# The Socioeconomic Distribution of Mortality Rates in Des Moines, Iowa, 1974

R. SCOTT FREY, PhD

THE ECONOMIC POSITION of a population group determines its life chances (1-3). Consider, for example, mortality: mortality rates of groups vary inversely with socioeconomic status (4-10). This generalization and the observation that cities consist of relatively distinct geographic areas inhabited by people who share similar socioeconomic characteristics (11) suggest the following hypothesis: the lower the socioeconomic status of the population in an area, the higher the mortality rates. This hypothesis is consistent with research findings reported over the past 100 years for cities in this country and abroad (5,7,8, 12-36).

Despite a substantial body of research examining the posited relationship, relatively little research has been based on data collected since 1970 (27, 29-36). Moreover, the findings of several recent studies suggest an increase in socioeconomic differentials of various cause-specific forms of mortality over the past few decades (34,35). Although it is not possible to examine the validity of the reported increase in socioeconomic differentials, a consideration of data collected since 1970 would add to the understanding of the current status of the relationship. Thus, with this objective in

mind, I report the results of an examination of variations in the 1974 rates for 15 categories of death (including total mortality) across socioeconomic areas within Des Moines, Iowa.

# **Data and Methods**

Mortality data. All deaths that occurred in the incorporated city of Des Moines in 1974 were identified from records supplied by the Des Moines-Polk County Health Department. The census tract was the unit for the geographic classification of each death. In the enumeration of the 1970 census (37), Des Moines was divided into 47 tracts. Four of the 1,846 reported deaths and 1 census tract (No. 35) were omitted from this analysis because of incomplete data. Each of the remaining 1,842 deaths was assigned to 1 of the 46 tracts according to residence and a general cause of death according to the Ninth Revision of the International Classification of Diseases, Clinical Modification (38).

Socioeconomic areas. The 46 census tracts were aggregated into 3 socioeconomic areas. Three areas were constructed because (a) the small number of cases for several causes of death necessitated using a small number of socioeconomic areas and (b) previous researchers have used three areas in their analyses (14,19,30). In effect, the use of three socioeconomic areas is defensible for methodological reasons.

Aggregation of census tracts proceeded in two steps. First, Shevky-Bell (39) social rank scores were computed for each tract according to the following formula:

$$SRS_i = OSS + ESS \div 2$$

where,  $SRS_i = \text{social rank score}$ for census tract i ( $i = 1, \ldots, 46$ ); OSS = 100 - [x (r - o)], OSS= occupation standardized score for census tract in 1970, r = occupation ratio score for census tract (number of craftsmen, foremen and kindred workers, operatives except transport, and laborers except farm per 1,000 employed persons), x = $100 \div$  range of occupation ratios (r) for all census tracts, and o =lower limit of distribution of occupation ratios (r) for all census tracts: ESS = 100 - [x(r - o)]. ESS = education standardized score for census tract in 1970, r = education ratio score for census tract (number of persons completing a maximum of 8 years of education per 1,000 persons 25 years and older),  $x = 100 \div$  range of education ratio (r) for all census tracts, and o = lower limit of distribution of education ratio (r) for all census tracts. Shevky-Bell social rank scores were used because they were used by previous researchers (16,17,25,26). Second, tracts were ranked from low to high according to social rank scores and then ag-

Tearsheet requests to Dr. R. Scott Frey, Assistant Professor, Department of Sociology, George Washington University, Washington, D.C. 20052.

gregated into three groups, consisting of a relatively equal number of tracts and representing areas of low (I), middle (II), and high (III) socioeconomic status. The three socioeconomic groups, social rank scores, population figures, and number of deaths are presented in table 1.

Standardized mortality ratios. Sexspecific standardized mortality ratios (40) were computed for each of 15 general categories of death (table 2) by socioeconomic area according to the indirect method. The formula for the ratio follows:

$$SMR_{i} = \frac{9}{\substack{j = 1 \\ j = 1}} \times 100$$
$$\sum_{j = 1}^{\infty} M_{j}p_{j}$$

where,  $SMR_i$  = standardized mortality ratio for socioeconomic area i (i = I, II, III); j = age category (1 = under 5 years, 2 = 5–14, 3 = 15-24, 4 = 25-34, 5 = 35-44, 6 = 45-54, 7 = 55-64, 8 = 65-74, 9 = 75+);  $m_j$  = death rate in age group j in 1974;  $p_j$  = population in age group j in 1970 (37);  $M_j$ = death rate in age group j in 1974 for standard population (Des Moines).

