¹⁹

It is possible, however, that part of the difference in the actuarial data is due to some inherent peculiarity in this material. For instance, it is noted in the Joint Report that the ratios of the basic data to the American Select Men table were higher in the older ages. In other words, there was an indication that the mortality rates based on the recent data were apparently too low in the early part of life. If this is true, then the ratios of the rates for the registration States to the basic data are too high at these ages.

A further point should be kept in mind, if any precise comparison of assurance and general mortality is desired. The data considered in this paper were based on amount of policy and death claims. Thus,

¹⁸ Mortality Statistics of Insured Wage-earners and Their Families. Experience of the Metropolitan Life Insurance Company industrial department, 1911 to 1916, in the United States and Canada. By Louis I. Dublin, with the collaboration of Edwin W. Kopf and George H. Van Buren. New York: Metropolitan Life Insurance Co., 1919.

¹⁹ Registrar General's Decennial Supplement, England and Wales, 1921, Part II, Occupational Mortality, Fertility, and Infant Mortality, London: 1927.

twice as much weight is given to deaths among persons carrying \$2,000 of insurance as to those among persons carrying the unit of \$1,000. To what extent the additional weight given to larger policies—i. e., those of persons in the higher economic or social levels or in the more responsible situations—affects the mortality rates in this group, is difficult to say.

Reference has already been made to the impossibility of making any satisfactory comparisons with respect to individual causes of mortality. The actuarial data as to cause of death were not obtained according to year of the policy and age at issuance. They were secured in two broad age groups. It is of interest, however, to make a brief comparison in these two broad groups. For the actuarial material, the groups are 15 to 39 and 40 and over (ages at issuance). The first group can be compared with some logic with the corresponding group of the general population, but the latter group can not, since there are relatively few persons exposed to risk in the higher ages, due primarily to the brief period covered by the actuarial data. It was felt that the most logical comparison would be with the age group 40 to 59 for the registration States. . The period of 1920-1926 is used for both sets of data to avoid the effect of the influenza epidemic. The rates and percentages according to cause are given in Table 5.

TABLE 5.-Death rates per 100,000 and percentage by cause in two age groups; registration States and actuarial data, 1980-1986

	of deaths oh cause	Actuarial	24
years -	Percentage of death due to each cause	Registra- tion States	84409494944444 8460949484444 000000044889180
15 to 59 years	Rate .	Actuarial	88888 40842 50842 50888 5095 5095 5095 5095 5095 5095 509
	R	Registra- tion States	66 86 86 86 86 86 86 86 86 86 86 86 86 8
	Percentage of death due to each cause	Actu- arial ¹	
40 to 59 years		Registra- tion States	0808547856644464 6 0808478666666666666666666666666666666666
40 to 5	Rate	Actu- arial ¹	699 699 881 881 881 881 881 881 881 881 881 8
	R	Registra- tion States	1, 101.8 101.8 101.8 101.8 10.0 25 25 25 25 25 25 25 25 25 25 25 25 25
	s of deaths ch cause	Actuarial	00148684040.04 001486940 00148868189400084
years	Percentage of death due to each cause	Registra- tion States	00 01 01 02 02 02 02 02 02 02 02 02 02 02 02 02
15 to 39 years	te	Actuarial	8. 59179949999890 8. 59175789571555888 9. 0400419944999599890
	Rate	Registra- tion States	622 622 623 623 623 623 623 623
-	Cause		All causes Tuberculosis of the lungs. Tuberculosis of the lungs. External causes of the rest. External sauses other than suicide External sauses other than suicide Preatmonia—all forms. Nephritis and Bright's disease Nephritis and Hright's disease Nephritis and typhitis. Appendicitis and typhitis. Appendicitis and typhitis. Typhoid and paratyphoid Cirrhosis of liver. All other causes.

140 and more.
 Organic disease of heart.
 A couldent due to occupation." These high rates are to be explained in part as being due to double accidental death provisions, zince the data are by policies, not by persons.
 Typhoid.

January 9, 1931

We have not felt that the rates themselves were of any great significance, owing to the differences which have been brought out in this paper. The percentages, however, put the comparison on a relative basis and make it much more interesting. Therefore the percentages have been plotted in Figure 5. The indications are more or less what would be expected, primarily a lower relative mortality

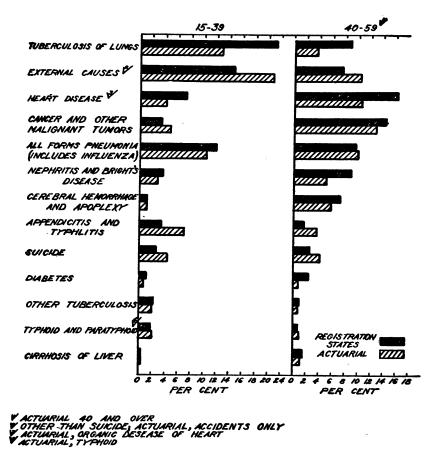


FIGURE 5.—Percentage of deaths, by cause, in two age groups, registration States and actuarial data, 1920-1926

from tuberculosis, heart disease, and some other degenerative diseases among the insured persons.

The particular bearing of these findings upon the major findings of this paper is the suggestion that the duration of selection due to the medical examination may not be identical for different causes of mortality, and may prove to be much greater for those causes given special stress in the medical examination. It is of interest that various acute causes of death—external causes, for instance—and cancer (where the medical examination could not be a selective factor of any moment) do not show an excess in the mortality rates for the registration States.

The following conclusions are suggested:

1. The insurance medical examination results in a lower mortality during the earliest policy years.

2. The duration of such selection would appear to last for three or four years for all causes, except possibly at the highest ages.

3. Most of the difference occurs in the first year or two of policy life.

4. The ratio of the mortality rate in successive policy years to that in the first year reaches a comparatively constant level at about 158.

5. For certain diseases, such as tuberculosis and heart disease, it is possible that the selective factor is of much longer duration.

The direct application of these conclusions in the field of public health should be discussed. Perhaps of most importance is the bearing which they may have upon the value of the medical examination, or the so-called periodic health examination, in the assessment of physical condition.

Life-insurance mortality data are increasingly important as a measure of the vitality of our people, because of the fact that the number exposed to risk is accurately known and the knowledge in regard to insured persons is much greater than that obtained in the course of securing Census data. This fact is now being realized and we may expect in the future that life-insurance records will be given more thorough analysis. Interpretation of such data is difficult without a knowledge of the effect of selection due to the medical examination. On the basis of the conclusions presented in this paper it is possible to show that such mortality rates can be used as a measure of health; in other words, that the medical examination in itself does not interfere seriously with the comparability of the material.

It is also important that it should be understood generally in connection with the analysis of life-insurance mortality data that if we exclude special mortality issued on industrial groups, these rates to-day are definitely below those of the general population.

28443°---31-----2

		66 and over		88244 88264 88278 88278 88278 88278 1188844 118884 118884 118884 118884 118884		82688888888888888888888888888888888888
ear 1		60 to 64		821 822 822 822 822 822 822 822 822 822		213 200 200 213 213 213 213 213 213 213 213 213 213
insurance year		• 56 to 59		864 897 897 897 897 897 897 897 897		550 11,12,0,0,0,0,0,0,0,0,0,0,0,0,0,0,0,0,0,
and by in		50 to 54		1, 276, 931 945, 328 6157, 228 8175, 728 8175, 728 8175, 728 8175, 728 8175, 728 8175, 728 817, 728 818, 950 818, 715 83, 715 83, 715 83, 715 83, 715 83, 715 848		7, 732 9,655 9,655 9,196 9,702 9,702 9,702 1,135
at entry		45 to 49		2, 224, 245 1, 217, 253 1, 217, 262, 567 1, 217, 263 730, 0667, 311 6677, 311 6677, 311 862, 253 963, 311 963, 358 963, 258 963, 258 943, 059 943, 059 944, 059 945, 059 945, 059 945, 059 945, 059 945, 059 945, 059 945, 059 945,		9 9 9 9 9 9 9 9 9 9 9 9 9 9 9 9 9 9 9
d by age	Age at entry	40 to 44		3, 315, 752 2, 405, 501 1, 708, 832 1, 147, 706 890, 065 892, 045 892, 589 151, 280 151, 280 150, 270 150, 270, 270 150, 270, 270, 270, 270, 270, 270, 270, 27		9,747 9,7947 9,7947 7,948 7,948 8,567 7,415 1,222 1,222 1,222 1,228 1,28
deaths classified	7	35 to 39) RI6K ¹	4, 164, 913 8, 012, 502 1, 579, 580 1, 579, 580 1, 579, 580 1, 150, 238 1, 150, 208 1, 15	DEATHS 2	11,054 11,054 8,410
of		30 to 34	EXPOSED TO	4, 224, 547 3, 002, 679 2, 194, 719 2, 194, 719 2, 194, 719 1, 587, 515 1, 587, 587, 515 1, 587, 587, 587, 587, 587, 587, 587, 587	NUMBER OF	9685 947 947 9487 9487 9487 9487 9473 9473 9473 9473 9473 9473 9473 947
and number	-	25 to 29	EX	8, 612, 033 2, 512, 033 1, 3512, 033 1, 345, 174 1, 345, 301 1, 345, 337 1, 366, 337 1, 376, 377 1, 376, 377 2, 37	NUN	5, 654 6, 757 9, 759 171 1, 503 171 1, 503 171 1, 503 1, 180 8, 148 1, 180 8, 148 8, 148 148 148 148 148 148 148 148 148 148
to risk an		20 to 24		2, 624, 578 1, 766, 178 1, 320, 320 1, 300, 320 1, 300, 300 1, 300, 300, 300 1, 300, 300, 300, 300, 300, 300, 300, 30		4,62,29,29,29,20,20,20,20,20,20,20,20,20,20,20,20,20,
exposed		15 to 19		1, 238, 003 913, 823 669, 710 669, 710 669, 710 861, 853 835, 662 835, 662 835, 662 137, 844 137, 844 147, 84414, 844 147, 84414, 844 147, 844 147, 84414		54 54 54 55 55 55 55 55 55 55 55 55 55 5
Number of persons		Insurance year		First- Becond Fourth- Fourth- Fitth- Fitth- Fitth- Beventh- Eighth- Bighth- Bighth- Bighth- Finth- Fielth- Fielth- Fuelth- Tweifth-		First. Becond Fourth Fourth Fifth Seventh Beventh Bighth Ninth Tenth Tweifth

APPENDIX TABLE

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¹ First 4 insurance years, based on 1930-1936 data; the remaining years, on 1915-1926 data. ³ Based on amounts insured and death claims paid, converted into persons on a unit of \$1,000.

PARASITES OF BATS

The United States Public Health Service has issued four bulletins as a key catalogue of the parasites of man, and a fifth bulletin of this series dealing with the parasites of monkeys in their relation to public health.

There has recently been published a sixth number of the series, National Institute of Health Bulletin No. 155, entitled "Key-Catalogue of the Parasites Reported for Chiroptera, with Their Possible Public Health Importance," by Ch. Wardell Stiles and Mabelle O. Nolan.

Some bats are used as food. The members of one family (the Desmodontidae), known as "vampires," attack man and livestock, sucking the blood and causing wounds which may become fly blown and form portals of bacterial infection. It is popularly believed that bedbugs are distributed by bats, but this view is due to confusing the common bedbug with closely allied bugs which live on bats and in bat haunts. Occasionally bats are kept as household pets. Some bats feed on mosquitoes, and thus potentially contribute to a reduction of these pests and the diseases they carry, although evidence is lacking which would justify our building so-called "batteries," or bat roosts, as a panacea against mosquitoes. Bats have numerous parasites, both external and internal, and of these no less than-11 species are reported as parasites both of bats and of man. Three additional parasites of man, including the causative agent of African sleeping sickness, are transmissible experimentally to bats.

