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Work stress, sleep deficiency and predicted 10-year cardiometabolic risk in a female patient care worker population

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Abstract

Objectives—The aim of this study was to investigate the longitudinal effect of work-related stress, sleep deficiency and physical activity on 10-year cardiometabolic risk among an all-female worker population.

Methods—Data on patient care workers (n=99) was collected two years apart. Baseline measures included: job stress, physical activity, night work and sleep deficiency. Biomarkers and objective measurements were used to estimate 10-year cardiometabolic risk at follow-up. Significant associations ($P<0.05$) from baseline analyses were used to build a multivariable linear regression model.

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Results—The participants were mostly white nurses with a mean age of 41 years. Adjusted linear regression showed that having sleep maintenance problems, a different occupation than nurse, and/or not exercising at recommended levels at baseline increased the 10-year cardiometabolic risk at follow-up.

Conclusions—In female workers prone to work-related stress and sleep deficiency, maintaining sleep and exercise patterns had a strong impact on modifiable 10-year cardiometabolic risk.

Keywords

Cardiometabolic risk; Nurses; Sleep maintenance; Physical activity; Follow-up; work-family conflict; sleep

Introduction

Cardiovascular diseases are the leading cause of adult death in the world, with approximately 32 % and 27 % of adult female and male fatalities in 2004 attributed to cardiovascular diseases [Mathers, et al. 2008]. In the U.S., recent statistics show that heart diseases and strokes caused 34 % of all deaths and represented a total annual cost of 475.3 billion dollars; making cardiovascular disease the foremost cause of death and the third-most expensive disease in America [Roger, et al. 2011]. The last decade has seen both increased awareness and reduction of the incidence and prevalence of cardiovascular disease, but it remains a major burden [Backe, et al. 2012].

Recent research has identified sleep quality and sleep duration as important factors in cardiovascular disease risk [Buxton and Marcelli 2010, Laugsand, et al. 2013]. Indeed, insufficient sleep duration and poor sleep quality appear to contribute to increased cardiovascular disease risk [Buxton and Marcelli 2010, Cappuccio, et al. 2010], and have been linked to elevated body mass index [Hasler, et al. 2004] [Kohatsu, et al. 2006], weight gain [Patel, et al. 2004, Patel, et al. 2008], obesity [Buxton, et al. 2013, Cizza, et al. 2005, Gangwisch, et al. 2005, Taheri, et al. 2004], and diabetes mellitus [Ayas, et al. 2003, Buxton, et al. 2013, Gottlieb, et al. 2005]. Additionally, modulations of cardiovascular function by sleep duration have been demonstrated showing elevations in blood pressure [Lusardi, et al. 1999, Meier-Ewert, et al. 2004], which is consistent with observed associations of sleep duration with blood pressure [Gottlieb, et al. 2006, Stranges, et al. 2010]. Perhaps most importantly, past research has shown that short sleep (<6 hours per day) predicts premature mortality [Dew, et al. 2003, Grandner, et al. 2010, Kripke, et al. 2002, Patel, et al. 2004, Wingard and Berkman 1983]. Long sleep duration (>8 hours per day) may also be associated with cardiovascular risk [Gallicchio and Kalesan 2009] [Buxton and Marcelli 2010, Hale and Do 2007].

Shift work, particularly night shift work, has been associated with sleep deficiency and has shown a dose-response relationship with coronary heart disease [Kawachi, et al. 1995, Åkerstedt 2003]. A vulnerable population with regards to both sleep deficiency and shift work are patient care workers [Buxton, et al. 2012, Kawachi, et al. 1995]. A recent study demonstrated that 57 % of a large patient care worker cohort reported sleep deficiency, specifically a lack of quality, duration and/or perceived sufficiency of sleep [Buxton, et al.

