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Model-based Analysis of Tuberculosis Genotype Clusters in the United States Reveals High Degree of Heterogeneity in Transmission and State-level Differences Across California, Florida, New York, and Texas

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Abstract

Background.—Reductions in tuberculosis (TB) transmission have been instrumental in lowering TB incidence in the United States. Sustaining and augmenting these reductions are key public health priorities.

Methods.—We fit mechanistic transmission models to distributions of genotype clusters of TB cases reported to the Centers for Disease Control and Prevention during 2012–2016 in the United States and separately in California, Florida, New York, and Texas. We estimated the mean number of secondary cases generated per infectious case (R_0) and individual-level heterogeneity in R_0 at state and national levels and assessed how different definitions of clustering affected these estimates.

Results.—In clusters of genotypically linked TB cases that occurred within a state over a 5-year period (reference scenario), the estimated R_0 was 0.29 (95% confidence interval [CI], .28–.31) in the United States. Transmission was highly heterogeneous; 0.24% of simulated cases with individual R_0 >10 generated 19% of all recent secondary transmissions. R_0 estimate was 0.16

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Supplementary Data

Supplementary materials are available at *Clinical Infectious Diseases* online. Consisting of data provided by the authors to benefit the reader, the posted materials are not copyedited and are the sole responsibility of the authors, so questions or comments should be addressed to the corresponding author.

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(95% CI, .15–.17) when a cluster was defined as cases occurring within the same county over a 3-year period. Transmission varied across states: estimated R_0s were 0.34 (95% CI, .3–.4) in California, 0.28 (95% CI, .24–.36) in Florida, 0.19 (95% CI, .15–.27) in New York, and 0.38 (95% CI, .33–.46) in Texas.

Conclusions.—TB transmission in the United States is characterized by pronounced heterogeneity at the individual and state levels. Improving detection of transmission clusters through incorporation of whole-genome sequencing and identifying the drivers of this heterogeneity will be essential to reducing TB transmission.

Keywords

tuberculosis; transmission heterogeneity; United States; clustering

Tuberculosis (TB) incidence in the United States fell by more than 70% between 1993 and 2017; reductions in transmission driven by progress in detecting and treating latent TB infection among persons recently exposed have been a key component of this decline [1, 2]. Even though a minority of new TB cases are due to recent transmission [3, 4], extensive public health resources are required for their investigation and control. In the absence of timely TB control measures, these recent transmission outbreaks can grow, leading to large numbers of cases within the local community [5]. This is especially pertinent for vulnerable populations, which include racial and ethnic minorities, persons living in congregate settings (including correctional facilities and homeless shelters), and patients with medical comorbidities who are susceptible to TB and poor TB outcomes [6–9]. The risk of outbreaks expanding into larger populations and becoming endemic increases as they become larger or more frequent [10]. Understanding more about past outbreaks and responding with focused strategies can help prevent future outbreaks and mitigate these negative impacts.

Transmission of *Mycobacterium tuberculosis* (*Mtb*) is heterogeneous, driven by pathogen, host, environmental, and societal factors [11]. A better understanding of this heterogeneity can help improve TB control efforts, including outbreak prevention and response. Molecular characterization of *Mtb* isolates through genotyping, which consists of matching isolates on the basis of both spacer oligonucleotide typing (spoligotype) and 24-locus mycobacterial interspersed repetitive unit-variable number of tandem repeats (MIRU-VNTR), can identify TB cases that are clustered and thus presumptively related by recent transmission [12]. Genotyping has also been used to understand the evolution and spread of *Mtb* [13, 14]. We used the US distribution of genotypically clustered TB cases to characterize recent transmission nationally and in the 4 states reporting more than half of all TB cases: California, Florida, New York, and Texas.

METHODS

Cluster Distribution Data

We used data from the US National Tuberculosis Surveillance System and the National Tuberculosis Genotyping Service for TB cases reported from the 50 US states and District of Columbia during 2012–2016 to infer the distribution of TB clusters in the United States and

independently in California, Florida, New York, and Texas. Cases were defined as clustered if they had matching spacer oligonucleotide typing (spoligotype) and 24-locus MIRU-VNTR genotyping results, were reported within specified geographic boundaries (ie, the same county or state), and occurred during 2 time periods (ie, 2012–2016 or 2014–2016). A cluster definition that included cases reported within a single state boundary and within 2012–2016 was defined as the reference scenario.

Branching Process Model

We used a branching process framework to describe recent transmission and cluster formation [15, 16]. In this framework, the number of secondary cases resulting from a single case over a course of its infection occurs probabilistically and is given by the "offspring distribution" of the branching process model, Z (Table 1). The mean of this distribution is R_0, the reproductive number, equal to the average number of secondary cases resulting from a single case. We assume that this probability distribution of secondary cases follows a Poisson distribution with parameter ν, that is, $\boldsymbol{P}(Z = z) \sim \text{Poisson}(\nu)$.