The bulletin gives a classification of the parasitic protozoa, worms, arachnoids, and insects reported for these hosts, together with a classification of the bats themselves, and under each species of bost is given a list of the particular parasites reported for that particular animal.

The publication is exceedingly technical and is intended principally for distribution to public health officers, bacteriologists, and zoologists.

THE PATHOLOGY OF GENERALIZED VACCINIA IN RABBITS

An account of the pathologic histology of local and focal lesions of the skin, mucosæ, and viscera of rabbits produced by Armstrong's heat-selected vaccine virus is detailed in National Institute of Health Bulletin No. 156, recently issued by the United States Public Health Service. Also, the literature of the histology of variola and vaccinia is reviewed.

The visceral lesions are essentially coagulation necroses; those of the skin and mucosæ also show coagulation necrosis in addition to varied other proliferative, degenerative, hemorrhagic, and inflammatory changes. The distribution of such focal lesions is summarized in tabular form according to organs and by routes of inoculation and lapse of time after inoculation.

This bulletin is intended for restricted distribution to persons especially interested in the subject.

COURT DECISION RELATING TO PUBLIC HEALTH

Relief not granted in action against city brought because marketability of oysters was affected by pollution of tidal waters.—(Connecticut Supreme Court of Errors: Lovejov v. City of Norwalk, 152 A. 210: decided Nov. 7, 1930.) The plaintiff owned some oyster grounds under the navigable tidal waters of Long Island Sound in Norwalk Harbor. The title to such grounds had originated, pursuant to statute, in designations made by the oyster committee of the town of Norwalk, and had come to plaintiff by assignment from former owners. substantial part had been purchased during and since 1925. Sewage from Norwalk had for more than 50 years been discharged into the tidal waters of Long Island Sound. The plaintiff had resided in Norwalk and had been in the ovster business for more than 30 years. and in all that period was familiar with the sewerage system maintained by Norwalk and its effect upon the tidal waters. Plaintiff's grounds, involved in this action, had been used exclusively for growing and fattening oysters transplanted from other localities. The discharge of sewage by the defendant city of Norwalk did not interfere with the health and growth of ovsters upon the plaintiff's beds, but did introduce bacteria into the tidal waters. In 1925 the State health department adopted the policy of forbidding the marketing of oysters for human consumption from grounds within the State's jurisdiction unless an authorizing certificate was obtained. For the season of 1927-28, the plaintiff was refused certificates for the harvesting of oysters direct from his Norwalk beds, but was permitted to transplant oysters therefrom to other beds owned by him. In September, 1927, Norwalk employed competent engineers to make a survey of the sewage-disposal problem, and plans were submitted for the construction of a disposal plant. In November, 1928, the voters approved a proposal to issue bonds and to proceed with the building of the proposed plant. It was then the city's purpose, if legislative approval could be obtained, to construct such plant without delay. The plaintiff brought an action against the city of Norwalk for injury to his oyster grounds, resulting from the sewage discharged by the city. The judgment of the trial court in favor of the city was upheld by the supreme court of errors. After speaking of the city's action relative to the proposed disposal plant, the supreme court said:

* * * We think that this situation affords no indication that the defendant city had acted unreasonably, or negligently failed to take steps toward correction of the conditions of which the plaintiff complains. * * *

The supreme court stated the controlling conclusions reached by the trial court as follows:

* * * That the acts found were confined to tidal waters and did not constitute a public nuisance; that the plaintiff or his predecessors in title received their grants of oyster grounds subject to the public right of employing tidal waters for drainage purposes, and the exercise thereof by the defendant was not in derogation of any right of the plaintiff. * * *

After a discussion relative to the discharge of sewage into tidal waters and the designation, under the statutes, of places for the planting of oysters, the supreme court stated:

It follows, as stated subsequently by the United States Supreme Court in that case [Darling v. Newport News, 249 U. S. 540, 39 S. Ct. 371, 63 L. Ed. 759], that, as the trial court held, the recipients of the designations and the plaintiff, as their successor in interest, took the same subject to such rights as existed to discharge sewage into the waters of Norwalk River and harbor, and to the risk of the pollution of the water naturally resulting therefrom. * *

The court also held against the plaintiff's further claim of an unconstitutional taking of his property without compensation.

DEATHS DURING WEEK ENDED DECEMBER 20, 1930

Summary of information received by telegraph from industrial insurance companies for the week ended December 20, 1930, and corresponding week of 1929. (From the Weekly Health Index issued by the Bureau of the Census, Department of Commerce)

	Week ended Dec. 20, 1930	Corresponding week, 1929
Policies in force	74, 932, 777	75, 191, 352
Number of death claims	13, 608	14, 578
Death claims per 1,000 policies in force, annual rate_	9.5	10. 1

Deaths 1 from all causes in certain large cities of the United States during the week ended December 20, 1930, infant mortality, annual death rate, and comparison with corresponding week of 1929. (From the Weekly Health Index issued by the Bureau of the Census, Department of Commerce)

[The rates published in this summary are based upon mid-year population estimates derived from the 1930 census]

	Week ended Dec. 20, 1930					onding 1929	Death rate ² for first 51 weeks	
City	Total deaths	Death rate ³	Deaths under 1 year	Infant mor- tality rate ³	Death rate ³	Deaths under 1 year	1930	1929
Total (78 cities)	7, 907	11. 9	638	4 51	13. 1	763	11. 9	12.7
Akron Albany * Atlanta White Colored	47 30 85 50 35	9.6 12.2 16.5 	2 0 8 5 3	18 0 82 79 86	8.5 18.6 16.3 (9)	6 3 10 4 6	7.9 14.7 15.5 (9)	9.3 16.3 16.0

See footnotes at end of table.

January 9, 1981

Deaths ¹ from all causes in certain large cities of the United States during the week ended December 20, 1930, infant mortality, annual death rate, and comparison with corresponding week of 1929—Continued

	We	ek ended	Dec. 20	, 1930		ponding 1929	Death 1 first 51	week
City	Total deaths	Death rate ¹	Deaths under 1 year	Infant mor- tality rate ³	Death rate ³	Deaths under 1 year	1930	1929
Baltimore 4	188	12.2	13	45	16.8	17	14.0	14
White Colored Birmingham	150 38 61	(⁶) 12.3	12 1 7	53 16 67	(⁶) 13.2	11 6 8	(9) 13.6	(9) 15
White Colored	26		3	48		3		
Boston	35 194	(⁰) 12.9	4 21	98 61	(9) 14.1	0 27	() 14.0	(9) 14 11
Bridgeport	34	12.0		17	12.4	6	10.9	ii.
Buffalo	141	12.8	1 8 3 1	17 36 60	14.6	13	12.9	14.
Cambridge Camden	30	13.8	8	60	8.7	. 2	11.8	12
	22	9.8	1	18	15.6	4	13.5	14,
Canton Chicago ^s	20 644	9.8 9.9	1 32	27 28 53	6.5 11.8	1	9.8 10.4	11. 11.
Cincinnati	132	15.3	34 9	53	18.2	69 11	15.6	17.
Cleveland	100	9. 2	15	45 89	9.9	22	11. Ŏ	12
Columbus	76	13.7	9	89	17.3	7	15.4	14.
Dallas	58	11. 5	4		13. 3	9	11.4	11.
White Colored	- 46		3 1			7		
Dayton	12 47	(⁶) 12. 2	1		(⁰) 9.3	2	(⁰) 10. 8	() IL
Denver		17.0	14	15 1 53	14.3	4 9	15.0	14.
Des Moines	29	10.6	1	0	11.1	5	11.6	11.
Detroit	262	8.6	45	69	11.1	32	9.2	11
Ouluth	23	11.8	. 1	27	13.4	1	11.5	· 11.
l Paso	35	17.8	3		11.9	1	17.1	19.
foll River \$ 7	25	11. 2 13. 2	5	110	9.1 16.8	2	11.0	11.
lint	23 35 25 29 20	6.6	1	23 12	6.5	4	11.7 8.9	13. 10.
rie all River • 7	55	17.8	5	10	9.8	1	11.2	10.
White	47		5			î.		
Colored	8	(9) 8.6	5 5 0 1		(9) 7.8	0	(9)	(⁰) 10.
Frand Rapids	28	8.6	1	15	7.8	0	10.1	10.
White	58 43	10. 3	12 11		17.1	5 2 -	12. 2	12.
Colored	15	(6)	1		(0)	2 -	(6)	60
Colored ndianapolis	9 0	12.8	2	15	(⁶) 16. 2	3 8	(⁰) 14.3	(9 14.
White	75		2	17		6 .		
Colored	15	(6)	0	0	(⁶) 12.9	6 2 8 1	()	(6)
ansas City, Kans	69	ÌÍ. 4	10	87	12.9	8	11.3	- 12
White	39 34	16.6	75	163 138	10.3	1	11.8	12
Colored	5	(6)	2	303	(6)	1-		(6)
White Colored ansas City, Mo	96	(⁶) 12.7	23	25	(9) 14.4	9	13.4	(9) 14.
noxville	18	8.8	0	0	9.5	1	13.4	13.
White Colored	13		0	0 -		1 -		
os Angeles	5 304	(9) 12.7	0 34	0	(⁶) 13. 5	0	9 11.1	(⁶) 11.
ouisville	96	16.3	4	103 34	13. 0	25 9	13.5	11.
White Colored	68		i	10		9	10. 0	30.
Colored	28	(⁰) 7.3	3	199	(⁶) 18.5	ŏ	(6)	()
owell 7	14	7.3	2	53	ìś. 5	3	(⁰) 13. 2	(⁰) 14.
emphis	23 90	11.7	1	28	10.2	1	10.5	11.
wen	90 41	18.6	7	82 0	16.3	10	16.9	18.
	49	()	7	235	(0)	5	6	<i>(</i> 1)
liwaukee	119	10.9	25	110	() 11.1	24	(•) 9.8	(⁰) 10.
inneapolis	95	10.7	11	72	12.5	24 4	10.8	10.1
eshville	46	16.3	2	31	19. 2	5 1	17. 2	18.
White	24 -		2	42 -		2		
ew Bedford 7	22 28	(⁰) 12.9	03		(⁶) 9.7	3	(⁰) 11. 0	() 11.9
w nite Colored ew Bedford 7 ew Haven w Orleans	36	12.9	3	77 15	9.7	3 2	11.0 12.5	11.9
CW Olicalis	149	17.0	14	78	18.6	16	12.5	13.4
white	99		8	68 _		9	****	A4. 0
Colored	50	()	ā l	97	()	7	(9)	()

See footnotes at end of table.