2012]. The Nurses Health Study showed a 4 % increase in the hazard ratio for ischemic heart disease with every 5 years nurses worked rotating night shifts [Brown, et al. 2009]. Furthermore, a large-scale, prospective study on the same population of married nurses demonstrated a significant increase in mortality risk in those who reported either short or long sleep duration [Patel, et al. 2004]. This increased risk was associated with hypertension [Patel, et al. 2004], and other studies have shown the independent association of short and long sleep duration with coronary events in this cohort [Ayas, et al. 2003].

Patient care workers also report significantly higher psychosocial stress than other occupational groups [Evans and Steptoe 2002]. When compared with other workers, nurses report more job strain due to higher demands and lower skill utilization [Evans and Steptoe 2002]. In addition, patient care workers report high levels of work-family conflict, an important cause of perceived stress [Kim, et al. 2012].

There is a well-established association between work-related stress and risk of cardiovascular disease [Backe, et al. 2012, Eller, et al. 2009, Schnall, et al. 2009], but the effect has been claimed to be significantly less than that of traditional risk factors like amount of physical activity [Kivimäki, et al. 2012, Steptoe and Kivimäki 2012].

The concept of work-related stress is constantly evolving. A recent review highlights the need for a broader understanding of how it relates to cardiovascular risk [Backe, et al. 2012]. Work-family conflict has been identified as a major cause of work-related stress in several studies, and has been implicated as a contributing factor to cardiovascular risk [Bellavia and Frone 2005, Berkman, et al. 2010]. The conflict of roles at work and in one's personal life is recognized as an independent factor from job strain and effort-reward imbalance [Bhui, et al. 2012, Netemeyer, et al. 1996], and might contribute to our understanding of the psychosocial work-environment in cardiovascular risk research [Frone, et al. 2011]. Two recent systematic reviews on cardiovascular risk and workplace stressors also underline the need for more longitudinal studies on female populations [Backe, et al. 2012, Eller, et al. 2009]. The association between high demands, low control and low co-worker support (iso-strain) at work and cardiovascular risk is consistently found in men [Belkic, et al. 2004], but the evidence in women is less conclusive [Backe, et al. 2012]. To address this, we focus the current analysis on a sample of women.

A study on a patient worker population combining several different measures of work stress controlled by other confounding factors such as physical activity [Hublin, et al. 2010], sleep duration [Chandola, et al. 2010] and night-work hours [Yang, et al. 2006].

In the current study we test the hypothesis that work-related stress, sleep deficiency and physical activity over time influences modifiable 10-year cardiometabolic risk in a female worker population. More specifically, we investigate the predictive associations of baseline characteristics on modifiable cardiometabolic risk factors in a female worker population, and investigate the longitudinal effects of iso-strain, work-family conflict, night work, physical activity and sleep deficiency on a non-self reported, modifiable cardiometabolic risk score [Marino, et al. 2014], in a female patient care worker population.

Materials and Methods

Study Design

This study had a longitudinal design where data was collected at two time-points from a group of patient care workers at a large hospital in Boston. Baseline collection spanned from October 2009 to January 2010 and the second was a follow-up from August to November 2011 (Figure 1). All participants provided written informed consent. Research protocols were approved by the designated institutional review board and consistent with the principles in the Declaration of Helsinki.

Study Sample and Data collection

Data collection at baseline has been described in previous articles [Buxton, et al. 2012] [Sorensen, et al. 2011] [Kim, et al. 2012]. In brief, the data presented at baseline was collected through a cross-sectional survey of patient care workers at two major hospitals in 2009. In this initial data collection, we randomly selected 2000 eligible workers and invited them via e-mail to participate in the survey on-line. Of these patient care workers initially contacted, 1572 participants completed at least 50 % of the survey. When collecting follow-up data, only employees from one of the two hospitals were contacted. Employees who completed the initial survey at this hospital (n=840) were randomized and divided into six equal groups. Each group received an initial contact via e-mail, reminding employees of their participation in the original survey and asking about their interest in the second phase of the study. If not responsive after the initial contact, employees received up to 7 additional contacts via email. A small number of subjects who expressed interest but subsequently became unresponsive were re-contacted by phone. Subjects who completed phase two of the study were compensated \$100. Out of 840 employees who were re-contacted after completing the survey, a total of 99 (11.8 %) females completed the phase-two visit and blood draw. A low participation rate at time two was expected as this was a much more comprehensive sampling with in-person collection of both biomarkers and self-report data. Test statistics (frequency distributions and mean differences) showed that these responders were similar to the overall sample on important baseline socio-demographic characteristics, with race/ethnicity (described below) being the only significant difference.