Modeling Individual-level Heterogeneity

To incorporate individual-level heterogeneity in transmission, we varied the value of ν between individuals; here, ν can also be interpreted as the individual-level reproductive number [16]. We compared 4 models, with each encapsulating a different level of heterogeneity represented by the distribution of ν (Table 2). In all 4 models, the mean of both ν and the off-spring distribution Z is R_0. Model I ("homogeneous" model) assumes that ν is constant ($\nu = R_0$) so that all individuals have the same infectious potential and the number of secondary cases resulting from each case is Poisson-distributed. Model II ("Susceptible-Infectious-Recovered-type" model) assumes that the individual reproductive number is distributed exponentially, similar to assumptions in standard to assumptions in standard Susceptible-Infectious-Recovered (SIR) type compartmental models. Model III ("overdispersed" model), a model previously used to capture heterogeneity in transmission of TB and other infectious diseases [14, 16–18], assumes a gamma distributed ν, resulting in a negative binomial distribution of secondary cases, with mean R_0 and shape parameter k. Finally, model IV ("long-tailed" model) allows for even greater heterogeneity using a Poisson lognormal distribution; this modeling approach is often used to describe species abundance [19, 20] and, recently, heterogeneity in TB transmission [21]. This model assumes that the individual reproductive numbers ν follow a lognormal distribution (ie, $\ln(\nu) \sim N(\mu, \sigma^2)$), where μ and σ are the mean and standard deviation, respectively, of an underlying normal model. Larger values of σ^2 are indicative of increased heterogeneity.

Model Comparison

We used a likelihood-based framework to evaluate and compare the fit of each model described above to the observed data. Using the likelihood function (described in detail in the Supplementary Materials, Section 4), we calculated the maximum likelihood estimates (MLEs), the parameters that yield the highest likelihood, and corresponding 95% confidence regions/intervals. We compared models using the Akaike information criterion,

$AIC = 2r - 2\ln(\widehat{\mathscr{L}})$, where r is the number of parameters in a model and $\widehat{\mathscr{L}}$ is the likelihood estimate.

Sensitivity of Model Inference to Data Censoring and Importation

To assess the sensitivity of model inference to possible imperfections in data, we conducted a simulation study in which we considered 2 mechanisms by which observed data could differ from true clustering. First, we assumed underreporting of clustered TB cases due to factors such as cases not being reported in local jurisdictions, cases not being culture-confirmed (eg, pediatric cases) or isolates not being genotyped, or cases being right- or left-censored over time. Second, we assumed overascertainment of clusters due to inclusion of imported cases of matching genotype (ie, not due to local or recent transmission). We generated synthetic cluster distributions by simulating the branching process models under various assumptions about R_0 and individual-level heterogeneity (taken as true parameter values), which also incorporated imperfections in data described above. For each synthetic cluster distribution, we then applied the likelihood-based inference method to estimate both R_0 and individual-level heterogeneity (estimated parameter values). By comparing true parameter values to their corresponding estimates, we inferred the sensitivity of each estimated parameter value to underreporting or overascertainment. (See Supplementary Materials, Sections 5 and 6, and Supplementary Figures 1–4 for additional details.)

RESULTS

Cluster Distributions

Of 35 313 genotyped TB cases reported during 2012–2016 in the United States, 13 159 (37%) were clustered under the reference definition of having the same genotyping result as at least 1 other case that was reported from the same state during the same period (Figure 1A). The remaining 22 154 (63%) unclustered cases did not have another same-state case with the same genotype during that time frame. Among clustered cases, 31% occurred in large (≥10 cases) clusters; the largest cluster included 148 cases. When we restricted the definition of clustering to cases reported within county boundaries and occurring within a 3-year period (2014–2016), 21% of cases were clustered, and 15% of clustered cases were in clusters of ≥10 cases. The largest cluster contained 65 cases (Figure 1B).

Model Comparison

Of the 4 models considered, model IV (long-tailed model), which assumed the highest level of individual-level heterogeneity, provided the best statistical fit. The MLE under model IV was statistically >1000 times more likely to explain the data than models I–III (Table 2). Much of this improved fit reflected a better ability to represent clusters of large size (ie, the long tail), which occurred more frequently than predicted under models I–III (Figure 2). This result was found consistently, regardless of cluster period or if clusters were defined within counties or states (see Supplementary Figures 5–8). Comparison of model fits across 4 states (Supplementary Materials, Section 9, and Supplementary Figures 11 and 12) also support this conclusion.

Individual-level Heterogeneity in Transmission

The estimated distribution of the individual reproductive number revealed substantial individual-level heterogeneity (Figure 3). For 95% of TB cases, the individual reproductive number was estimated to be less than 1; these cases generated only 38% of secondary cases. Only 5% of TB cases were estimated to have individual reproductive number of 1 or greater, and these individuals generated 62% of secondary cases. Of note, 0.24% of cases were estimated to have a reproductive number larger than 10 and were projected to generate 19% of secondary cases.

Cluster Ascertainment Criteria and Inference

The estimated value of R_0 ($\widehat{R_0}$) varied by geographic area and the time frame used to define clustering. For example, $\widehat{R_0}$ in the reference scenario (state boundaries, and 5-year time frame between 2012 and 2016) was 0.29 (95% confidence interval [CI], .28–.31; Figure 4, hatched orange region); this estimate fell to 0.25 (95% CI, .24–.27; Figure 4, orange region) when a 3-year time frame between 2014 and 2016 was used, 0.19 (95% CI, .18–.2; Figure 4, hatched green region) when county boundaries but a 5-year window were used, and 0.16 (95% CI, .15–.17; a 45% reduction; Figure 4, green region) when both 3-year time frame and county boundaries were used (see Supplementary Table 1). Estimates of heterogeneity in transmission were less sensitive to the choice of cluster definition. For example, the estimated percentage of secondary cases originating from cases with R_0 >10 fell between 16% and 19% regardless of cluster definition (Supplementary Figure 9).