Deaths from all causes in certain large cities of the United States	during the week
ended December 20, 1930, infant mortality, annual death rate,	and comparise
with corresponding week of 1929—Continued	•

•	Wee	k ended	Dec. 20,	193 0		ponding 1929	Death rate ² for first 51 weeks	
City	Total deaths	Death rate ³	Deaths under 1 year	Infant mor- tality rate ³	Death rate ³	Deaths under 1 year	1930	1929
New York	1, 449	10.8	112	47	12. 2	142	10. 7	11. 3
Bronx Borough	198	8.1	8	23	9.3	18	7.8	8. 2
Brooklyn Borough	488	9.7	56	59	10.6	47	9.7	10. 2
Manhattan Borough	567	16.0	37	47	17.1	60	16.0	16.3
Queens Borough Richmond Borough	152	7.2	8	32	9.6 17.9	14	7.0 13.9	7.7
Richmond Borougn	44 93	14.5 10.9	36	58 31	17.9	39	13.9	15.9 12.7
Newark, N. J	64	10.9	3	31	13.3	4	11.9	12.7
Oakland Oklahoma City	49	13.8	2	36	9.2	5	11.0	10.9
Omaha	44	10.7	3	. 36	13.5	i	13.5	13.5
Paterson	45	17.0	5	87	14.0	$\hat{2}$	12.1	13.4
Philadelphia	471	12.5	35	52	13.7	46	12.5	13.1
Pittsburgh	173	13.4	22	78	13.0	14	13.8	14.8
Portland, Oreg	75	13.0	3	37	9.9	5	12.2	12.7
Providence	76	15.8	10	93	15.0	7	12.9	14.5
Richmond	49	13. 9	5	73	17.2	5	14.8	16. 2
White	32		2	44		3		
Colored	17 70	(⁶) 11. 2	3 5	128 44	(⁶) 12.9	2 3	(⁶) 11.6	(⁶) 12.3
Rochester	199	11. 2	13	44	14.1	15	14.0	14.6
St. Louis St. Paul	199 52	10.0	13	20	10.5	4	10.1	10.6
Salt Lake City 4	38	14.1	$\tilde{2}$	32	10.9	5	12.6	12.9
San Antonio	64	13.0	2	(າ)ີ	18.3	15	14.3	14.7
San Diego	48	16.8	4	` 84	14.6	3	14.5	15.1
San Francisco	191	15.8	4	27	14.0	10	13.3	13. 1
Schenectady	15	8.2	1	31	14.2	3	11.0	12.2
Seattle	101	14.5	6	61	8.3	5	11.0	11.2
Somerville	25	12.6	4	126	12.7	3	9.7	9.3 12.9
Spokane	26 43	11.7 14.9	1 5	26	16.8 14.1	1	12.4 12.1	12.9
Springfield, Mass	43 46	14.9	57	86 86	14.1	4 6	11.7	12.7
Syracuse Tacoma	20	9.7	il	27	7.9	1	12.4	11.7
Toledo	73	13.0	6	55	12.7	2	12.7	13.7
Trenton	31	13.2	4	77	21.7	5	16.6	17.1
Utica	39	19.8	$\tilde{2}$	56	17.9	ī	14.5	15.5
Washington, D. C.	132	14.1	4	23	15.6	6	15.2	15.4
White	81		2	17		3		
Colored	51	(6)	2	36	(6)	3	(6)	(6)
Waterbury	13	6.7	1	24	6.8	2	9.3	9.3
Wilmington, Del. ⁷	32	15.9	2	48	12.4	4	14.6	13.8
Worcester	56	14.8	1	14	15.8	6	12.6	12.6
Yonkers	31	11.9	2	48 57	11.8 15.6	67	8.1 10.4	9.4 12.4
Youngstown	35	. 10.7	*	57	10.0		10.4	14.9

¹ Deaths of nonresidents are included. Stillbirths are excluded. ² These rates represent annual rates per 1,000 population, as estimated for 1930 and 1929 by the arithmetical method. ³ Deaths under 1 year of age per 1,000 live births. Cities left blank are not in the registration area for births.

births.
Data for 73 cities.
Deaths for week ended Friday.
For the cities for which deaths are shown by color the colored population in 1920 constituted the following percentages of the total population: Atlanta, 31; Baltimore, 15; Birmingham, 39; Dallas, 15; Fort Worth, 14; Houston, 25; Indianapolis, 11; Kansas City, Kans., 14; Knoville, 15; Louisville, 17; Memphis, 38; Nashville, 30; New Orleans, 26; Richmond, 32; and Washington, D. C., 25.
Population Apr. 1, 1930; decreased 1920 to 1930; no estimate made.

PREVALENCE OF DISEASE

No health department, State or local, can effectively prevent or control disease without knowledge of when, where, and under what conditions cases are occurring

UNITED STATES

CURRENT WEEKLY STATE REPORTS

These reports are preliminary, and the figures are subject to change when later returns are received by the State health officers

Reports for Weeks Ended December 27, 1930, and December 28, 1929

Cases of certain communicable diseases reported by telegraph by State health officers for weeks ended December 27, 1930, and December 28, 1929

	Diph	theria	Infl	uenza	Me	asles	Meningococcus meningitis	
Division and State	Week ended Dec. 27, 1930	Week ended Dec. 28, 1929						
New England States: Maine New Hampshire Vermont	2	42	3 12	6	18 3	4 14 41	0 0 0	1 0 0
Massachusetts Rhode Island Connecticut	62	103 7 23	6	6 	273 1 118	171 1 19	3 0 0	0 2 0 3
Middle Atlantic States: New York New Jersey Pennsylvania		157 132	¹ 25 13	¹ 28 11	151 120 406	191 56	6 2 0	13 4
East North Central States: Ohio Indiana Illinois	69 35 146	89 21 212	7 1 12	55 24	24 138 208	351 22 303	2 4 7	12 28 9
Michigan Wisconsin West North Central States: Minnesota	16 14 10	114 11 24	2 41 2	3 22 1	49 191 24	163 485 119	1 5 1	28 5
Iowa. Missouri. North Dakota South Dakota	12 25 3 13	10 34 8 1	3 1	11	656 5	162 11 98 3	24 3 0 0	1 5 6 1
Nebraska Kansas South Atlantic States:	4 24	15 23	2 1	8	2 10	174 116	3 1	3 1
Delaware Maryland ' District of Columbia West Virginia Neth Garalia	3 39 10 11	1 23 6 17	12 2 16	1 19 17	18 12 31	1 13 90	0 0 1 0	02000
North Carolina South Carolina Georgia Florida	23 12 23 6	67 27 34 9	9 588 42 1	12 903 148 5	50 32 25	6 24 7	1 1 0 0	3 5 6 0
¹ New York City o	only.		3	Week en	ded Frie	lays.		

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Cases of certain communicable diseases reported by telegraph by State health officers for weeks ended December 27, 1930, and December 28, 1929—Continued

	Diph	theria	Inft	lenza	Me	asles	Menin men	gococcus ngitis
Division and State	Week ended Dec. 27, 1930	Week ended Dec. 28, 1929						
East South Central States: Kentucky		20			19	10	1	0
Tennessee Alabama Mississippi	20 39 22	14 19 18	46 36	109 62	24 122	16	1 0 4	6 0 1
West South Central States: Arkaneas Louisiana Oklahoma ⁸	3 12 24	8 36 49	25 10 41	69 24 113	1	7 10 19	1 0 1	1 6 8
Texas. Mountain States: Montana Idaho	33	112 	222 	40 	81 1 10	51 38 22	1 0 1	2
Wyoming Colorado New Mexico Arizona	1 8 6 3	5 6 28 11			43 121 28	3 14 3 1	1 2 0 1	5 1 4 0 3
Utah ³ Pacific States: Washington		14	15	4	1 6	66 15	0	3 2 5
Oregon California	5 46	67 67	27 57	29 39	52 109	11 203	0 8	14
	Polion	nyelitis	Scarle	t fever	Smallpox		Typhoid fever	
Division and State	Week ended Dec. 27, 1930	Week ended Dec. 28, 1929						
New England States: Maine	3	0	19	38	0	0	2	0
New Hampshire Vermont Massachusetts Rhode Island Connecticut Vident Martin Street	0 6 0 0	0 0 1 0 1	2 4 222 31 55	13 13 293 23 63	0 0 0 0 0	0 6 0 0 0	2 0 8 0 1	0 1 2 0 0
Middle Atlantic States: New York New Jersey Pennsylvania East North Central States:	1 0 1	1 0	436 142 370	312 161	7 0 0	3 0	8 5 19	8 3
Ohio Indiana Illinois Michigan	2 0 4 0	3 1 3	381 172 389 134	289 137 455 251	45 53 53 12	136 138 90 57	18 3 17 15	15 0 3 1 0
Wisconsin West North Central States: Minnesota Iowa	1 3 2 0	0 0 0	122 61 104	92 98 43 57	9 4 17 3	29 8 79 50	2 2 0 4	8 2 1
Missouri North Dakota South Dakota Nebraska Kansas	1 1 2 2	1 0 1 0	128 9 6 37 46	44 11 54 127	2 20 22 47	13 14 32 24	0 1 1 1	0 0 0 1
South Atlantic States: Delaware Maryland ³ District of Columbia	0 0 1	0 0 0	12 75 23	5 50 25	0 0 0	0	0	0
West Virginia North Carolina South Carolina Georgia	1 1 1 0	0 3 2 0	62 22 23 28	54 60 32 42	13 0 2 0	14 11 3 0	2 8 1 5 6	7 1 8 5 1 7
Florida East South Central States: Kentucky Tennessee	0 0 0	0 0 1	8 43 22	3 52 20	0 10 0	2 31 5	1 5 5	1 0 8
Alabama Mississippi	2 0	0	52 19	29 17	6 5	22	12 5	4 7

Week ended Friday.
 Figures for 1930 are exclusive of Oklahoma City and Tulsa.

Cases of certain communicable disea for weeks ended December 27, 2	uses reported by telegraph	, by State health officers
for weeks ended December 27, 2	1930, and December 28,	1929—Continued

	Poliomyelitis		Scarlet fever		Smallpox		Typhoid fever	
Division and State	Week ended Dec. 27, 1930	Week ended Dec. 28, 1929						
West South Central States:								
Arkansas	0	0	9	19	3	16	6	
Louisiana	i	ŏ	1 11	20	16	ň	ğ	5
Oklahoma ³	Ô	ı i	55	48	19	137	7	12
Texas	ŏ	ō	19	61	12	54	3	12
Mountain States:	Ţ						-	-
Montana	0	0	24	28	11	10	0	0
Idaho	i i	l ŏ	1	4	Ō	7	Ó	Ó
Wyoming	1	Ó	6	6	Ó	3	0	1
Colorado	0	1	55	28	1	23	1	0
New Mexico	0	0	8	22	0	4	4	0
Arizona	0	0	3	9	0	16	0	1
Utah ²	0	0	5	14	0	0	0	0
Pacific States:								
Washington	1	1	46	50	13	123	1	6
Oregon	2	0	8	38	9	14	0	1
California	12	1	76	208	21	60	7	4

Week ended Friday.
Figures for 1930 are exclusive of Oklahoma City and Tulsa.

SUMMARY OF MONTHLY REPORTS FROM STATES

The following summary of cases reported monthly by States is published weekly and covers only those States from which reports are received during the current week.

State	Menin- gococ- cus menin- gitis	Diph- theria	Influ- enza	Ma- laria	Mea- sles	Pel- lagra	Polio- mye- litis	Scarlet fever	Small- pox	Ty- phoid fever
November, 1930										
Alabama California Idaho Illinois Maryland Minnesota Minnesota Missouri Mow Hampshire Now Hampshire New Hampshire North Carolina Oklahoma ¹ Oklahoma ¹ Oregon Rhode Island South Dakota Texas Virginia Washington Wisconsin	13 15 5 11 23 9 9 2 6 21 1 1 	467 316 722 111 728 83 83 351 111 11 222 510 276 18 44 43 33 328 370 89 89	200 132 26 53 63 4 28 111 28 186 51 111 36 36 148 1,28 23 104	248 2 72 1 10 10 125 537 28	178 529 248 300 373 13 27 55 1, 290 194 5 5 5 5 483 66 645	30 5 	19 146 8 2 48 6 6 60 35 3 4 4 21 30 8 3 21 30 8 3 24	409 408 136 48 1, 336 254 244 577 224 577 224 86 65 5 65 5 47 164 438 180 384	3 86 43 10 96 8 0 42 63 24 6 27 65 65 65 65 65 89 29	84 58 25 70 85 85 85 85 85 85 85 85 85 85 85 85 85

¹ Exclusive of Oklahoma City and Tulsa.