Measures

Registered nurses measured height, weight, and blood pressure, and collected blood samples. Blood pressure was taken using a calibrated, clinical-standard arm cuff (Welch Allyn Spot Vital Signs monitor model# 4200B-E1) and systolic blood pressure (SBP) was used as a continuous measure. Body Mass Index (BMI) was calculated as weight (kg) per meter squared (m²) of height. Subjects were given a *smoking status* as current smoker if they answered yes to the following question “Have you smoked a cigarette, even a puff, in the last 7 days? (Yes; No)”. Venous blood samples provided by all 99 subjects were assayed to determine cholesterol and glycosylated hemoglobin values.

Assay details—All assays at follow-up were performed by CLIA certified laboratories. Glycosylated hemoglobin (Hb A_{1C}) was assayed using a Roche P-Modular Tina-Quant Immunoassay, with an intra-assay precision coefficient of variation (CV) of 0.8–1.5%, an

inter-assay CV of 1.3–2.0%, and a lower limit of detection of 2.9%. Total cholesterol was measured enzymatically in serum using a Roche/Hitachi analyzer with an intra-assay CV of 0.8%, an inter-assay CV of 1.7%, and a lower limit of detection of 3mg/dL. HDL cholesterol was measured enzymatically in serum via a Roche/Hitachi analyzer with an intra-assay CV of 0.60–0.95%, an inter-assay CV of 1.2–1.3%, and a lower limit of detection of 3mg/dL.

Assessment of 10-year cardiometabolic risk—Cardiometabolic 10-year risk was assessed based on five non-self report modifiable cardiometabolic risk factors, initially developed in the Framingham Study [Wilson, et al. 1998] modified by D’Agostino [D’Agostino, et al. 2008], and further developed by Marino et al. in 2013 using the Framingham offspring study to focus solely on modifiable factors [Marino, et al. 2014]. The most recent modification evaluates risk through the addition of glycosylated hemoglobin (Hb A_{1c}) systolic blood pressure and BMI. In brief, while controlling for age and in gender-stratified risk models, the model assessed current smoking, and continuous total cholesterol, HDL-cholesterol, systolic blood pressure, HbA_{1c} levels, and BMI. Details of the model are described elsewhere [Marino, et al. 2014]. Each of the items in the component score was kept continuous to keep them sensitive to small changes and to reflect the potential impact of interventions. Some (n=2) of the participants had missing data on more than one of these risk factors and were excluded from analyses.

Covariates—The covariates were determined *a priori* and reported at baseline. We selected covariates that have been associated with cardiovascular health, sleep and psychosocial stress. All covariates have been described in previous studies [Buxton, et al. 2012, Kim, et al. 2012, Sorensen, et al. 2011].

Socio-demographic factors were obtained through participants reporting their *age* (years), *gender*, *race/ethnicity* (Hispanic, White, Black and mixed race/others), *occupation* (staff nurse, patient care associate (PCA) and others), *education* (GED or less; Some College; College Degree; Graduate School), *ability to pay bills* (great deal of difficulty; some difficulty; a little difficulty; no difficulty; don’t know; refused), height (inches) and weight (pounds). *Body mass index* (BMI) was measured by a nurse at follow-up using weight and height (kilograms per square meter).

Night work was quantified from detailed administrative payroll data and calculated as average night work-hours per month (between 10 PM and 6 AM), calculated from October 2008 until August 2009, making the assessment a year before initiating the survey for all workers. Excluding shifts shorter than 4 hours, the variable was trichotomized into 0–6 hours, 6–72 hours and more than 72 hours per month, over the past year, during months worked, as described previously [Buxton, et al. 2012].