State-level Variation

The cluster distribution of TB cases and the corresponding estimates of individual-level reproductive numbers varied considerably at the state level. For example, the proportion of clusters with ≥10 cases was nearly 8-fold larger in Texas compared with New York (Figure 5, blue line compared with red line). Consequently, the estimated mean individual-level reproductive numbers varied by a factor of 2: 0.19 (95% CI, .15–.27) in New York, 0.28 (95% CI, .24–.36) in Florida, 0.34 (95% CI, .3–.4) in California, and 0.38 (95% CI, .33–.46) in Texas (see Supplementary Table 2). There were substantial differences in estimated degree of transmission. For example, the contribution of individuals with R_0 >10 to the total secondary cases varied from 9.5% (from 0.13% of individuals) in Florida to 20% (from 0.3% individuals) in California (Supplementary Figure 10).

Sensitivity of Model Inference to Data Censoring and Importation

Under- and overascertainment of clusters had a predictable effect on the inference of R_0. R_0 was underestimated (bottom left quadrant in Supplementary Figure 3A) when cases were underreported and overestimated (top right quadrant in Supplementary Figure 3A) when cases were overascertained. In both instances, the degree of under- or overestimation was linearly associated with the level of underreporting or overascertainment. Estimates of individual-level heterogeneity in transmission (σ) were unaffected by underreporting and were only slightly underestimated when the observed clusters included overascertainment of imported cases (Supplementary Figure 3B).

DISCUSSION

This model-based analysis of genotype-clustered TB cases in the United States revealed that there is substantial heterogeneity in transmission. We estimated that 95% of individual cases transmit to less than 1 secondary case each and contribute to only 38% of overall secondary transmission. By contrast, 0.24% of cases were estimated to transmit to 10 or more secondary cases, resulting in 19% of all secondary cases. This degree of heterogeneity is larger than described with prior models (ie, negative binomial distribution) [16, 22] but is remarkably consistent with data and prior analysis from the United Kingdom and the Netherlands [21] (Supplementary Materials, Section 10, and Supplementary Figures 13 and 14). Taken together, these results suggest that heterogeneity in TB transmission (in low-burden settings) may be larger than previously thought and may be driven by common underlying processes (eg, propensity for transmission in outbreak-prone settings).

The characteristics of *Mtb* transmission varied across states. For example, even though TB incidence is similar in Texas and New York, the estimated R_0 in Texas was twice as high, suggesting that more cases in Texas reflect recent transmission, whereas more cases in New York may represent reactivation of latent infection or importation. These findings are consistent with estimates of recent transmission from the Centers for Disease Control and Prevention [23] and from a previously published transmission model [24]. This may be reflecting differences in demography, immigration patterns, mobility, population density, relative sizes of key populations (eg, people experiencing homelessness or incarceration), variation in *Mtb* strain pathogenicity and virulence, and societal context [25]. Also important to consider are differences in the size of jurisdictions, particularly counties, which can vary considerably between states. Outbreaks are unlikely to be confined to small and/or highly interconnected counties (eg, in cities) but may be contained in larger counties with more self-contained populations.

Conventional genotyping has known limitations that can lead to underestimation or overestimation of clustering. Underestimation may occur if true transmission-linked cases are not detected (eg, individuals move out of a jurisdiction or are reported elsewhere), do not have a specimen culture showing *Mtb*, or do not have an *Mtb* specimen that was genotyped (eg, technical challenges associated with particular loci [26]). In the United States, most cases are likely diagnosed, and more than 95% of cultured cases have genotyped specimens [27], making the choice of geographic area and period for the definition of clustering important. Our estimates suggest that defining clusters at the county level rather than the state level could reduce estimates of transmission within clusters by more than one-third. Some left and right censoring in observed clusters may also occur over time. For example, cases may be missed toward the beginning of an outbreak (ie, not included in the data) or toward the end (ie, not yet diagnosed). Such under- or overascertainment might cause estimates of recent transmission (ie, R_0) to fall or rise proportionally. The choice of time periods that we examined seemed to have less impact on estimates of transmission. None of the mechanisms mentioned above substantially affected estimates of individual-level heterogeneity in transmission, which remained high in all of our sensitivity analyses. When choosing the appropriate administrative level at which to define clusters, it is important to additionally consider the geographic size and population of administrative units, their

interconnectedness, the relative value of a sensitive vs specific definition, and the level at which any response could be organized.

Conventional genotyping methods may also overestimate clustering by falsely attributing transmission links to cases that share common ancestry but are not related by recent transmission. TB often has a decades-long latency period, genotyping cannot be performed during this latent period, and molecular changes occur slowly; these factors can limit the use of conventional genotyping to estimate transmission. For example, cases that result from a commonly circulating (endemic) strain might reactivate at similar times and thus could share a genotype but not reflect recent transmission events. Genotyping clusters defined by 24-loci MIRU-VNTR could encompass transmission events up to 3 decades in the past [28]. Studies that have compared genotyping with whole-genome sequencing (WGS), considered a gold standard, have found that only about 50%–60% of genotyped clusters could be confirmed by WGS [29, 30]. This discrepancy tended to be much more pronounced for clusters that consisted of nonnative individuals, suggesting that importation (ie, transmission that occurred in distant past or outside of the country) can also contribute to overestimation [26, 29]. Furthermore, accurate clustering by conventional genotyping varies by *Mtb* strain, with some genotypes and lineages having higher rates of internal diversity by WGS, meaning they are less likely to be related by recent transmission [31, 32]. These genotypes are also often concentrated in specific geographies, so the amount of inaccurate clustering can vary by state. As a result, varying imprecision in clustering could account for some of the variation in our estimates of heterogeneity of transmission between states.