November, 1930

Actinomycosis:	Cases
California	1
Anthrax:	
California	1
Illinois	1
Chicken pox:	-
Alabama	176
California	902
Colorado	282
Idaho	33
Illinois	
Louisiana	43
Maryland	294
Minnesota	671
Missouri	405
Montana	280
North Carolina	230 580
Oklahoma 1	47
Oregon	223
Rhode Island	76
South Dakota	122
Virginia	492
Washington	305
Wisconsin	1, 894
Conjunctivitis:	
Montana	4
Oklahoma ¹	2
Dengue:	_
Alabama	2
Diarrhea:	
Maryland	20
Diarrhea and dysentery:	
Virginia	118
Dysentery:	
California (amebic)	2
California (bacillary)	13
Illinois	37
Illinois (amebic)	4
Illinois (bacillary)	2
Louisiana	4
Maryland	21
Minnesota	1
Oklahoma 1	11
Washington	4
German measles:	
California	36
Colorado	7
Illinois	13
Maryland	4
Montana	3
North Carolina	23
Rhode Island	2
Washington	23
Granuloma, coccidioidal:	
California	1
Hookworm disease:	
Louisiana	81
Impetigo contagiosa:	
Colorado	4
Maryland	22
Oregon	18
Lead poisoning:	
Illinois	2
	- 1
Exclusive of Oklahoma City and Tulsa.	

Leprosy:	Cases
California	. 1
Idaho	. 1
Louisiana	. 1
Washington	1
Lethargic encephalitis:	
Alabama	2
California	3
Colorado	2
Illinois	2
Louisiana	3
Texas	1
Wisconsin	i
Mumps:	•
•	31
Alabama California	617
	134
Colorado	
Idaho	5
Illincis	757
Louisiana	1
Maryland	31
Missouri	59
Montana	61
Oklahoma 1	4
Oregon	106
Rhode Island	9
South Dakota	17
Washington	125
Wisconsin	477
Ophthalmia neonatorum:	
Celifornia	3
Illinois	48
Maryland	2
Missouri	10
	2
Oklahoma 1	4
Paratyphoid fever:	
California	3
Illinois	1
Minnesota	8
North Carolina	1
Texas	4
Washington	7
Puerperal septicemia:	
Ulinois	4
Washington	1
Rabies in animals:	-
California	69
Illinois	1
Louisiana	6
Maryland	2
Missouri	6
Oregon	1
Rhode Island	2
Scabies:	-
Maryland	4
Oregon	9
Septic sore throat:	11
Illinois	7
Maryland	-
Missouri	14
Montana	1
North Carolina	14
Oklahoma 1	32
Oregon	3

January 9, 1931

Septic sore throat—Continued.	Cases	Undulant fever:	Cases
Rhode Island	. 1	California	. 18
South Dakota	. 4	Illinois	. 5
Tetanus:		Louisiana	
California	. 8	Maryland	. 6
Illinois	. 2	Minnesota	. 2
Louisiana.	. 5	Missouri	. 10
Oklahoma 1	. 1	Oregon	. 2
South Dakota	. 1	Washington	. 7
Trachoma:		Vincent's angina:	
California	122	Colorado	. 3
Illinois	. 1	Illinois	. 3
Maryland	1	Maryland	9
Minnesota	1	Washington	
Missouri	61	Whooping cough:	ŧ
Montana	1	Alabama	79
Oklahoma.	3	California	426
South Dakota	1	Colorado	81
Trichinosis:		Idaho	38
California	2	Illinois	491
Illinois	3	Louisiana	24
Tularaemia:		Maryland	102
Illinois	9	Minnesota	99
Louisiana	1	Missouri	54
Maryland	1	Montana	111
Minnesota	3	North Carolina	370
Missouri	3	Oklahoma	21
South Dakota	1	Oregon	66
Virginia	6	Rhode Island	55
Wisconsin	2	South Dakota	29
Typhus fever:		Virginia	254
Alabama	6	Washington	139
Maryland	2	Wisconsin	57 3
¹ Exclusive of Oklahoma City and Tulsa			

¹ Exclusive of Oklahoma City and Tulsa.

Cases of Certain Communicable Diseases Reported for the Month of August, 1930, by State Health Officers

State	Chick- en pox	Diph- theria	Measles	Mumps	Scarlet fever	Sman- pox	Tuber- culosis	Ty- phoid and paraty- phoid fever	Whoop- ing cough
Maine. New Hampshire	12	15 6	17	74	44 6	0	77	20	142
Vermont	19	3	18	16	ň	ŏ	14	2	63
Vermont Massachusetts	70	183	227	144	192	ŏ	374	63	551
Rhode Island	5	11	8		13	ŏ	56	6	37
Connecticut	17	24	32	23	31	ŏ	109	ő	129
					•-	, i	100	v	
New York	184	237	603	243	234	1	1,656	116	1, 418
New Jersey	30	145	165	43	76	ō	463	67	272
Pennsylvania	159	183	495	168	232	ŏ	497	206	829
-									
Ohio	157	101	97	61	234	42	578	201	401
Indiana	8	45	25	11	73	116	162	68	109
Illinois	100	258	75	172	235	79	855	173	652
Michigan	79	121	194	72	170	57	420	70	706
Wisconsin	136	51	255	140	108	23	185	32	930
Minnesota	33	48	19		59	11	158	27	101
Iowa	7	12	4	9	27	43	36	8	38
Missouri	25	78	66	29	78	54	217	148	79
North Dakota	4	8	15	34	29	30	14	25	64
South Dakota	7	30	7	••	10	29	iil	19	9
Nebraska	25	19	28	6	17	36	13	20	58
Kansas.	15	40	45	25	17 71	27	75	73	112
					•• •			10]	

¹ Pulmonary.

State	Chick- en pox	Diph- theria	Measles	Mumps	Scarlet fever	Small- pox	Tuber- culosis	Ty- phoid and paraty- phoid fever	Whoop- ing cough
Delaware. Maryland. District of Columbia Virginia. West Virginia. North Carolina. South Carolina. Georgia. Florida	13 7 90 15 32 46 4 5	8 41 13 98 44 244 86 47 11	13 18 26 208 43 16 13 45 2	8 13 22 24 22	6 34 14 135 58 136 26 61 11	0 0 4 19 11 2 3 0	13 233 80 190 46 135 91 38	29 240 21 312 245 244 305 248 21	4 123 22 285 145 437 233 75 12
Kentucky ³ Tannessee, Alabama, Mississippi Arkansas Louisiana, Oklahoma ³ Texas.	10 6 142 18 1 2	43 51 61 5 41 25 99	27 85 96 20 47	14 24 156 4 6 1	76 80 26 14 23 29 74	9 2 9 12 0 43	¹ 233 370 263 ¹ 13 ¹ 178 52	529 186 161 126 154 248 138	115 143 379 44 34 32
Montana. Idaho	11 11 2 	3 7 28 32 11	14 13 2 65 14 37	26 2 61 13 9	32 8 15 34 6 7	11 8 0 4 12 1 1 0	33 6 114 59 118 18	13 8 3 47 28 27 0	87 60 15 199 12 29 4
Washington Oregon California	43 24 130	29 19 147	81 82 305	75 55 333	44 24 136	42 20 46	115 37 760	21 29 85	188 136 377

Cases of Certain Communicable Diseases Reported for the Month of August, 1930, by State Health Officers—Continued

¹ Pulmonary

² Reports received weekly.

³ Exclusive of Oklahoma City and Tulsa.

Case Rates per 1,000 Population (Annual Basis) for the Month of August, 1930, Based on Provisional Populations

State	Chick- en pox	Diph- theria	Measles	Mumps	Scarlet fever	Small- pox	Tuber- culosis	Ty- phoid and para- typhoid fever	Whoop- ing cough
Maine New Hampshire Vermont Massachusetts Rhode Island Connecticut	. 19	0. 22 . 15 . 10 . 51 . 19 . 18	0.25 .59 63 .14 .23	1.09 .52 .40 .17	0.65 .15 .36 .53 .22 .23	0.00 .00 .00 .00 .00 .00	1. 13 ¹ . 13 1. 03 . 96 . 80	0.29 .10 .07 .18 .10 .04	2,09 2,06 1,52 .63 .94
New York New Jersey Pennsylvania	.09	. 22 . 42 . 22	. 56 . 48 . 60	. 23 . 13 . 20	. 22 . 22 . 28	. 00 . 00 . 00	1.54 1.35 .61	. 11 . 19 . 25	1. 32 . 79 1. 01
Ohio Indiana Illinois Michigan Wisconsin	. 28 . 03 . 16 . 19 . 55	. 18 . 16 . 40 . 29 . 20	. 17 . 09 . 12 . 47 1. 02	. 11 . 04 . 27 . 17 . 56	. 41 . 27 . 36 . 41 . 43	. 07 . 42 . 12 . 14 . 09	1. 02 . 59 1. 32 1. 02 . 74	. 36 . 25 . 27 . 17 . 13	. 71 . 40 1. 01 1. 71 3. 73
Minnesota Iowa Missouri North Dakota South Dakota Nebraska Kansas	. 15 . 03 . 08 . 07 . 12 . 21 . 09	. 22 . 06 . 25 . 14 . 51 . 16 . 25	. 09 . 02 . 21 . 26 . 12 . 24 . 28	.04 .09 .59 .05 .16	. 27 . 13 . 25 . 50 . 17 . 14 . 44	.05 .21 .18 .52 .49 .31 .17	. 72 . 17 . 70 . 24 . 19 . 11 . 47	. 12 . 04 . 48 . 43 . 32 . 17 . 46	. 46 . 18 . 26 1. 10 . 15 . 49 . 70
Delaware Maryland District of Columbia Virginia	.09 .17 .44	.39 .30 .31 .48	. 64 . 13 . 63 1. 01	. 15 . 09	.30 .25 .34 .66	.00 .00 .00 .02	.64 1.68 1.93 .92	1. 43 1. 73 . 51 1. 52	. 29 . 80 . 53 1, 39

¹ Pulmonary.

State	Chick- en pox	Diph- theria	Measles	Mumps	Scarlet fever	Small- pox	Tuber- culosis	Typhoid and para- typhoid fever	Whoop- ing cough
West Virginia	.12	. 30 . 90 . 58 . 19 . 09	. 29 . 06 . 09 . 18 . 02	. 15 . 10 . 18	. 39 . 50 . 18 . 25 . 09	. 13 . 04 . 01 . 01 . 00	. 31 . 92 . 37 . 30	1.66 .90 2.07 1.01 .17	.98 1.62 1.58 .30 .10
Kentucky ³ Tennessee Alabama Mississippi Arkansas Louisiana Oklahoma ³ Teras	.05 .03	. 19 . 23 . 36 . 03 . 23 . 14 . 20	.12 .38 .56 .00 .11 .27	.06 .11 .91 .03 .03 .01	. 34 . 36 . 15 . 09 . 13 . 16 . 15	.04 .01 .05 .08 .00 .24	¹ 1.05 1.64 1.54 ¹ 08 ¹ 1.00 .30	2.38 .83 .94 .80 .86 1.41 .28	. 52 . 63 2. 22 . 28 . 19 . 18
Montana Idaho Wyoning Colorado New Mexico Arizona	. 29 . 10	. 20 .07 .18 .32 .88 .30	. 31 . 34 . 10 . 74 . 38 . 99	. 57 .05 .10 .69 .36 .24	. 13 . 70 . 21 . 78 . 39 . 16 . 19	. 24 . 21 . 00 . 05 . 33 . 03	.72 .16 1.29 1.62 3.17	.29 .29 .21 .16 .53 .77 .73	1.91 1.58 .78 2.26 .33 .78
Utah ³	.52 .32 .30 .27	. 22 . 23 . 30	.61 1.01 .63	. 56 . 68 . 69	.33 .30 .28	.00 .32 .25 .09	¹ . 64 . 86 . 46 1. 56	.00 .16 .36 .17	.52 1.41 1.67 .78

Case Rates per 1,000 Population (Annual Basis) for the Month of August, 1930, Based on Provisional Populations—Continued

¹ Pulmonary

¹ Exclusive of Oklahoma City and Tulsa.