Work-related stress was assessed by self-reported *job demands*, *decision latitude*, *coworker/supervisor support* and *work-family conflict*. A modified version of the Job Content Questionnaire [Karasek Jr 1979, Karasek, et al. 1998, Lusardi, et al. 1999] measured *job demands*, *decision latitude* and *co-worker/supervisor support*. *Job demands* were assessed through 5 items that were weighted and summed yielding a scale from 12 to 48 [Karasek, et al. 1998]. *Decision latitude* was assessed through 9 items and created as a weighted sum of

decision authority and skill discretion from the Job Content Questionnaire [Karasek, et al. 1998]. *Co-worker/supervisor support* was assessed through 5 items (3 and 2 respectively) with 5 response categories summed, yielding a scale from 2 to 10 [Karasek, et al. 1998]. *Supervisor support* was summed and scaled giving a scale of 3 to 15 [Karasek, et al. 1998]. Social support was defined as the sum of co-worker and supervisor support. *Iso-strain* was a composite variable assessed by combining the risk categories job demands, decision latitude and social support and dichotomizing these three variables into low/high categories [Bhui, et al. 2012, Harvey, et al. 2011]. Iso-strain was defined as the combination of high job demands, low control and low social support [Bhui, et al. 2012].

Work-family conflict was assessed by the work-family conflict scale [Netemeyer, et al. 1996] consisting of five items with response categories from 1–5 summed, scaled and giving a range of 5–25: (1) “The demands of my work interfere with my family or personal time”. (2) “The amount of time my job takes up makes it difficult to fulfill family or personal responsibilities”. (3) “Things I want to do at home do not get done because of the demands my job puts on me”. (4) “My job produces strain that makes it difficult to fulfill my family or personal duties”. (5) “Due to work-related duties, I have to make changes to my plans for family or personal activities”. Response categories ranged from 5 = “Strongly Agree” to 1 = “Strongly Disagree”, making a higher score an indicator of more work-family conflict.

Physical Activity (PA) outside of work was assessed with a measure adapted from the Centers for Disease Control and Prevention Behavioral Risk Factor and Surveillance System Physical Activity measure [Centers for Disease Control and Prevention 2009]. Respondents were asked about their participation in vigorous and moderate PA of at least 10 minutes’ duration outside of work. Each activity was defined by number of days per week and the total duration (hours and minutes) per day. *Recommended physical activity* was defined as reporting at least 30 minutes of moderate or vigorous activity on at least 5 days a week or at least 20 minutes of vigorous activity on at least 3 days a week [US Department of Health and Human Services].

Sleep deficiency [Luyster, et al. 2012] in the last 4 weeks was assessed by self-report. Sleep was measured using three self-reported variables: sleep duration, sleep quality and sleep insufficiency. *Short sleep duration* was assessed by asking subjects how many hours of sleep they got each night over the past four weeks [Buxton, et al. 2012, Sorensen, et al. 2011]. Sleep quality was determined based on the frequency of *sleep maintenance problems*, assessed by asking subjects how often they woke up in the middle of the night or early morning, with four response categories: “not during the past month”; “Less than once per week”; “Once or twice per week”; or “three or more times a week” [Buysse, et al. 1989]. *Sleep insufficiency* was determined based on subject response on a five-component scale ranging from “never” to “always” to the question “how often during the past 4 weeks did you get enough sleep to feel rested upon waking up?” [Sorensen, et al. 2011], similar to the Centers for Disease Control measures used in the BRFSS study to assess U.S. national sleep sufficiency by state [McKnight-Eily, et al. 2008]. Sleep outcomes were dichotomized into “Yes/No” based on participant response on the component scales. In brief, a response of <6 hours per night yielded a “Yes” on *short sleep duration*, a response of “three or more times a week” for interrupted sleep yielded a “Yes” on *sleep maintenance problems* and a response

of “always” or “sometimes” yielded a “Yes” on *sleep insufficiency*. *Sleep deficiency* was defined as having responded “Yes” to one or more of these sleep outcomes. The distribution and overlap of the sleep variables are illustrated in figure 2.