Recent and local transmission can be corroborated through identification of epidemiologic links [33]. However, conducting epidemiologic investigations is challenging in populations with large numbers of clustered cases, especially if transmission occurs in settings and venues where identifying contacts may be unfeasible (eg, populations in which routine investigations are challenging, such as those experiencing homelessness and substance abuse) [34, 35]. Novel methods that incorporate transmission and epidemiological dynamics, contact patterns and network structures, genetic diversity, and evolutionary dynamics of *Mtb* may improve our ability to infer transmission by integrating molecular and genomic data [18, 36–41].

The high degree of heterogeneity in the individual reproductive number estimated here might not only reflect individual-level factors but environmental conditions and societal and healthcare provider–related factors that individuals experience. Communities or populations in which background *Mtb* infection rates are higher, comorbidities and risk factors are more prevalent, living conditions are more crowded and less ventilated (eg, homeless shelters, correctional facilities), or more barriers to accessing healthcare exist are likely to experience higher rates of transmission. They are also likely to encounter challenges in the detection and treatment of latent TB infection among persons recently exposed to TB, resulting in larger and more frequent TB clusters [42, 43]. Better characterization of larger clusters, for example, by incorporating data on settings of possible transmission, demography, health equity measures, geography, and other risk factors [44], might help better identify factors that drive high transmission. In our analysis, we did not account for variability in infectiousness among clustered cases (eg, sputum smear positivity and grade, radiographic

evidence of cavitary disease among cases with pulmonary TB disease). Linkage of such clinical data with WGS and phylogenetic analyses of TB cases in large clusters might provide insight into the frequency and mechanism of high transmission, thus improving resource allocation by excluding probably unrelated cases from outbreak investigations. These insights could also help TB programs proactively identify (or potentially predict) the cases and contacts at highest risk of resulting in secondary cases and high transmission events. In this way, further transmission can be prevented through intensified, focused interventions to ensure complete identification, evaluation, and treatment of recently infected contacts. Further work to determine if some strains of *Mtb* are associated with increased transmission due to pathogenicity or virulence factors could also help TB programs prioritize certain cases and contacts for follow-up to prevent further transmission.

In conclusion, this model-based analysis of molecular surveillance data in the United States suggests that, although the overall rate of recent TB transmission is generally low, a small fraction of TB cases probably plays an important role in driving transmission at the population level. Understanding the drivers of this heterogeneity, by identifying populations, settings, and activities that are more frequently associated with large outbreaks, could improve outbreak prevention and response (through early and accurate detection of large clusters), reduce TB transmission, and improve TB-related resource allocation in the United States and more broadly.

Supplementary Material

Refer to Web version on PubMed Central for supplementary material.

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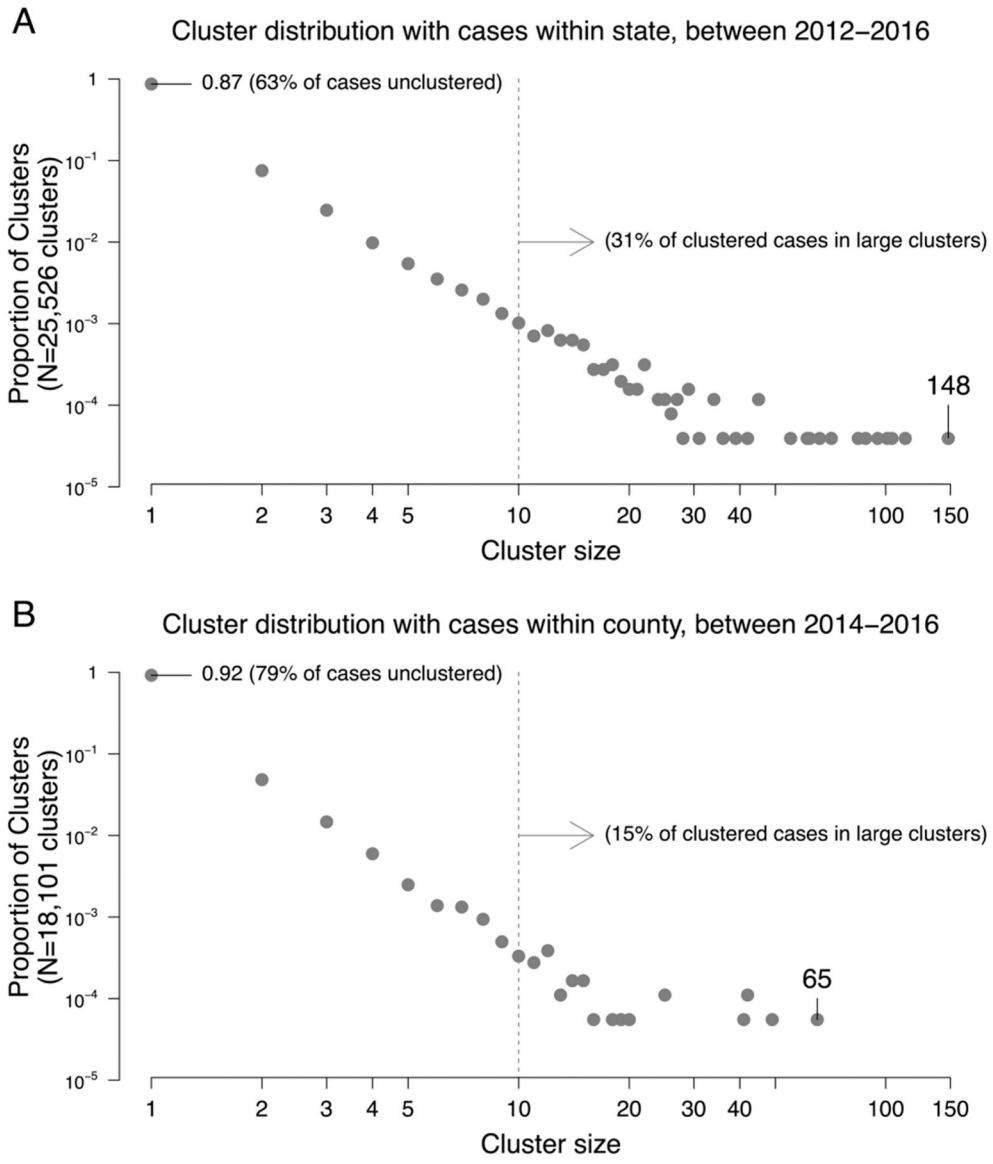

Figure 1.