GENERAL CURRENT SUMMARY AND WEEKLY REPORTS FROM CITIES

The 94 cities reporting cases used in the following table are situated in all parts of the country and have an estimated aggregate population of more than 31,690,-000. The estimated population of the 87 cities reporting deaths is more than 30,100,000. The estimated expectancy is based on the experience of the last nine years, excluding epidemics.

	1930	1929	Estimated expectancy
Cases reported			
Diphtheria:	1 100		
46 States 94 cities	1,483	1, 756	
Measles:	584	767	1,066
	9.470	0.040	
	3, 470	3, 846 663	
94 cities Meningococcus meningitis:	1, 181	003	
46 States	73	165	
94 cities	39	100	
Poliomyelitis:	38	100	
46 States	91	24	
Scarlet fever:	51	24	
46 States	3, 908	3,887	
94 cities	1,435	1,500	1,205
Smallpox:	1, 100	1,000	1,200
46 States	634	1,067	
94 cities	57	142	37
Typhoid fever:			
46 States	314	207	
94 cities	53	32	30
Deaths reported			
Influenza and pneumonia:			
87 cities	723	1,009	
Smallpox:		_,	
87 cities	0	0	
	-	•	

Weeks ended December 20, 1930, and December 21, 1929

³ Reports received weekly.

The "estimated expectancy" given for diphtheria, poliomyelitis, scarlet fever, smallpox, and typhoid fever is the result of an attempt to ascertain from previous occurrence the number of cases of the disease under consideration that may be expected to occur during a certain week in the absence of epidemics. It is based on reports to the Public Health Service during the past nine years. It is in most instances the median number of cases reported in the corresponding weeks of the preceding years. When the reports include several epidemics, or when for other reasons the median is unsatisfactory, the epidemic periods are excluded, and the estimated expectancy is the mean number of cases reported for the week during nonepidemic years.

If the reports have not been received for the full nine years, data are used for as many years as possible, but no year earlier than 1921 is included. In obtaining the estimated expectancy, the figures are smoothed when necessary to avoid abrupt deviation from the usual trend. For some of the diseases given in the table the available data were not sufficient to make it practicable to compute the estimated expectancy.

	ł	Diph	theria	Influ	ienza			Pneu-	
Division, State, and city	Chicken pox, cases reported			Cases Deaths reported reported		Measles, cases re- reported	Mumps, cases re- reported	monia, deaths reported	
NEW ENGLAND									
Maine: Portland	2	1	0	1	1	0	0	2	
New Hampshire: Concord	0	0	0		0	0	0	2	
Vermont:	-		, , , , , , , , , , , , , , , , , , ,		-		-		
Barre Burlington	0	0	0		0	0	0	0	
Massachusetts:	-	-	•			-	Ť	•	
Boston Fall River	56 22	41 4	33 6	1	0	59 0	13 5	25 3	
Springfield	14	5	0		0	3	5	2	
Worcester Rhode Island:	38	6	6		-	1	-		
Pawtucket	1 9	1 10	. 4 5		0	0	0	2	
Connecticut:	÷		-		-	-		-	
Bridgeport Hartford	3	7	0		0	1	. 3	2	
New Haven	8	2	0		0	8	8	1	
MIDDLE ATLANTIC									
New York:									
Buffalo	35 180	18 195	16 71	23	2 5	6 98	38 35	16 154	
New York Rochester	180	195	7	20 	Ó	1	2	5	
Syracuse New Jersey:	20	4	0		0	2	0	3	
Camden	4	6	1		0	53	7	4	
Newark Trenton	70 7	23 3	7	3	0	0	62	17 2	
Pennsylvania:			-		-	-	· .]	-	
Philadelphia Pittsburgh	110 73	71 23	16 18	5	2	16 10	28 17	49 28	
Reading	12	3	Õ		Ō	5	23	1	
EAST NORTH CENTRAL									
Ohio: Cincinnati	6	14	1		4	10	15	10	
Cleveland	170	44	14	7	Ō	2	72	12	
Columbus Toledo	3 110	8 10	2 10	1	1	2	7	2	
Indiana:					1	1	0	3	
Fort Wayne Indianapolis	2 19	5 10	1 - 3 -		0	0	Ó	12	
South Bend	4	1	Ō.		0	. 0	0	2 2	
Terre Haute	3	1	-			-	-	-	
Chicago	105	132	122	4	7	11	60 0	30 1	
Springfield Michigan:	3	-	_			-	-	-	
Detroit	97 26	67 4	33 0	1	3	8	18 1	20 1	
Grand Rapids	<i>6</i> 2	ī							

		Diph	theria	Influ	uenza			Pneu-	
Division, State, and city	Chicken pox, cases reported	Cases, estimated expect- ancy	Cases reported	Cases reported	'Deaths reported	Measles, cases re- reported	Mumps, cases re- reported	monia, deaths reported	
EAST NORTH CENTRAL—Continued									
Wisconsin: Kenosha Madison Milwaukee Racine Superior	32 48 132 55 5	1 2 20 3 0	0 2 9 0 0		0 	0 1 5 0 1	7 10 68 0 0	0 9 0 0	
WEST NORTH CENTRAL									
Minnesota: Duluth Minneapolis St. Paul Iowa:	7 56 32	1 20 12	0 11 0	2	0 0 2	0 1 0	1 10 0	1 4 8	
Davenport Des Moines Sioux City Waterloo Missouri:	2 1 7 10	0 3 1 0	0 0 0 0			0 0 1 0	0 0 4 0		
Kansas City St. Joseph St. Louis North Dakota;	32 0 35	8 2 43	3 0 23		2 0	3 1 724	0 0 11	2 5	
Fargo Grand Forks South Dakota:	9 0	0	0		0	8	5 9	0	
Aberdeen Nebraska: Omaha	3 12	0	0			1	2		
Kansas: Topeka Wichita	6 6	2 3	8 0 1	1	0 1 0	0 1 0	0 0 0	6 2 7	
SOUTH ATLANTIC									
Delaware: Wilmington	3	1	1		0	1	0	3	
Maryland: Baltimore Cumberland Frederick	105 0 0	30 1 0	13 0 1	9	1 0 0	20 0 0	7 0 0	23 0 0	
District of Columbia: Washington Virginia:	17	17	13	. 1	1	16	0	8	
Lynchburg Norfolk Richmond Roanoke	3 8 2 10	3 3 9 3	1 - 0 - 4 - 2 -		0 0 0 1	1 0 7 0	1 0 4 0	0 2 3 2	
West Virginia: Charleston Wheeling North Carolina:	3 10	12	0	1	1 0	0	6 0	1 0	
Raleigh	2 6		1		0	0	0 1	2 3	
Charleston Columbia	0 20	1	1	35	0	0 3	0 10	6 10	
Jeorgia: Atlanta Brunswick Savannah	3 0 0	5 0 1	4 0 1	18 12	2 1 3	19 0 0	0 0 1	8 0 0	
Miami	•2	3	3		0	0	2	1	
St. Petersburg Tampa	0	0 2	5		8	2	0	1 0	

City reports for week ended December 20, 1930-Continued

Division, State, and city		Diph	theria	Influ	lenza			Pneu-
	Chicken pox, cases reported	Cases, estimated expect- ancy	Cases reported	Cases reported	Deaths reported	Measles, cases re- reported	Mumps, cases re- reported	monia, deaths reported
BAST SOUTH CENTRAL								
Kentucky: Covington	1	1	1		0	0	2	
Tennessee:	19				3	0	2	
Memphis Nashville Alabama:	3	6 2	6 3		0	4	Ö	
Birmingham	12	5	3	6	1	42	0	1
Mobile Montgomery	0 1	1 2	1 0		1	0 0	0	
WEST SOUTH CENTRAL								
Arkansas: Fort Smith	1	2	0			1	0	
Little Rock	5	2	ŏ		0	Ô	ŏ	1
New Orleans	1	13 2	17	2	4	1	0	19
Shreveport Oklahoma:			2		0		0	
Muskogee Oklahoma City	6 0	2 3	3	5	Ő	1	Õ	(7
Tulsa Texas:	8	4	5			7	1	
Dallas Fort Worth	26 10	14 5	15 1	1	02	3 0	7	4
Galveston Houston	2 0	17	7 10		0	0	0	0
San Antonio	ĭ	5	· 6		2	ŏ	ŏ	6
MOUNTAIN								
Montana: Billings	2	0	0		o	0	0	C
Great Falls	6 10	0 0	Ō		0	0	0	Č
Helena. Missoula	0	1	Ŭ		ŏ	ŏ	ŏ	0
Idaho: Boise	3	1	0		0	0	0	2
Colorado: Denver	34	8	0		2	7	3	16
Pueblo New Mexico:	2	ĭ	Ŏ		Ō	10	Ó	ĩ
Albuquerque	17	0	1		0	0	1	1
Arizona: Phoenix	0	0	0		0	0	0	3
Utah: Salt Lake City	20	4	2		0	2	3	6
Nevada: Reno	0	0	0		0	0	0	0
PACIFIC								
Washington:		_				0	20	
Seattle Spokane	19 1	52	18 0			Ō	0	· · · · · · · · · · · · · · · · · · ·
Tacoma Oregon:	8	3	4		0	0	1	2
Portland	27	11	0		<u>o</u>	2	10 1	8 0
Salem California:	0	0	0		0		1	-
Los Angeles	16 9	40 2	14 1	41	4	2	8	36 6
San Francisco	14	17	4		ŏ	ĭ	4	ž

City reports for week ended December \$0, 1930-Continued

28443°---31-----3

	Scarle	t fever		Smallp)I	Tuber-	Ту	phoid f	Whoop-		
Division, State, and city	Cases, esti- mated expect- ancy	Cases re- ported	Cases, esti- mated expect- ancy	Cases re- ported	Deaths re- ported	re-	Cases, esti- mated expect- ancy	Cases re- ported	Deaths re- ported	ing cough, cases re- ported	Deaths, all causes
NEW ENGLAND											
Maine: Portland	2	5	0	0	0	1	0	1	0	8	19
New Hampshire: Concord	0	0	0	0	0	0	0	0	0	0	10
Vermont: Barre	o	0	0	0	0	0	0	0	0	0	15
Burlington Massachusetts: Boston	1 70	0 53	0	0	0	7	0 1	1	0	27	94
Fall River Springfield	28	- 8 - 12	Ŭ O	Ŏ	Ŭ 0	32	Ô	0 1	0 1	17	29 37
Worcester Rhode Island:	11	13	0	0	0	3	0	0	0	1	56
Pawtucket Providence	1 8	8 15	0 0	0 0	0 0	02	0	0 0	0 0	1 7	15 76
Connecticut: Bridgeport Hartford	9	12	0	0	0	2	0	0	0	2	34
New Haven	5	4	ŏ	0	0	2	ŏ	1	Ō	3	36
MIDDLE ATLANTIC											
New York: Buffalo	26	25	0	0	0	11 103	1 9	0 5	0 1	31 134	137 1, 44 9
New York Rochester Syracuse	192 7 12	164 50 10	0 0 0	0	0 0 0	105 5 0	0 0	ŏ	0	134 14 10.	1, 11 5 69 46
New Jersey: Camden	4	7	0	0	0	0	0	0	0	6	22
Newark Trenton	17 4	23 10	0 0	0 0	0 0	5 0	1 0	0 0	0 0	17 0	96 31
Pennsylvania: Philadelphia Pittsburgh	80 36	113	0	0	0	28 11	2 1	2 0	1 0	28 14	471 173
Reading	30 2	55 2	ŏ	ŏ	ŏ	2	ō	ŏ	ŏ	0	28
EAST NORTH CENTRAL											
Ohio: Cincinnati	16	35	1	0	0	5	0	4	0	8	132
Cleveland Columbus	37 11	83 7	1	0	0 0	12 2	0	2 0	0	17 -0	160 76
Toledo Indiana:	13	11	1	1	0	6	1	0	0	3	74
Fort Wayne Indianapolis South Bend	3 10 3	1 52 6	0 5 1	0 5 0	0 0 0	0 4 0	0 0 0	2 2 0	0 1 0	0 8 2	24
Terre Haute Illinois:	3	Ő	ò	ŏ	ŏ	ŏ	ŏ	ŏ	ŏ	ő	17
Chicago Springfield	115 1	183 6	1 0	0 1	0	34 0	2 0	50	0 0	64 0	644 12
Michigan: Detroit	94	55	1	4	Q	26	1	0	0	46	262
Flint Grand Rapids_ Wisconsin:	12 11	20 	1 0	0	0	1	0 0	0	0	0	20
Kenosha Madison	2 2	1 6	0	0		0	0	0	0	0	8
Milwaukee Racine	28 6	18 3	0 0	0	0	7	0 0 0	0 0	0	12 2	119 14
Superior WEST NOBTH	3	0	0	0	0	0	0	0	0	5	9
CENTRAL											
Minnesota: Di Inth	10	1	0	0	0	1 2	0	0	0	0	23 95
Minneapolis St. Paul	50 27	13 4	1 2	ŏ	ŏ	ő	ŏ	ŏ	ŏ	11	90 57