Data analysis

In initial analyses, the characteristics of workers were compared on 10-year cardiometabolic risk score (%) treated continuously for analysis purposes. For dichotomous variables we used the independent sample t-test to compare means. For ordinal and continuously measured characteristics we used simple linear regression or a one-way ANOVA. Significant variables from baseline analyses were entered into a multivariable linear regression. As baseline assessments only had self-report measures, no measure of risk change is included in the analyses.

Results

Our participants (n=99) were predominantly white (91 %), female (100 %), nurses (68 %) with a college degree (65 %) and a mean age of 40.8 (SD 11.9, range 21–62) years. Fourteen (13.7 %) of total respondents had iso-strain, 61 (59 %) met recommended levels of physical activity and 64 (63 %) had sleep deficiency. The participants' score on the individual cardiometabolic risk factors demonstrated means and standard deviations of 112.1 mmHg (SD 12.1) on systolic blood pressure, 5.6 % (SD 0.68) on Hb A_{1C}, 26.2 kg/m² (SD 5.7, range 17.6 – 46.1) on BMI and 67.6 mg/dL (SD 16.7) on HDL cholesterol. The participant characteristics measured by cardiometabolic risk (range 1 – 46 %) are presented in table I.

Multivariable associations

Significant variables from baseline analyses were included in a multi-linear model in order to control for covariation. Table II reports the unadjusted regression coefficient (B). The adjusted regression coefficient (β) reported in the following paragraph, represents the relation between X¹ and Y averaged over all values of X², hence showing the most influential predictor of the significant covariates. Sleep maintenance problems at baseline significantly predicted cardiometabolic risk score at follow-up and this association remained after adjusting for covariates ($\beta=0.26$). In addition, whether or not the patient care worker reported doing the recommended level of physical activity at baseline, predicted increased cardiometabolic risk at follow-up in the adjusted model ($\beta= 0.29$), as did the occupational category the worker belonged to ($\beta= 0.24$). The overall model fit was ($X^2_{red}=0.23$) and a Durbin-Watson test yielded $d=1.59$, which is within the critical value standards for non-autocorrelation, meaning the unobserved error terms are not correlated (see table II for details of model results).

Discussion

In this longitudinal study of 99 female patient care workers, we investigated how baseline measures of sleep deficiency, physical activity, work stress and night work predicted 10-year cardiometabolic risk from biomarkers at 2-year follow-up. Sleep maintenance problems, occupational category, and not meeting the CDC's recommended amount of

physical activity at baseline, significantly predicted increased 10-year cardiometabolic risk. These predictive associations remained when controlled for covariates.

Our finding that sleep maintenance problems independently predict increased cardiometabolic 10-year risk is consistent with a recent large-scale prospective study on reduced sleep quality and incidence of heart failure which identified a dose-response relationship between cumulative increase in insomnia symptoms and risk of heart failure [Laugsand, et al. 2013]. Earlier studies on insomnia symptoms and cardiovascular risk further support a link to heart failure as well as cardiovascular disease [Chien, et al. 2010, Newman, et al. 2000]. A recent study of insomnia and heart failure [Laugsand, et al. 2013] describes an increased relative risk of heart failure for women compared to men, which was attributed to either a higher prevalence of insomnia symptoms in females [Ohayon 2002] or the lower baseline risk for heart failure in women [Laugsand, et al. 2013]. Our results support further investigation into whether insomnia symptoms pose a greater threat to women than men when assessing cardiometabolic risk.