Genotype cluster distribution of tuberculosis (TB) cases in the United States. Shown are the frequency of observed genotype TB clusters of various sizes in the United States based on cases reported within a given state and occurring within a 5-year time period (2012 to 2016) *(A)* and cases reported within a given county and occurring within a 3-year time period (2014 to 2016) *(B)*. Genotypic clusters are defined as cases with matching spoligotype and 24-locus mycobacterial interspersed repetitive unit-variable number of tandem repeats occurring within the specified geographic boundary and during the specified time window. Both axes are plotted on a log scale.

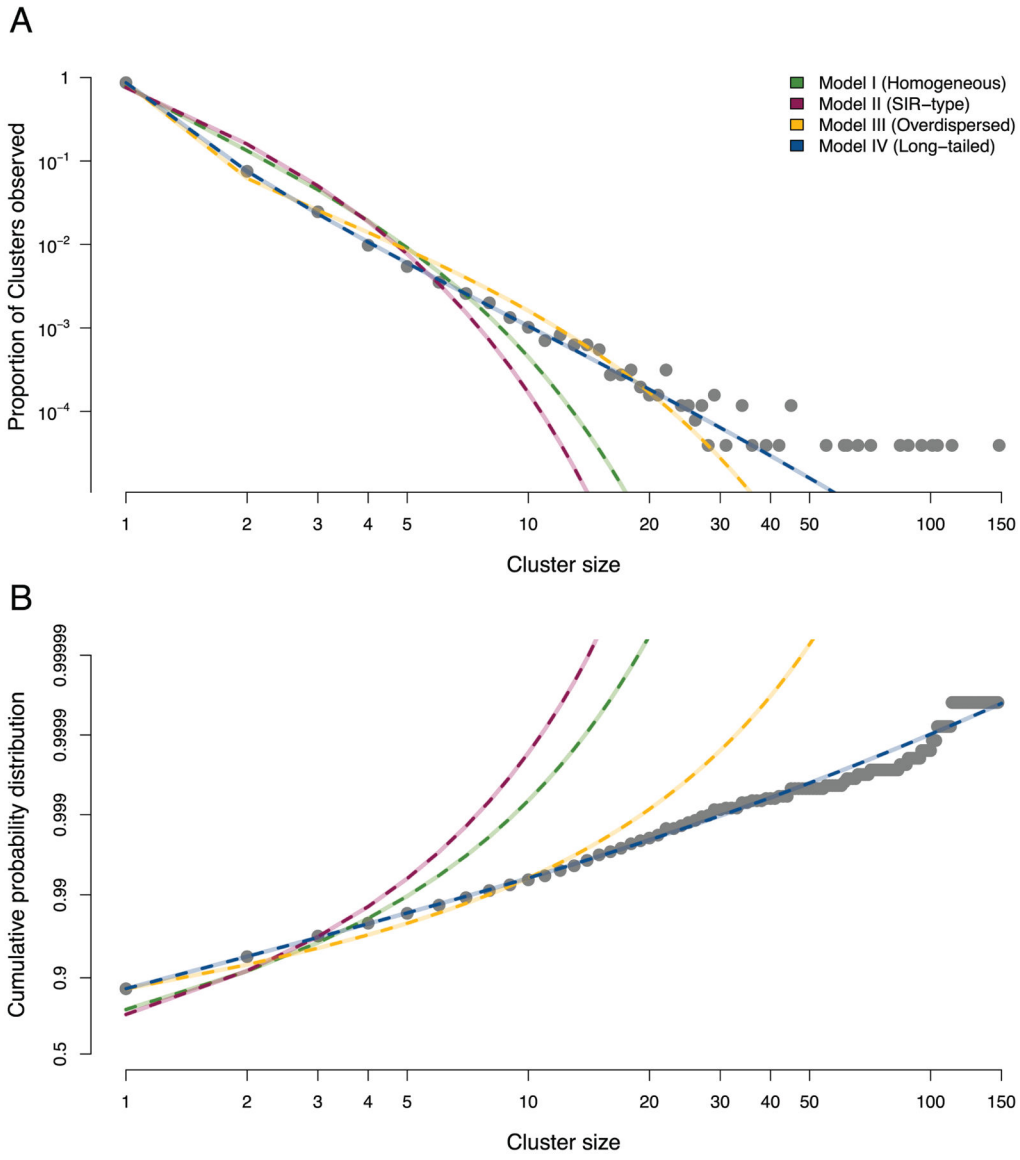

Figure 2.