As used here, the standardized mortality ratio is an age-standardized ratio that compares the extent to which observed deaths in a socioeconomic area vary as a percentage of the expected number of deaths if age-specific rates for the city occurred in the socioeconomic area. The ratio is 100 when observed and expected deaths are equal, and a score of less or more than 100 indicates a rate deviating from the expected rate.

Shortcomings. The following shortcomings are associated with the data and methods used in the analysis:

#### Standardized mortality ratios for 15 categories of death (including total mortality) by socioeconomic area and sex. Des Moines, Iowa, 1974



—the reliability of mortality data (41) and census tract data (42),

—the use of 1970 census tract population characteristics and 1974 tract mortality rates when population characteristics of tracts may have changed between 1970 and 1974, —census tract homogeneity (43, 44),

—the manner in which mortality rates were computed (45-47),

--standardizing infant deaths (children under 1 year old) according to the population under 5 years old because data for the number of live births were not available,

-the use of a small number of

cases in computing mortality ratios for several causes of death, and

—the inability to generalize aggregate findings to the individual level (48,49).

However, I assume that these shortcomings do not vitiate the results reported here regarding the socioeconomic distribution of mortality rates. Such an assumption is realistic because the results presented in the next section are consistent with previous research findings.

# Results

The results are presented in table 2 and the chart. The data reveal

Table 1. Census tracts grouped into three socioeconomic areas by Shevky-Bell social rank (SR) scores, Des Moines, Iowa, 1970

Tract	······	• • • • • • • • • • • • • • • • • • •	SR scores	Population	Deaths	
		Area I (low)				
36 .		• • • • • • • • • • • •	.001	1.034	16	
38 .			22,961	1.393	16	
42			34,624	2.396	26	
18 .			38.431	2,828	31	
17 .			42.304	2,694	27	
21			44.562	5.696	66	
37			45.006	3,204	40	
25			46.254	647	38	
20			47.645	1.687	19	
13			54.044	4,195	48	
12			54 987	4,100	62	
10	•••••••••	••••	54 990	4,520	35	
14	•••••		55 050	2,507	39	
27	•••••	••••	55 166	2,010	30	
27 .	•••••••••	• • • • • • • • • • • • •	56 690	3,091 1 AEA	42	
20 .	•••••		50.009	_1,4 <b>0</b> 4	37	
			-	Area II (middle)		
33			57 713	2 040	40	
44	•••••	• • • • • • • • • • • •	59 070	2,545	40	
40	•••••	• • • • • • • • • • • • •	50.272	4,102	19	
40 . E	••••••	•••••	50.311	6,240 5,600	57	
5.	••••••	• • • • • • • • • • • •	58.404	5,032	52	
3.	•••••		59.405	4,025	31	
39.	•••••	• • • • • • • • • • • •	60.284	1,524	9	
43 .		• • • • • • • • • • • •	61.249	6,674	41	
15 .	••••		61.979	3,284	39	
2.	• • • • • • • • • •	• • • • • • • • • • •	62.715	7,518	60	
1.	••••••		63.465	9,069	41	
4.			63.884	5,427	55	
6.			65.428	4,697	50	
34 .			65.515	2,558	43	
16 .			65.747	2,262	43	
47.			65.928	4,934	25	
26.	•••••	• • • • • • • • • • • •	66.023	2,687	25	
		-		Area III (high)		
24.		- 	69.033	2,012	16	
45.			70.708	6,351	44	
11			73.819	5,976	43	
41 .			75,706	3.661	37	
40 .			79.870	3.694	25	
29			83,084	4.712	72	
8			84,406	12 412	92	
9			84,750	8 586	80	
7.			85 858	14 163	78	
28	•••••		85 915	A A37	53	
10		• • • • • • • • • • • •	85 059	5 AQ5	53	
21	••••••	•••••	00.000	0,400	00	
30.04	•••••		31.400	2,233	3	
30.01		• • • • • • • • • • • •	91.003	2,007	17	
20.02	•••••	••••••••••••	97.991	4,208	35	
32 .	•••••	• • • • • • • • • • •	100.000	2,573	19	

partial support for the hypothesispartial in the sense that not all forms of mortality vary inversely and monotonically with socioeconomic status. For instance, an inverse and monotonic pattern is seen for deaths of males due to infectious diseases, malignant neoplasms, nervous diseases, circulatory diseases, genitourinary diseases, infant deaths, and homicide. Females show a pattern consistent with that of males for total mortality and mortality due to infectious diseases. malignant neoplasms. diabetes. nervous diseases, circulatory diseases, infant deaths, and violence. Similarly, the total population shows an inverse and monotonic pattern for total morbidity and mortality from infectious diseases, malignant neoplasms, nervous diseases, circulatory diseases, genitourinary diseases, infant deaths, and homicide