Short sleep duration has previously been linked to increased blood pressure [Buxton and Marcelli 2010, Gottlieb, et al. 2006]. Recent findings from a large American cohort also demonstrated that insomnia symptoms led to a 5-fold elevated risk for hypertension when combined with short sleep duration (< 5 hours), whereas insomnia symptoms and adequate sleep duration did not show significant associations with hypertension [Vgontzas, et al. 2009]. Our results support a predictive link between sleep maintenance problems and cardiometabolic risk, but we did not find evidence for such a “combination effect” when investigating sleep deficiency. However, this could be due to our small sample or all female population; future studies should include combination/composite variables to further the understanding of this link.

Our results failed to show a link between night/shift work hours per month and 10-year cardiometabolic risk, when controlling for covariates. Rotating night shifts has previously been associated with increased risk for ischemic heart disease [Brown, et al. 2009] and exhibits a dose-response relationship with coronary heart disease in the Nurses Health Study [Brown, et al. 2009, Kawachi, et al. 1995]. However, there are other studies showing no association between night shifts and ischemic heart disease [Frost, et al. 2009]. A recent review on night work and ischemic heart disease looked at 16 studies with a relative risk varying from 0.64 to 2.25 [Frost, et al. 2009]. However, overlap between sleep maintenance problems and working night shifts could perhaps be expected as sleep deficiency has been suggested as a mechanism explaining earlier findings [Brown, et al. 2009]. In the current study, there was no significant correlation between working night shifts and reporting sleep maintenance problems. Our results are in line with a recent review showing only a marginally increased relative risk when obtaining night work hours from company records, compared to when night work is self-reported [Frost, et al. 2009].

Failing to exercise moderately or vigorously for 150 minutes or more during a week is a well-established risk factor in cardiovascular disease [Sattelmair, et al. 2011, Wen, et al. 2011]. The results from our study are in accordance with numerous prospective cohort studies and meta-analyses done in several countries [Sattelmair, et al. 2011, Wen, et al.

2011]. Our occupational group, patient care workers, have been investigated through the Nurses Health Study, demonstrating that even light physical activity was associated with lower cardiovascular risk [Lee, et al. 2001].

Interestingly, several studies have documented a link between short sleep duration and reduced physical activity [Briones, et al. 1996, Weaver, et al. 1997]. A recent laboratory study demonstrated how induced short sleep duration (5.5) in 18 healthy subjects led to a 31 % decrease in daily activity and 24 % reduction in moderate and vigorous activity [Bromley, et al. 2012]. When looking at this in our population, though not significant, we observed a trend that a lower percentage of the participants' reporting short sleep duration met the recommended amounts of physical activity than participants reporting adequate sleep duration. Such a link between sleep and physical activity underlines the need for a specific focus on sleep adequacy, duration and maintenance in occupational health interventions.

No measures of job stress were significant in the adjusted analyses in our study. A review on psychosocial stress and ischemic heart disease risk highlights the contradictory evidence in female populations when it comes to job stress [Eller, et al. 2009].

This study has limitations, especially that the sample size is small and was not chosen randomly, which limits the generalizability of our results. However, an all female population is not common in cardiometabolic risk studies. These results should be considered a contribution to a growing research base on female cardiometabolic risk. Another limitation is that we have only two time points approximately two years apart that does not allow for investigations of mediating mechanisms or directionality. But, the extensive questionnaires at baseline, administrative payroll data on night work status, and lengthy follow-up with biomarkers still make a strong case for the relevance of variables showing significant predictive value. Another limitation of note is that our prospective design does not preclude reverse causality. We do not have information on events before the beginning of the study and cannot account for other factors influencing causality. Moreover, we do not have a measure or question addressing postmenopausal status. The average menopausal age in American women is 51 years, and the range is 40–61 years, and about half of our sample is within this range, which is an important risk factor.