Fitting branching process models to genotype cluster distributions of tuberculosis (TB) in the United States. We fit branching process models to the cluster distribution consisting of genotyped TB cases occurring within US state boundaries over a 5-year time period (shown in Figure 1A). We considered 4 model assumptions to describe the underlying individual-level heterogeneity. Model I (homogeneous) hypothesized that there was no individual-level heterogeneity except by stochastic chance alone (green); model II (SIR-type) hypothesized that the individual reproductive number followed an exponential distribution (purple); model III (overdispersed), a gamma distribution (yellow); and model IV (long-tailed), a lognormal distribution (blue). Shown are frequency distributions *(A)* and cumulative probability distributions *(B)* corresponding to the best-fit models of each type (shown by colored dashed lines) against the data (shown in gray dots). Cluster size and frequency distributions are

plotted on a log scale, and the cumulative probability distribution is plotted on a logit scale. Abbreviation: SIR, Susceptible-Infectious-Recovered.

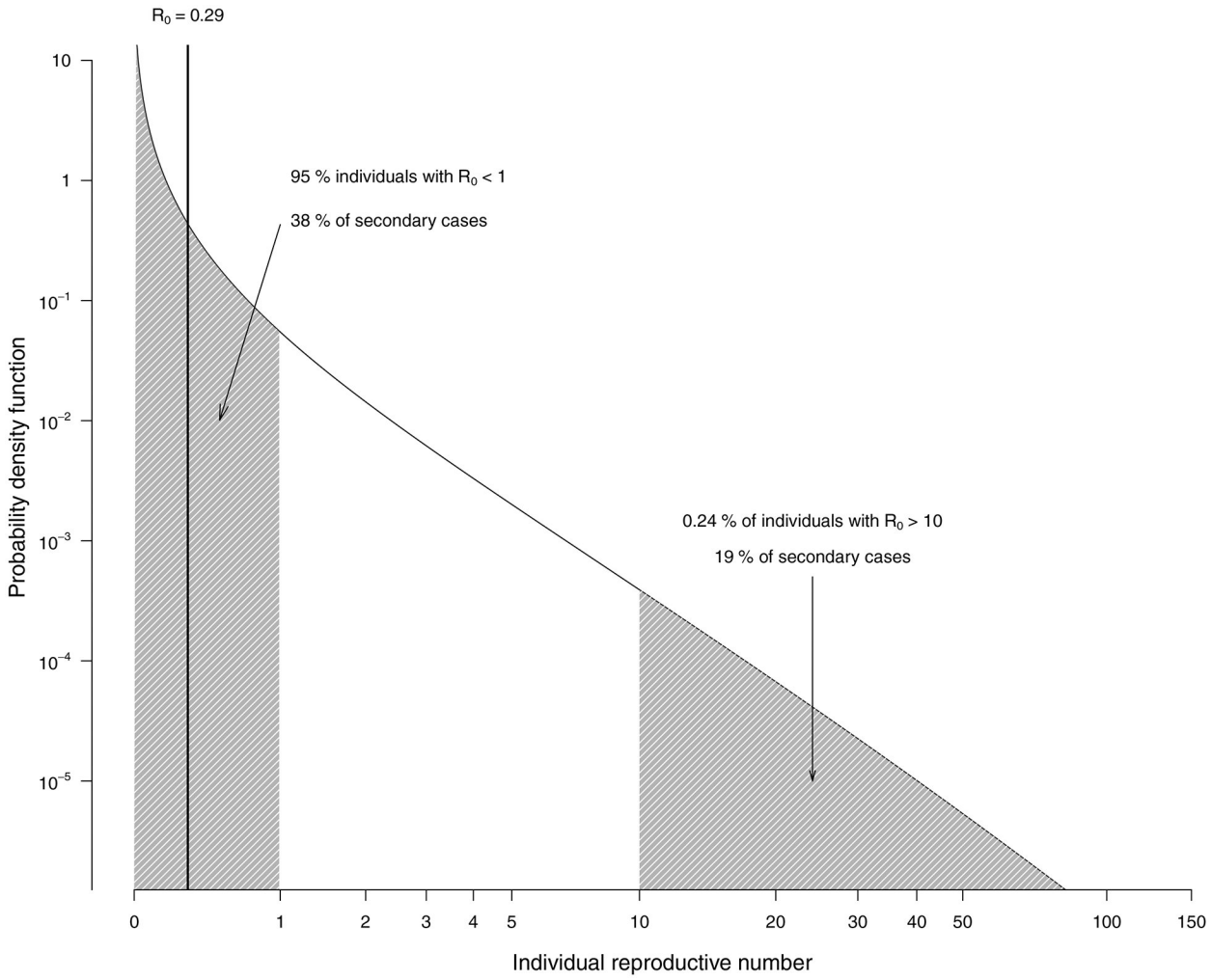

Figure 3.

Underlying individual-level heterogeneity of tuberculosis (TB) transmission. Shown is the probability density function corresponding to the best-fit Poisson lognormal model, describing the distribution of the individual reproductive number under the reference scenario (clustering based on genotyped cases reported within state boundary and occurring between 2012 and 2016). The solid vertical line shows the mean of the distribution (ie, $\widehat{R_0}$), or the estimated average number of secondary transmission cases resulting from a single TB case.