City reports for week ended December 20, 1930-Continued

. .	Scarle	t føver		Smallpo	X	Tuber-	Ту	phoid f	Whoop-		
Division, State, and city	Cases, esti- mated expect- ancy	Cases re- ported	Cases, esti- mated expect- ancy	Cases re- ported	Deaths re- ported	culo- sis, deaths re- ported	mated	Cases re- ported	Deaths re- ported	ing cough, cases re- ported	Deaths, all causes
WEST NORTH CEN- TRAL-contd.											
Iowa:						•					
Davenport Des Moines	1 10	29	1	6 3			0	0		0	29
Sioux City Waterloo	2 2	19 0	1	02			0	0		07	
Missouri:							-	-			
Kansas City St. Joseph	15	10 5	1	0	0	4	0	0	0	5	96 20
St. Louis	33	76	Ŏ	ŏ	Ŏ	ğ	ĺ	4	Ō	6	199
North Dakota: Fargo	1	0	0	0	0	0	0	0	0	0	5
Grand Forks South Dakota:	1	0	0	0			0	0		0	
Aberdeen	0	0	0	.0			0	0		0	
Nebraska: Omaha	5	11	1	16	0	2	0	0	0	1	44
Kansas:								0			
Topeka Wichita	2 5	0 5	0	0 7	0	2	0	ŏ	0 1	03	18 47
SOUTH ATLANTIC											
Delaware:	2		o	0	0	2	0	o	0	2	32
Wilmington Maryland:		4					-		-		
Baltimore Cumberland	28 1	34 2	0	0	0	17 0	2 0	2 0	0	2 0	188 9
Frederick	Ô	õ	ŏ	ŏ	ŏ	ŏ	ŏ	ŏ	ŏ	ŏ	5
District of Colum- bia:	•										
Washington	23	22	0	0	0	14	1	2	0	0	132
Virginia: Lynchburg	0	0	0	0	0	0	0	0	0	0	6
Norfolk Richmond	17	1 8	0	0	0	2 2	0	0	0	0	46
Roanoke	3	ĭ	ŏ	ŏ	ŏ	ĩ	ō	ŏ	Ŏ	Ō	13
West Virginia: Charleston	1	0	0	0	0	1	0	1	0	0	20
Wheeling North Carolina:	2	3	0	0	0	1	0	0	0	0	15
Raleigh	0		0				0				
Wilmington Winston-Salem	$\frac{1}{2}$	2 1	0	0	0	03	0	0	0	3 0	12 16
South Carolina:	0	0	0	0	0	0	0	· 0	0	0	24
Charleston Columbia	1	2	ŏ	ŏ	ŏ	3	ŏ	ŏ	ŏ	ŏ	31
Georgia: Atlanta	5	19	0	0	0	5	0	0	0	4	85
Brunswick	0	0	Ó	0	Ō	0	Ō	0 0	0	0	6 30
Savannah Florida:	1	3	1	0	0	1				-	
Miami	2 0	1	0	0	0	1	0	0	0	0	32 10
St. Petersburg. Tampa	1	3	ŏ	0	ŏ	ĭ)	ŏ	1	ŏ	0	28
EAST SOUTH CENTRAL											
Kentucky:											
Covington Tennessee:	2	6		0	0	0	0	0	1	0	19
Memphis	6	17	0	0	0	6	0	5	0	0	90 46
Nashville Alabama:	3	0	0	0	0	0	0	1	1	0	
Birmingham	4	6 3	1	0	0	7	1	0	0	0	61 24
Mobile Montgomery	i	1	ŏ	ŏ			ŏ	ŏ.		ĭ	
WEST SOUTH CENTRAL											
Arkansas:											
Fort Smith	0	0	0	0	1		0	0 .	· · · · · ·	3	

City reports for week ended December 20, 1930-Continued

Sanuary 9, 1981

80

	Scarle	t fever	Smallpox			Tube	r- ?	ryphoid	fever	fever Whoop			
Pirision, State, and city	Cases, esti- mated expect- ancy	Cases re-	Cases, esti- mated expect- ancy	re-	d por	•	re-	s esti mat	d re-	s Deaths re- d ported	ing cough,	Deaths, all causes	
WEST SOUTH CENTRAL—COD.					-								
Louisiana: New Orleans Shreveport	8	12	0	0)	0	1		2	7 2	1	149	
Oklahoma: Muskogee Oklahoma	1	0	0	6		0	(D	0	0 0	0		
City Tulsa	2 2	6 7	0	2		0				0	0	49	
Texas: Dallas Fort Worth Galveston Houston San Antonio	7 3 1 3 2	3 3 0 5 0	1 1 0 1 0	0 5 0 4 0		0000000					14 0 0 0 0	58 55 8 58 64	
MOUNTAIN Montana: Billings Great Falls Hekna Missoula Idaho:	1 2 1 1	0 8 0 0	0 0 0 0	11 0 2 0		0000		B B			8 8 0 1	6 7 5 1	
Boise Colorado:	0	1	1	0		O	6		o (0	4	10	
Denver Pueblo	13 2	18 1	0 0	0		00	0 2				15 7	94 7	
New Mexico: Albuquerque Arizona:	0	1	0	0		0	C			0	0	6	
Phoenix	1	0	0	0		0	4			0	. 0	17	
Salt Lake City. Nevada:	2 0	6 0	1 0	0 0		0	3				11 0	3 8 9	
Reno PACIFIC	U	0	U	U		Ů	1		' '		U		
Washington: Seattle Spokane Tacoma Oregon:	8 8 3	11 5 7	1 4 4	0 3 2) 0		11 0 2	20	
Portland Salem	7 0	2 2	7 0	0 0		00	3 0				1 0	75	
California: Los Angeles Sacramento San Francisco.	32 2 17	14 0 4	1 0 2	0 0 0		0 0 0	31 1 9		1	0	8 0 13	304 188	
<u></u>			ingococ eningiti		Lethargic cephaliti		en- is	Pellagra		Poliomyelitis (iu paralysis)			
Division, State, a	Division, State, and city		s Dea	ths (Cases De		aths	Cases	Deaths	Cases, esti- mated expect- ancy	Cases .	Deaths	
NEW ENGLAI Massachusetts: Boston Worcester			0	0	0		0	0	0	0 0	23	0	
Pennsylvania:		-	7 1	2 0	0		0	0 0	- 0	1	0	0 0 0	
			1 2	0	0	Y.;	0 0	0	0	0	0	Dal	

City reports for week ended December 20, 1930-Continued

¹ Typhus fever: 3 cases; 1 case at New York City, N. Y.; 1 case at Savannah, Ga.; and 1 case at Dallas, Tex.

	Menin men	gococcus ingitis	Letha ceph	rgic en- alitis	Pel	lagra	Poliom	yelitis (i paralysis	i nfantile ;)
Division, State, and city	Cases	Deaths	Cases	Deaths	Cases	Deaths	Cases, esti- mated expect- ancy	Cases	Deaths
EAST NORTH CENTRAL									
Obio: •									
Cincinnati Cleveland	03	01	0	0	0		0		
Indiana		-	Ŭ	, v	v	U U	ľ		l u
Indianapolis	0	1	0	0	0	0	0	0	0
Illinois: Chicago	4	o	0	l o	0	i o	0	1 1	a
Michigan:	-								-
Detroit Wisconsin:	6	2	0	0	0	0	0	0	0
Milwaukee	0	0	0	0	0	0	0	1	0
	-			-	-			_	
WEST NORTH CENTRAL									
Minnesota:			•		•				
Minneapolis St. Paul	0	0	0	0	0	0	0	3 1	0
Iowa:	v	v			U	v	v		U
Des Moines	1	0	0	0	0	0	0	0	0
Waterloo Missouri:	1	0	0	0	0	0	0	0	0
Kansas City	1	1	0	0	0	0	0	1	0
St. Joseph	1	0	0	0	0	0	0	0	0
St. Louis	3	2	0	0	0	0	0	0	0
SOUTH ATLANTIC									
Delaware:									
Wilmington	1	0	0	0	0	0	0	0	0
Maryland:	1	0	0	0	· 0		0	0	•
Baltimore District of Columbia:	1	U		U		0	U U	v	0
Washington	0	1	0	0	0	0	0	1	0
Virginia: Norfolk	0	o	o	0	0	o	0	1	0
South Carolina:	•	v	U U	v I	•	v I	v	- 1	v
Charleston	1	0	0	0	0	0	0	0	Ç
Columbia	0	0	0	0	0	2	0	0	9
Georgia: 1 Atlanta	1	1	0	0	0	0	0	· 0	0
	_	-	-			-	-		
EAST SOUTH CENTRAL									
Fennessee: Memphis	1	2	0	o	0	0	0	o	0
	-	-1	Ŭ,	°		Ĩ	•	-	•
WEST SOUTH CENTRAL								1	
Arkansas:								0	
Little Rock Louisiana:	0	0	0	0	0	1	0		0
New Orleans	1	1	0	0	2	2	0	0	0
Oklahoma:	0			0	0	0	0	1	0
Oklahoma City	U	0	0	° I	v I		v	-	U
Dallas 1	0	0	0	0	1	1	0	0	0
Fort Worth	1	0	• 0	0	0	0	0	0	0
MOUNTAIN					1				
Montana:									
Missoula	1	0	0	0	0	0	0	0	0
Phoenix	0	2	0	0	0	0	0	0	0
Jtah:	۷I	-	v I	v		۳I	1		-
Salt Lake	2	0	0	0	0	0	0	0	0
PACIFIC	- 14 I			1					
alifornia:	1						1		
Los Angeles	0	0	0	0	2	0	0	0	0
Sacramento	1	0	0	0	0	0	0	1 6	0 1
San Francisco	1	0	0	U	U I	U		U	1

City reports for week ended December 20, 1930-Continued

¹ Typhus fever: 3 cases; 1 case at New York City, N. Y.; 1 case at Savannah, Ga.; and 1 case at Dallas, Tex.

January 9, 1981

The following tables give the rates per 100,000 population for 98 cities for the 5-week period ended December 20, 1930, compared with those for a like period ended December 21, 1929. The population figures used in computing the rates are approximate estimates, authoritative figures for many of the cities not being The 98 cities reporting cases have an estimated aggregate populaavailable. tion of more than 32,000,000. The 91 cities reporting deaths have more than 30,500,000 estimated population.