The relationship between sleep maintenance problems, physical activity and cardiometabolic risk is a widely studied connection. Yet, there is less evidence in female patient care workers, a population with very high stress levels and high prevalence of sleep deficiency. The Nurses Health Study is perhaps the largest cohort with the most extensive follow-up done on a nurse population. However, its participants consist only of married nurses and do not include patient care associates or single nurses, who are vulnerable workers with regards to both social support and workload. This study also implements biomarker data in a high priority population providing objective measurements to motivate change in occupational health practices. In addition, our data includes non self-report variables such as administrative payroll data on night work status, and measured cardiometabolic risk [Marino, et al. 2014], a much more sensitive outcome than cardiovascular disease or death as registered by death certificates, self-report or hospitalization due to modest accuracy of

diagnoses and referral biases [Frost, et al. 2009]. Sleep has previously been suggested as a third pillar of health, alongside nutrition and exercise, in cardiometabolic disease risk [Buxton, et al. 2013]. Our results support this view and highlight the need for sleep deficiency to be prioritized in occupational health programs, especially those targeting cardiovascular health in women.

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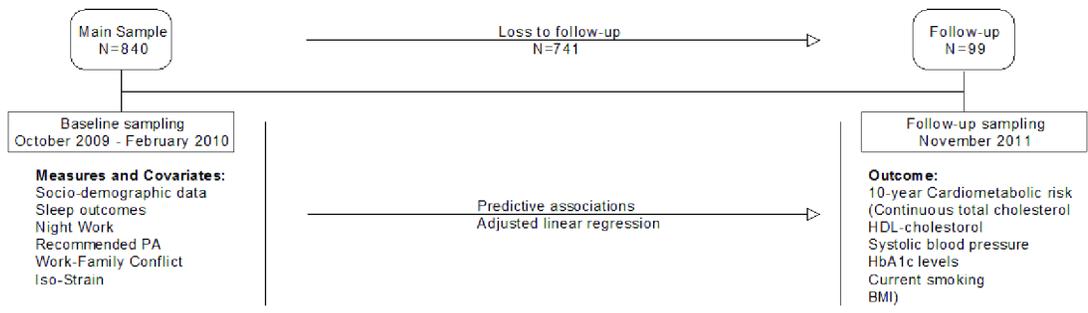


Figure 1.

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Sleep Deficiency, 1 of any component; n=61, 62%

N = 99 Missing = 1

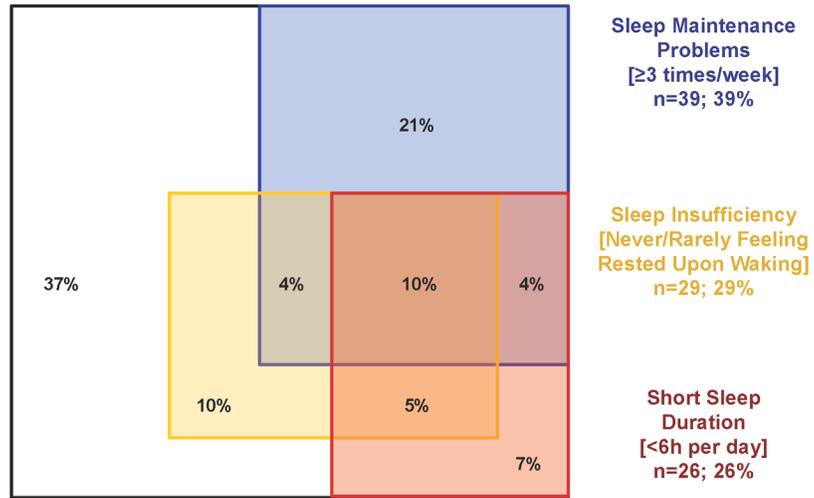


Figure 2.

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Table 1

Mean (SD), correlation (r), variance (r^2/F) of participant characteristics on 10-year cardiometabolic risk score within covariates. Covariates are either dichotomous (yes; no) or ordinal/continuous (range). Categorical variables are shown with both test statistic and mean (SD). Reference categories are bolded. All characteristics may not add up to n=99, because of missing data.