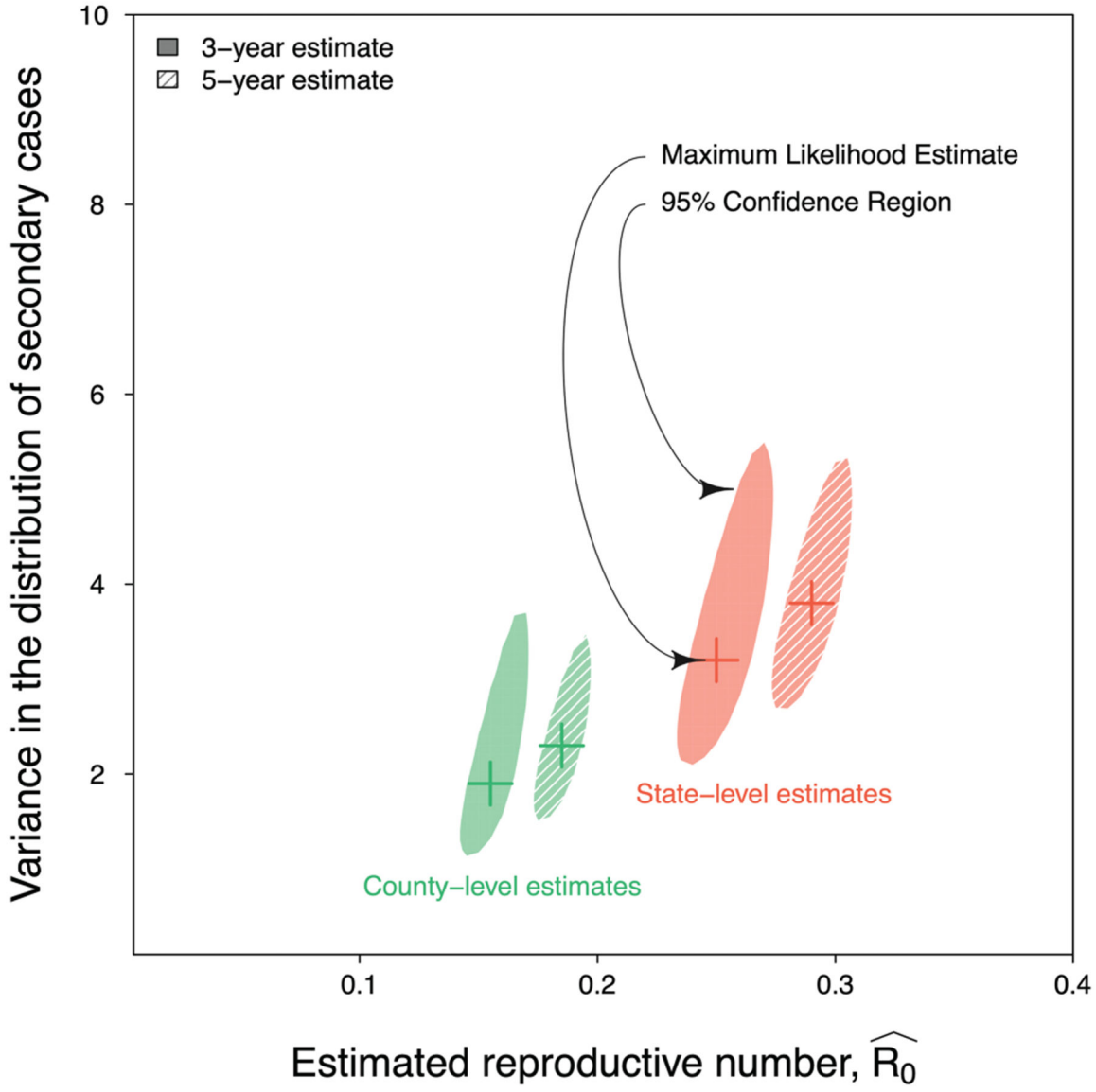

Figure 4.

Comparing model-based inferences under different definitions of tuberculosis clusters in the United States. We fit Poisson lognormal models to 4 cluster distributions, each using a different geographic boundary and time window for cluster ascertainment. Shown are the estimated mean reproductive number (ie, $\widehat{R_0}$) and variance of the distribution of secondary cases when the clusters were defined to include cases reported within state boundaries and occurring within 5-year time window (hatched orange), state boundaries and occurring within a 3-year time window (orange), county boundaries and occurring within a 5-year time window (hatched green), and county boundaries and occurring within a 3-year time window (green). The crosses indicate maximum likelihood estimates, and shaded areas indicate estimated 95% confidence intervals.