Although several categories of death do not vary inversely and monotonically with socioeconomic status, ratios for the lowest socioeconomic area (I) are higher than the ratios for the upper two areas for most categories of death. Males from area I display higher ratios than males from areas II and III for 12 of 15 categories of death: infectious diseases, malignant neoplasms, diabetes, nervous diseases, circulatory diseases, digestive diseases, genitourinary diseases, infant deaths, violence, accidents, homicide, and total mortality. The ratios for females in area I are higher than those in the upper two areas for 11 of 14 categories of death: infectious diseases, malignant neoplasms, diabetes, nervous diseases, circulatory diseases, digestive diseases, genitourinary diseases, infant deaths, violence, accidents, and total mortality. The total population's pattern is similar to that of males and females for 12 of 15 categories of death: infectious diseases, malignant neoplasms, diabetes, nervous diseases, circulatory diseases, digestive diseases, genitourinary diseases, infant deaths, violence, accidents, homicide, and total mortality. In effect, the findings indicate that most categories of death vary as an inverse function of socioeconomic status.

#### Table 2. Standardized mortality ratios <sup>1</sup> for 15 categories of death by socioeconomic area and sex, Des Moines, Iowa, 1974

	Socioeconomic area			
Category of death <sup>2</sup>	l (low)	ll (mlddle)	lll (high)	<b>Ρ</b> (χ²)
Infectious diseases (000–139)	174	96	63	NS
Male	<sup>3</sup> 159	3 98 °	<sup>3</sup> 68	NS
Female	189	<sup>3</sup> 95	<sup>3</sup> 58	NS
Malignant neoplasms (140-239)	139	100	79	.001
Male	139	91	85	.020
Female	135	110	75	.010
Diabetes (250)	162	77	86	NS
Male	<sup>3</sup> 202	³ 38	<sup>3</sup> 104	NS
Female	150	98	76	NS
Nervous diseases (320–389)	130	101	83	NS
Male	3 1 1 1	<sup>3</sup> 104	<sup>3</sup> 90	NS
Female	<sup>3</sup> 152	<sup>3</sup> 97	<sup>3</sup> 76	NS
Circulatory diseases (390–459)	118	102	88	.010
Male	106	99	98	NS
	130	105	81	.001
Respiratory diseases (460–519)	101	104	96	NS
	90	94	111	NS
	112	116	81	NS 010
Digestive diseases (520–579)	100	60	101	.010
	1/0	42	110	.010
	100	01	91	
Malo	141	93	04 70	
	150	377	/0	
Infont doothe (760–770)	172	- / /	92 59	020
Malo	155	102	56	.020 NS
	200	01	3 4 7	NG
	104	109	Q1	NS
Malo	3 84	03	115	NS
Female	118	121	74	NS
Violence (F800–F969)	167	78	85	001
Male	169	75	87	010
Female	164	84	82	NS
Accidents (E800–E949)	177	76	81	.010
Male	173	76	84	.050
Female	184	75	77	.100
Suicide (E950–E959)	108	86	108	NS
Male	107	73	120	NS
Female	<sup>3</sup> 109	119	³ 79	NS
Homicide (E960-E969)	256	<sup>3</sup> 72	<sup>3</sup> 48	.010
Male	273	³ 77	³ 21	.001
Female	• • •		••	
Total mortality	129	97	87	.001
Male	123	91	95	.001
Female	135	104	79	.001

<sup>1</sup> Ratio of observed to expected deaths imes 100.

<sup>2</sup> Categories based on the Ninth Revision of the International Classification of Diseases, Clinical Modification (38).

<sup>3</sup> Based on fewer than 5 cases.

## Conclusions

This study was designed to test the hypothesis that, in 1974, mortality rates varied inversely over three socioeconomic areas in Des Moines. If the shortcomings of the data and the method are not viewed as overly problematic, the hypothesis cannot be rejected. That is, socioeconomic differentials in mortality rates continue to exist. However, the failure to find a pattern in which rates for all categories of death vary consistently as a negative and monotonic function of socioeconomic status suggests that several categories of death do not vary strongly with socioeconomic status.

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