Work Conditions/Demographics	Cardiometabolic risk score (Range 0–100 %)		P-Value
Independent Variables			
Sleep Deficiency	Yes (n=63)	No (n=36)	
	9.29 % (10.5 %)	6.03 % (6.2 %)	0.058
Sleep Duration	< 6 hours (n=26)	>= 6 hours (n=73)	
	9.5 % (9.0 %)	7.54 % (9.3 %)	0.36
Insufficient Sleep	Yes (n=29)	No (n=69)	
	6.13 % (8.8 %)	8.86 % (9.3 %)	1.000
Sleep Maintenance Problems	Yes (n=39)	No (n=59)	
	10.51 % (10.6 %)	6.5 % (7.8 %)	0.048
Recommended Physical Activity	Yes (n=61)	No (n=38)	
	5.6 % (6.0 %)	11.9 % (11.7 %)	0.003
Iso-strain	Yes (n=14)	No (n=74)	
	8.82 % (9.2 %)	8.36 % (9.7 %)	0.87
Job Demands (12–48)	$r=0.09$ $r^2=0.008$		0.37
Decision Latitude (2–10)	$r=0.36$ $r^2=0.001$		0.73
Coworker Support (2–10)	$r=0.14$ $r^2=0.02$		0.17
Supervisor Support (3–15)	$r=0.09$ $r^2=0.009$		0.36
Work-Family Conflict (5–25)	$r=0.04$ $r^2=0.002$		0.68
Occupation (Staff Nurse ; Patient Care Associate (PCA); Other)	(Welch's F) 28.2 (2, 50.2)		<0.001
Staff Nurse	6.5 % (7 %)		
PCA	1.2 % (0.6 %)		
Other	13 % (12.3 %)		
Race\Ethnicity (White ; Hispanic; Black)	$F(2, 92)$ 0.14		0.87
White	8.2 % (9.2 %)		
Hispanic	5.4 % (5.1 %)		
Black	8.1 % (12.9 %)		
Education (GED or Less ; Some College; Degree; Graduate School)	(Welch's F) 1.3 (3, 4.6)		0.38
GED or Less	17.4 % (20.5 %)		
Some College	7.8 % (9 %)		
Degree	6.6 % (8.1 %)		
Graduate School	13 % (11.0 %)		
Economic Status (Difficulty paying bills?) (Little difficulty w/ bills ; ... At least some difficulty w/ bills)	$F(4, 91)$ 0.6		0.7
A Great Deal of Difficulty w/ Bills	5.2 % (3.6 %)		
At Least Some Difficulty w/ Bills	8.4 % (8.8 %)		
A Little Difficulty	10 % (10.2 %)		

Work Conditions/Demographics	Cardiometabolic risk score (Range 0–100 %)	P-Value
No Difficulty	7.6 % (9.3 %)	
Refused to Answer	2.9 % (2.9 %)	
Night work (0–6 hrs monthly ; >6 hrs but < 72; 72 +)	(Welch's F) 6.1 (2, 25.6)	0.007
0–6 hrs monthly	8.8 % (9.6 %)	
>6 hrs but < 72	3.9 % (4.8 %)	
72 +	10.6 % (9.7 %)	

Significant P-values (<0.05) are bolded

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Table II

Multivariable linear regression showing associations of baseline characteristics on 10-year cardiometabolic risk at follow-up. Categories are listed in parentheses.

Independent Variables	Cardiometabolic risk model	
	Unstandardized Coefficient (95 % CI)	P-value
Sleep Maintenance Problems (yes; no)	0.046 (0.01, 0.048)	.01
Occupation (Staff nurse ; PCA; Other)	0.027 (0.006, 0.48)	.01
Recommended Physical Activity (yes; no)	0.054 (0.019, 0.09)	.003
Night work Categories (<=6, 6-72, 72+)	0.015 (-0.012, 0.041)	.27

Significant *P*-values are bolded. The interpretation of the coefficient is the expected change in *y* for a one-unit change in *x* when the other covariates are held fixed—that is, the expected value of the partial derivative in *y* for a one unit change in *x* when other covariates are held fixed. This is sometimes called the unique effect of *x* on *y*.

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