Summary of weekly reports from cities November 16 to December 20, 1930-Annual rates per 100,000 population, compared with rates for the corresponding period of 19291 DIPHTHERIA CASE RATES

					Week e	nded—				
	Nov. 22, 1930	Nov. 23, 1929	Nov. 29, 1930	Nov. 30, 1929	Dec. 6, 1930	Dec. 7, 1929	Dec. 13, 1930	Dec. 14, 1929	Dec. 20, 1930	Dec. 21, 1929
98 cities	102	2 186	89	139	3 92	146	4 90	134	\$ <u>\$</u> 6	128
New England	113	117	80	177	111	112	117	117	6 130	168
Middle Atlantic	54	123	50	123	61	110	50	112	65	106
East North Central	125	302	123	167	113	191	7 122	170	\$ 120	167
West North Central	108	169	108	114	99	121	95	148	87	110
South Atlantic	141	135	60	144	9 104	127	• 113	107	° 91	107
East South Central	310	239	155	157	162	226	155	137	94	123
West South Central	183	446	164	259	10 159	362	11 147	2.3	10 219	225
Mountain	26	2 89	77	17	12 0	157	26	61	17	61
Pacific	73	60	111	56	76	84	64	58	97	56

MEASLES CASE RATES

98 cities	129	3 72	109	74	3 146	98	• 167	113	^{\$} 194	109
New England Middle Atlantic East North Central South Atlantic East South Central West South Central Mountain Pacific	164 80 31 751 59 169 4 318 33	56 34 94 81 24 14 27 3 107 280	148 73 28 636 40 74 11 275 12	70 33 101 100 22 0 38 131 249	202 89 28 933 \$57 175 10 12 13 51 31	81 54 93 216 4 14 46 165 377	250 89 7 27 1,055 9 74 337 11 8 146 31	85 47 133 202 28 14 61 104 464	6 173 91 8 29 1, 387 9 128 310 10 20 163	92 59 94 210 39 0 133 139 418

SCARLET FEVER CASE RATES

	200	³ 218	178	212	¥ 207	252	4 229	277	\$ 236	249
New England	217	249	241	258	246	276	237	375	* 312	310
Middle Atlantic	168	127	156	116	187	148	196	172	219	176
East North Central	266	347	224	361	259	409	7 318	438	300	355
	214	223	137	183	194	231	205	271	273	235
	198	163	172	139	• 211	159	9 241	193	193	253
East South Central	236	157	243	137	337	144	425	89	223	48
West South Central	101	156	142	118	10 100	156	11 94	137	1 (80	99
Mountain	275	3 267	223	348	¹⁹ 120	392	206	322	292	583
Pacific	102	261	97	266	113	355	83	340	97	244

¹ The figures given in this table are rates per 100,000 population, annual basis, and not the number of cases reported. Populations used are estimates as of July 1, 1930, and 1929, respectively. ases reported. Populations used are estimates as of July 1, 1930, and 1929, respectively. * Reno, Nev., not included. * Baleigh, N. C., Shreveport, La., and Denver, Colo., not included. * South Bend, Ind., Raleigh, N. O., Fort Smith, Ark., and Shreveport, La., not included. * Hartford, Conn., Grand Rapids, Mich., Raleigh, N. C., and Shreveport, La., not included. * Hartford, Conn., not included. * South Bend, Ind., not included. * Grand Rapids, Mich., not included. * Baleigh, N. C., not included. * Bahreveport, La., not included. * Bohreveport, La., not included. * Bort Smith, Ark., and Shreveport, La., not included. * Denver. Colo., not included.

¹⁰ Denver, Colo., not included.

Summary of weekly reports from cities November 16 to December 20, 1930-Annual rates per 100,000 population, compared with rates for the corresponding period of 1929 -Continued

					Week e	nded-				
	Nov. 22, 1930	Nov. 23, 1929	Nov. 29, 1930	Nov. 30, 1929	Dec. 6, 1930	Dec. 7, 1929	Dec. 13, 1930	Dec. 14, 1929	Dec. 20, 1930	Dec. 21, 1929
98 cities	3	3 24	8	14	\$7	19	4 15	23	\$ 9	23
New England Middle Atlantic. East North Central West North Central. South Atlantic. East South Central West South Central Mountain. Pacific.	0 0 33 0 0 4 43 7	0 333 50 2 0 38 371 111	0 0 4 66 0. 0 4 34 9	0 13 48 0 0 11 35 75	0 0 1 47 • 0 0 10 4 13 205 12	0 0 26 64 0 19 78 60	0 73 120 •0 118 146 7	2 0 29 56 0 0 34 78 118	⁶ 0 0 86 47 90 0 ¹⁰ 16 112 12	0 0 31 60 0 7 34 52 113

SMALLPOX CASE RATES

TYPHOID FEVER CASE RATES

98 cities	15	* 13	10	5	3 10	5	48	6	19	5
New England Middle Atlantic East North Central South Atlantic. East South Central West South Central Mountain Pacific	15 5 9 23 26 13 90 51 12	11 10 9 12 19 34 34 2 36 5	11 3 4 8 29 13 75 9 7	2 5 6 4 34 15 26 2	7 8 10 6 9 17 13 10 28 12 17 12	2 4 4 2 6 48 0 26 10	18 7 7 6 4 20 11 25 0 7	7 6 3 6 7 14 8 9 7	6 10 3 8 10 8 11 40 10 28 9 7	0 4 3 8 4 0 38 17 2

INFLUENZA DEATH RATES

91 cities	11	38	9	11	¥ 10	17	13 10	16	\$ 10	19
New England Middle Atlantic East North Central South Atlantic East South Central West South Central Mountain Pacific	7 8 5 6 22 15 38 60 9	4 9 6 9 4 30 16 29 6	2 11 7 0 9 29 15 26 9	4 5 10 21 17 15 55 17 13	4 6 8 12 9 19 15 10 37 13 34 3	11 14 9 27 28 60 47 17 13	4 8 7 5 21 9 22 29 10 12 9 9	7 9 15 12 19 60 78 0 19	6 2 5 8 10 15 9 19 37 10 25 17 12	9 18 14 15 13 52 66 26 26 28

PNEUMONIA DEATH RATES

			1		1		1	1	:	
91 cities	119	* 101	112	106	\$ 102	136	13 108	150	\$ 115	158
	110		71	92	66	74	109	135	¢ 108	157
New England	115	88	71							
Middle Atlantic	140	106	125	101	107	139	109	156	133	165
East North Central	83	96	78	84	78	126	7 85	116	\$ 70	117
West North Central	136	102	92	126	130	126	145	174	95	180
South Atlantic	143	94	165	129	° 143	131	9 121	191	9 128	184
East South Central	199	254	155	224	177	239	140	216	125	216
West South Central	123	129	165	156	10 139	238	10 176	230	10 147	234
Mountain	163	3 107	223	157	18 137	165	154	192	215	235
Pacific	61	28	86	104	74	138	74	107	156	138

Reno, Nev., not included.
Raleigh, N. C., Shreveport, La., and Denver, Colo., not included.
South Bend, Ind., Raleigh, N. C., Fort Smith, Ark., and Shreveport, La., not included.
Hartford, Conn., Grand Rapids, Mich, Raleigh, N. C., and Shreveport, La., not included.
Hartford, Conn., not included.
Grand Rapids, Mich., not included.
Raleigh, N. C., not included.
Bartort, La., not included.
Shreveport, La., not included.
Bouth Bend, Ind., Raleigh, N. C., and Shreveport, La., not included.

FOREIGN AND INSULAR

CANADA

Provinces—Communicable diseases—Week ended December 20, 1930.—The Department of Pensions and National Health of Canada reports cases of certain communicable diseases for the week ended December 20, 1930, as follows:

Provinc o	Cerebro- spinal fever	Influenza	Polio- myelitis	Smallpox	Typhoid fever
Prince Edward Island 1					
Nova Scotia ¹ New Brunswick					7
Quebec Ontario		5	1	1	11 10
Manitoba Saskatchewan	_ 1				1
Alberta			1		
					6
Total	1	5	8	1	. 36

¹ No case of any disease included in the table was reported during the week.

Quebec Province—Communicable diseases—Week ended December 20, 1930.—The Bureau of Health of the Province of Quebec, Canada, reports cases of certain communicable diseases for the week ended December 20, 1930, as follows:

Disease	Cases	Disease	Cases
Chicken pox	194	Mumps	27
Diphtheria	66		1
Erysipelas	8		94
German measles	2		46
Influenza	5		11
Measles	56		31

JAMAICA

Communicable diseases—Four weeks ended December 6, 1930.— During the four weeks ended December 6, 1930, cases of certain communicable diseases were reported in Kingston, Jamaica, and in the Island of Jamaica outside of Kingston as follows:

	Ċ	ases		Ca	S65
Disease	Kings- ton	Other localities	Disease	Kings- ton	Other localities
Corebrospinal meningitis Chicken pox Dysentery Erysipelas Leprosy	1 1 1 1 1	1 4 3 1 1	Poliom yelitis Scarlet fever Tuberc: losis Typhoid fever	1 43 20	2 1 66 48

From medical officers of the Public Health Service, American consuls, International Office of Public Hygiene, Pan American Sanitary Bureau, health section of the Lagua of Nations, and other sources. The reports contained in the following tables must not be considered as complete or final as regards either the list of countries included or the figures for the particular countries for which reports are given.

CHOLERA

[O indicates cases; D, deaths; P, present]

		-							We	Week ended—	1						
Place	June 29- July 26, 1020	July 27- Aug. 23-1030	Aug 24- Sept.	Sept.	0	October, 1930	1930			пөтол	November, 1930	8		Å	December, 1930	r, 1030	
		0001 (07	0007 °0	1930	4	Ħ	18	ส			15		8	-	81	8	R
A fçhanistan. China:	A O	4															
			20		Ì	Ì	İ	-					Ť	İ	İ	İ	
Canton			N		-	$\frac{1}{1}$	İT	-	$\frac{1}{11}$	$\frac{1}{11}$			Ħ	$\frac{1}{1}$	$\overline{\Pi}$	T	
			34	82	20	9	4	-	$\frac{1}{11}$	-				T	ÎÌ	İİ	
Shensi Province			Б	2			8		$\frac{1}{1}$					T		İİ	
Swatow Tientsin			- 7										-			İ	
India	D 26, 121	42, 893 22, 358	51, 551	11, 109 1	10, 172 7	7, 863											
Bassein			1 1	· · · ·		5	ļ,		-				-		Ħ	Ħ	
			- 01		, 0 01	-	; ∞	19	- 60			20	2-1	*	Ī	Ī	
	R R D D D	88 80	27	40	4 4	1-4	04	l- 4	101	19	00 4	44	60 44	~~	T		
						010			-							Ī	
Rangoon		(1 -	- 7			 	Π						Ī		İİ	ÌÌ	
Tuticorin			-		Īİ	•						-	$\overline{\Pi}$	T	Ħ	İİ	
India (French): Chandarnacor			-		-		-				-		-			İ	
Pondicharty			•			-								-	-	İİ	
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FEVER-Continued
YELLOW
AND
FEVER,
TYPHUS
SMALLPOX,
PLAGUE,
CHOLERA ,

CHOLERA—Continued [U indicates cases; D, deaths; P, present]

January 9, 1931

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Indo-Chira (French) (sea also table above): Annam 1		23 16 88 144 671 273	4 40 4 40 4 0	37	° 13 3		6 33	66 66	18	16 14	8	60 60		
i Pianas (an abalan in the Dhilinnine Islands and an his														

¹ Figures for cholera in the Philippine Islands are subject to correction. ⁴ During the period from Aug. 24 to Sept. 26, 1930, 28 cases of cholera with 17 deaths were reported in Manitum, Surigao Province, P. L. ⁴ Reports incomplete.