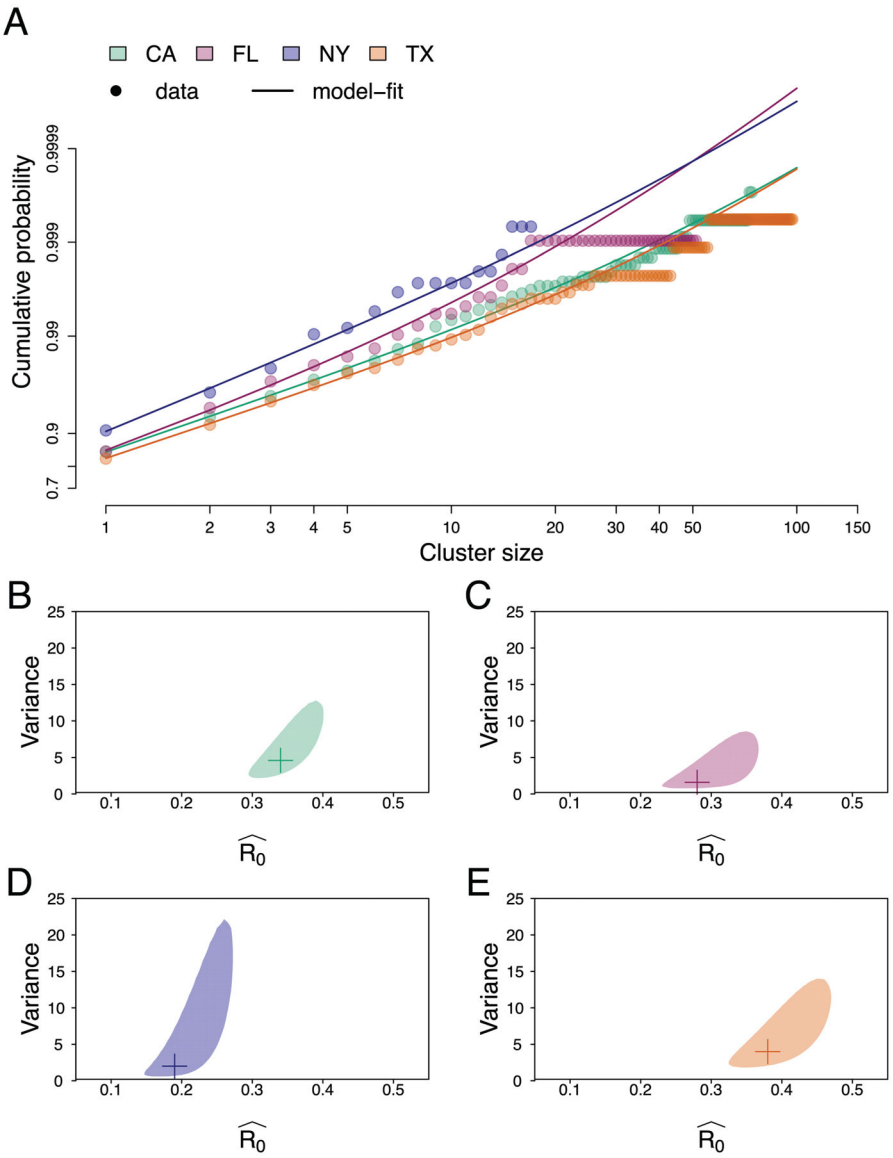

Figure 5.

State-level heterogeneity in tuberculosis cluster distributions and transmission across 4 US states, California, Florida, New York, and Texas, between 2014 and 2016. *A*, Colored circles show cluster distributions in California (green), Florida (violet), New York (blue), and Texas (orange) as cumulative probabilities (ie, the probability of a cluster of given size or less). The colored lines show cumulative probabilities corresponding to the Poisson lognormal model with maximum likelihood estimates. *B–E*, Shown are the estimated reproductive numbers $\widehat{R_0}s$ and variances (in the distribution of secondary cases) for California *(B)*, Florida *(C)*, New York *(D)*, and Texas *(E)*. The shaded colored region indicates the estimated 95% confidence region, and the cross in the middle indicates the maximum likelihood estimate.

Table 1.

Notations and Symbols Used in the Study, Detailed Descriptions, and Underlying Assumptions

Notations and Symbols	Description	Underlying Assumption
R_0	Reproductive number, or average number of secondary cases resulting from a single case	Theoretical concept
v	Individual reproductive number, expected number of secondary cases resulting from each individual	Assumed to vary based on the underlying models; see Table 2
$\widehat{R_0}$	Estimated reproductive number, based on maximum likelihood estimate	Estimates are aimed to capture cases resulting from recent transmission (and exclude cases resulting from reactivation that occur at longer time scales) Estimates are based on genotyped cluster data and subject to limitations of the clustering method (including missing cases, cases not cultured or genotyped, overascertainment of clustering)
Z	Offspring distribution of a branching process that describes the probability distribution of the number of secondary cases resulting from a single case	Varies based on the underlying models; see Table 2

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Table 2.

Description of Four Models of Individual-level Heterogeneity and Comparison of Their Statistical Fits to the Reference Data

Model	Model Description	Underlying Distribution of Individual Reproductive Number, ν; Resulting Distribution of Secondary Cases, *Z*; Variance of *Z*	Maximum Likelihood Estimate, Log Scaled (Difference in Log Likelihood Units Relative to the Highest Estimate)	Relative Likelihood Compared With the Best Model[a]
I, Homogeneous[b]	Assumes no individual-level heterogeneity, that is, all individuals have the **same** reproductive number"	ν is constant; $Z \sim \text{Poisson}(R_0)$; R_0	−16 787.68 (−1450.19)	<1/1000
II, SIR-type[b]	Reflecting assumption in standard Susceptible-Infectious-Recovered-type compartmental models, assumes exponentially distributed individual reproductive numbers	ν is exponentially distributed; $Z \sim \text{geometric}(R_0)$; $R_0(1+R_0)$	−17 804.98 (−2468.19)	<1/1000
III, Overdispersed	Assumes that the number of secondary cases from an individual are overdispersed and the degree of overdispersion is estimated	ν is gamma distributed; $Z \sim \text{negative binomial}(R_0, k)$ k is the dispersion parameter, smaller values relate to larger heterogeneity; $R_0\left(1+\frac{R_0}{k}\right)$	−15 507.78 (−170.99)	<1/1000
IV, Long-tailed	Assumes that individual-level heterogeneity is lognormally distributed (allowing for even larger heterogeneity)	ν is lognormally distributed; $Z \sim$ Poisson lognormal (μ, σ^2) μ, σ^2 are, respectively, mean variance of the underlying normal distribution; $R_0\,[1 + R_0\,(exp(\sigma^2) - 1)]$	**−15 336.79 (Ref)**	

[a] Relative likelihood is given by the quantity $\exp((AIC_{min} - AIC)/2)$, where AIC is the Akaike information criterion and AIC_{min} is the AIC score corresponding to the best-fit model, or the model with the lowest score.

[b] Poisson and geometric models are specific instances of the negative binomial model. Negative binomial model with dispersion parameter $k \to \infty$ is a Poisson model, and k = 1 is a geometric model.