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January 9, 1981

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FEVER-Continued
YELLOW
, AND
FEVER
TYPHUS
SMALLPOX,
PLAGUE,
CHOLERA,

PLAGUE [C indicates cases; D, deaths; P, present]

									M	Week ended	ded						
Place	July 26,	27- 27- Aug.	Aug. 24- Sept.	Sept.	0	October, 1930	, 1930			Nover	November, 1930	930		Å	cemb	December, 1930	
	ncet	0001 1000		27, 1930	4	=	18	52		80	15	ន	8		13	ล	5
Algeria: Algera:	°	2	=	-	-	6	69	<u>v</u> or	~	- 1	61						
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	2		-91	16		1 8								$\overline{111}$	ŤŤ		
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	213	229 1	191	- 65	× -	8	40	8									
China: Ch	- ~	101		,	••	$\overline{\Pi}$	•••							• ~~			
Mancharia-Tangliau and Nungan		30	8ª-	67	T	ŤŤ		P.						\uparrow			
Durbu bask indice: Batavia and West JavaC D	88 84	88-	62 92	222	14 14	ສສ	45	41	36 37	58							
Java and Madura. Ecuador (see table below). Eguador (see table below).	217	188	260	75	89	95	67	124	140	107	130			İ İ			
Alexandria.	8999	11 9	10 8	36		ю-н	∞ ⊣	86	-	00 FO FT	69	$\overline{\Pi}$	- 8		4-1-	1	
Aswan. Beni-Suef	8	1			İİİ	ÌÌÌ				-	$\frac{1}{11}$	$\frac{1}{11}$	8	Tİİ	-		

January 9, 1931

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PLAGUE-Continued

[C indicates cases; D, deaths; P, present]

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Place			Ju ₂ J	June J 29- July 26, A	July 27- Aug.	Aug 24- Sept.	Sent.	Oct	October, 1930	930		No	November, 1930	, 1930			Decem	December, 1930	
					,		27, 1530	4 11		18 2	1	∞	15	8	8 	•	13	8	2
Tunisia: Sfax district Tunis			00	 ອາອ													L L		
Union of Socialist Soviet Republics: Salsk Region	egion		סטס	- 10 4	5														
Union of South Africa: Cape Province			υp							1	1								
Orange Free State			DOD		'	10101													
Place	May, 1930	June, 1930	July, 1930	Aug., 1630	Sept., 1530	Oct., 1530			Id	Place			May, 1930		June, J 1930	July, 1930	Aue., 1930	Sept., 1930	Oct., 19 30
British East Africa (see also table above): Kenya Ecuador: Guayaquil	121	107	18	87	53	50		Madagascar (see also table above)—Con Tananarive Province	(see al ve Pro	so tabl	e above)	L Con		14	16	***	88	88	
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Autosuta Frovince		000	24 24 1	22008	52~~5			Thies 1 Trhies 1 Traouane 1	101					88°°73	828888	35388292 2	88883 2	282483	5255 <i>56</i>
	-	0 m -		¥53	9 <u>1</u>														

January 9, 1931

SMALLPOX

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Flace	June 1-28, 1930	June 29–July 26, 1930	July 27- Aug. 23, 1930	Aug. 24- Sopt. 20, 1930	Sept. 27.	-0	October, 1930	1930			Novet	November, 1980	086		December, 1930	
					1930	4	Ħ	18	52	1	80	15	ន	×8.	8	2
Algeria: Algeria:		-						, 								1
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British East Africa (see also table below): Tanganyika	1.610	168	242	522	2	5	•	21		-	29					
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	4554	24 13	- ⁸	-92-	-		ø	19	7	8	33	=	-	200	∞ «	
Quebec. 0 Montreal. 0 Baskatchewan. 0	4	ся ю 	NC 1- 00	-		6				5					1	
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	n 4 8	- നജ	6												P4	
	P 1	Ρ	Ρ	Р	Ь	Ч	Ч	Р	Ъ	<u> </u>	P.	$\frac{1}{1}$				
Foreigners only	5004	4	4 3	80°				3 1			-				~	
					-			-		+	+		-	1		

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SMALLPOX—Continued

[C indicates cases; D, deaths; P, present]

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									We	Week ended-	Ļ					
Place	June 1-28, 1930	June 29–July 26, 1930	July 27- Aug. 23, 1930	Aug. 24- Sept. 20, 1930	Sept. 27.	0	October, 1930	1930		Z	0 Vem	November, 1930	8	Ă	December, 1930	ъ,
					1930	4	=	18 25				15 2	80 23	8		13
Chosen (see table below). Colombia: Barrannilla																
	9	10	101			-		-								
Costa Rica: Port Limon. Curacao (alastrim).	- 19	2	1			•		•								
Dutch East Indies: Borneo	12	61								-+		_			-	•
Java- Batavia and West Java C	13 1	00	12	Ħ	10	61.			010	13						
East Jaya and Madura. C Saneri Islands	n I	°1	36-36-	4	-	-	-	-	N 0	-						
			6	ŝ					~							
France Steel table below). Great France England and Wales	556		344	341	82	74	44	125		87	53	107	116	160	8	1
Bradford Cardific Construction	-					-	1	-		-						
Leeds. London and Great Towns.	498 753	408 408	178 285	268 268	83	2 5 25	នន	88	28-	88	48	87°	882	1961	57	
Btoke-on-TrantC Beotland	32	61	61	210												
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India.	12,962 3,531	7, 630 2, 348	4, 877	3, 131	611 202	88 12 28	567 113		\square					\mathbb{H}		

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SMALLPOX-Continued

[C indicates cases; D, deaths; P, present]

	•					•										
					 -					Week	Week ended-					
Place		June 1-28, 1930	June 29-July 26, 1930	July 27- Aug. 23, 1930	Aug. 24- 24- Sept. 1930	Bept. 27.	Oct	October, 1930	8		Nov	November, 1930	1930		December, 1930	and a set
						1930	4	11 18	52	-	∞	15	8	8	•	
Somaliland, British: Boales											-					
Spain Straits Settlements		9	5 11	C1 00	100				2	10.00	=*	21	≈	21 0		
Sudan (Anglo-Egyptian)	204	8°9	° 21 •	~ <u>9</u> ,	128-1	25	1	- 23		-	-	101	-	0100		4
Budan (French) (see table below). Switzerland: Berne Canton	<u>, c</u>	ø		~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~~	20	- 8	ŀ	•	<u> </u>		<u> </u>		İ		1	Ģ
Syria (see table below). Turis (see table. Turisia: Turis	: > 0 : 1	8	•	•		- 61										
Turkey (see table below). Union of South Africa: Development	C		9		م			Ą	ب م	م م	ρ	٩				
Orange Free State			- Р-	P. P	- P	ρ		_			4 P	4 P4 P	İİ		$\frac{1}{11}$	
Upper Volta. On vessel: S. S. Manca, from Honolulu to San Francisco.	000	13	1	4	4	4						-				
	May.	June.	.4lul		August, 1930	930		September, 1930	er, 1930		Oct	October, 1930	30	Nove	Novemb er , 1980	1080 1
entra .	1930	1930	1930	1-10	11-20	21-31	1-10	11-20		21-30	1-10	11-20	21-81	1-10		11-20
Indo-China (see also table above)	305	213	338	89	8			12	22	8	82	62	21			8.
Sudan (Franoh).	274	20		-		8°			<u>е</u> ,			5.			\mathbf{H}	•
Syria: Beirut	1			64					<u> </u>		İ	4				

Place	May, 1930	May, June, July, Aug., Sept., Oct., 1930 1930 1930 1930 1930 1930 1930	July, 1930	Aug., 1930	Sept., 1930	Oct., 1930	Place	May, 1930	June, 1930	July, 1930	Aug., 1930	May, June, July, Ang., Sept., Oct., 1930 1930 1930 1930 1930 1930	1660 1660
British East Africa (see also table above): Kenya Uganda	171 78	142	142 186 424		424	424	France. Mexico: Durango(see also table above). D Morocco.	12 4 81	6.0	~~~~	60 60	[ci	
	8623		8	2		2	Turkey	16					
SeishinD	301-	-	1	2		, , , ,							
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TYPHUS FEVER

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•					27, 1930	4	11	18	25	1	8	15 2	22	8		13	ଛ
Algeria: Algeria: C	~	6	5	~							-			8	 ·	 	
Constantine Department	514	c3 c3	- m	1		61	-	5	-					0 T-1 CO			
Bulgaria. C	16	10	1	4	2		-	; m			-1	63	-				
China: Manchuria—Harbin (see also table below)	00		21	5													
Chosen (see table below). Czechoslovakta (see table below). Czechoslovakta (see table below).			-1					·		<u> </u>							
Egypt: AlexandriaC Behaira ProvinceC	45	15	1	ю - 1	1							61					ſ
D Cairo	4	5	2	5		1			<u></u>								-
Great Britain: Scotland-Dunfermline		1	*	1	**************************************	•		Π	1							$\frac{1}{11}$	

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TYPHUS FEVER—Continued

[C indicates cases; D, deaths; P, present]

										Week ended	pepue						
Flace	Jun e 1-28, 1930	June 29-July 26, 1930	July 27-Aug. 23, 1930	Aug. 24-Sept. 20, 1930	Sept.	Ō	October, 1930	1930			Чотеп	November, 1930	8		December, 1930	lber, 1	8
					1930	4	11	. 81	55			15	22	8	9	13	8
Greece (see table below). Ireland: Irish Free State—									<u> </u>				l				
	6	8					<u></u> -			<u>+-</u> -						<u>††</u>	
Casteber Westport Recommon County-		- 5 -														İŤ	
Roscommon Rtokkow County-Shillelegh Letvis (see table below)	4	9															
ties in Federal Dis-				1 1-													
	15.1 %	° I	Do 69 00	~ 61 63			1 1	• -		-		* T	- 01	- ~ -		ÎÌÌ	
Palestine Poland	117 117 11	684	or 45 ∞ -	- 23	13	-100	9		7	00	15.2		-9-9	8	200		-
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Tunisia	18	78 -	10		-	1		-16 -16				$\frac{1}{1}$		$\overline{\Pi}$	8	T	

January 9, 1981

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Sept. 1930	22	
Aug., 1830	619	2
July, 1930	18 7	nfectio
Tune, 1930	0 6 F	C bably laboratory infection).
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Place		.R Coast: Uly 10, 1980 A. Mouros, 4, 1980 (death)
		tt: 0, 1930 10, Aug fonrov agos, J
	ruania. key gosiavia	FEVER Gold Coast: July 10, 1930 - Albosso, Aug. Liberla, Monrovii Nigeria, Lagos, Ju
Nov., 1930	16 4	YELLOW FEVER Cases Gold C Jul Cases I Jul Liberia Nigeria
Oct., 1930	44	H O
	- 40	
Aug., E		
July, 1 1930	41 0 0 0 0 T	
June, 1930	8919 19	May 23, 1930.
Place	hina: Harbin (see also table above) C boen: Seoul	Brasil: Campos, Rio de Janeiro Province, May Para, June 23, 1880
	June, July, Aug., Sept., Oct., Nov., 1830 1330 1330 1330 1330 1330	June, 1uity, Aug., Bept., Oct., 1880, 1880, 1880, 1880, 1880, 1880, 1880 Nov., 1880 Nov., 1880 Nug., Sept., 0ct., 1880, 1880, 1880, 1880 Nov., 1880 Nug., 1880, 1880, 1880, 1880 Nug., 1880, 1880, 1880 Nug., 1880, 1880, 1880 Nug., 1880, 1880, 1880 Nug., 1880, 1880 Nug., 1880 Nug., 1880, 1880 Nug., 1880, 1880 Nug., 1880

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